

ON AN EQUIVALENT DEFINITION OF FREE ENTROPY

SERBAN TEODOR BELINSCHI

ABSTRACT. In his second paper on free entropy, Voiculescu suggested five other possible definitions of the entropy of an n -tuple of selfadjoint random variables. We show that two of them are equivalent to the original one.

1. Introduction

Let X_1, \dots, X_n be n selfadjoint random variables in the tracial W^* -probability space (\mathcal{M}, τ) . Denote by \mathcal{M}_k^{sa} the space of $k \times k$ selfadjoint matrices over \mathbb{C} , let Tr_k be the (unnormalized) trace on \mathcal{M}_k^{sa} , and denote by τ_k the normalized trace, $\tau_k = \frac{1}{k} \text{Tr}_k$. Then $(\mathcal{M}_k^{sa})^n$ becomes a real Hilbert space with the norm

$$\|(A_1, \dots, A_n)\|_2 = (\text{Tr}_k(A_1^2 + \dots + A_n^2))^{1/2}.$$

Denote by λ the Lebesgue measure corresponding to this norm (i.e. the Lebesgue measure normalized so that a cube with edges of length one has measure equal to one).

Let $\varepsilon > 0, m > 1, R > \max\{\|X_1\|, \dots, \|X_n\|\}$, and let $\Psi_{m,k} : (\mathcal{M}_k^{sa})^n \rightarrow \mathbb{R}_+$ be defined by

$$\Psi_{m,k}(A_1, \dots, A_n) = \sum_{q=1}^m \sum_{1 \leq i_1, \dots, i_q \leq n} |\tau_k(A_{i_1} \dots A_{i_q}) - \tau(X_{i_1} \dots X_{i_q})|.$$

Following Voiculescu's definition of microstate-based free entropy $\chi(X_1, \dots, X_n)$, let

$$\Gamma_R(X_1, \dots, X_n; m, k, \varepsilon) = \{(A_1, \dots, A_n) \in (\mathcal{M}_k^{sa})^n : \Psi_{m,k}(A_1, \dots, A_n) < \varepsilon, \|A_j\| \leq R\}$$

$$\chi_R(X_1, \dots, X_n; m, k, \varepsilon) = \log \lambda(\Gamma_R(X_1, \dots, X_n; m, k, \varepsilon))$$

$$\chi_R(X_1, \dots, X_n; m, \varepsilon) = \limsup_{k \rightarrow \infty} (k^{-2} \chi_R(X_1, \dots, X_n; m, k, \varepsilon) + 2^{-1} n \log k)$$

$$\chi_R(X_1, \dots, X_n) = \lim_{\varepsilon \rightarrow 0, m \rightarrow \infty} \chi_R(X_1, \dots, X_n; m, \varepsilon)$$

$$\chi(X_1, \dots, X_n) = \sup_{R > 0} \chi_R(X_1, \dots, X_n).$$

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In Section 7.1 of [3], two alternative quantities were defined by Voiculescu as follows. Let $\mathcal{D}_{1,R}(X_1, \dots, X_n; m, k, \varepsilon)$ be the set of probability densities $f : (\mathcal{M}_k^{sa})^n \rightarrow \mathbb{R}_+$ (i.e. $\int f d\lambda = 1$) so that $\int f(A_1, \dots, A_n) \Psi_{m,k}(A_1, \dots, A_n) d\lambda \leq \varepsilon$, and $f(A_1, \dots, A_n) = 0$ unless $\|A_j\| \leq R$ ($1 \leq j \leq n$). Then

$$\chi_R^{(1)}(X_1, \dots, X_n; m, k, \varepsilon) = \sup \left\{ - \int f \log f d\lambda : f \in \mathcal{D}_{1,R}(X_1, \dots, X_n; m, k, \varepsilon) \right\}.$$

$\chi^{(1)}$ is now defined following the same steps as in the definition of χ . One can also define χ_∞ and $\chi_\infty^{(1)}$ by setting $R = \infty$ from the beginning, that is, by leaving out the conditions on the norms of the A_j 's. Since the characteristic function of $\Gamma_R(X_1, \dots, X_n; m, k, \varepsilon)$ (properly normalized) belongs to $\mathcal{D}_{1,R}(X_1, \dots, X_n; m, k, \varepsilon)$, we conclude that

$$\chi_R^{(1)}(X_1, \dots, X_n; m, k, \varepsilon) \geq \chi_R(X_1, \dots, X_n; m, k, \varepsilon).$$

It follows that $\chi \leq \chi^{(1)}$, and, likewise, $\chi_\infty \leq \chi_\infty^{(1)}$.

It has already been proved in [1] that $\chi_\infty = \chi$. In this paper we shall prove that $\chi_\infty^{(1)} = \chi_\infty$, which will obviously imply that $\chi_\infty^{(1)} = \chi^{(1)} = \chi = \chi_\infty$. As an immediate consequence, we obtain a new way to define free entropy replacing $\chi_R(X_1, \dots, X_n; m, k, \varepsilon)$ by:

$$\chi_\infty^{(1)}(X_1, \dots, X_n; m, k, \varepsilon) = \log \int e^{-s(\Psi_{m,k}(A_1, \dots, A_n) - \varepsilon)} d\lambda,$$

and continuing as in the original definition. Here s is an appropriate positive number depending on $X_1, X_2, \dots, X_n, m, \varepsilon$ and k .

2. The main result

Given a probability density $f : \mathbb{R}^p \rightarrow [0, +\infty)$, i.e. $\int_{\mathbb{R}^p} f(x) d\lambda(x) = 1$, we denote by $I(f) = - \int_{\mathbb{R}^p} f(x) \log f(x) d\lambda(x)$ the classical entropy of f . A density of special interest is the Gaussian density $G_{a^2}^{(p)}$ defined on \mathbb{R}^p by $G_{a^2}^{(p)}(x) = \left(\frac{1}{2\pi a^2}\right)^{\frac{p}{2}} e^{-\frac{\|x\|_2^2}{2a^2}}$.

We will need the following extension of Shannon's classical inequality (see [2], pp. 55–57).

Lemma 2.1. *Let $\Omega \subset \mathbb{R}^p$ be a set of positive Lebesgue measure, and $\varepsilon > 0$ be fixed. If φ is a positive continuous function on Ω such that $\int_\Omega \varphi \geq \varepsilon \lambda(\Omega)$, $\varphi^{-1}([0, \varepsilon])$ is nonempty, and both $e^{-\varphi}$ and $\varphi^2 e^{-\varphi}$ are integrable on Ω , then there is a number $s \in (0, \infty)$ depending on φ, p , and ε such that $M :=$*

$$\max \left\{ I(f) : f \geq 0, f = 0 \text{ on } \mathbb{R}^n \setminus \Omega, \int_\Omega f(x) d\lambda(x) = 1, \int_\Omega f(x) \varphi(x) d\lambda(x) = \varepsilon \right\}$$

is attained for $f = \left(\int_\Omega e^{-s\varphi(x)} d\lambda(x)\right)^{-1} e^{-s\varphi}$ and equals $\log \int_\Omega e^{-s(\varphi(x) - \varepsilon)} d\lambda$.

Proof. For each $s > 0$, let $f_s(x) = (\int_{\Omega} e^{-s\varphi(y)} d\lambda(y))^{-1} e^{-s\varphi(x)}$. Consider the function $h : (0, \infty) \rightarrow (0, \infty)$ given by $h(s) = \int_{\Omega} \varphi(x) f_s(x) d\lambda(x)$. Since both $e^{-\varphi}$ and $\varphi^2 e^{-\varphi}$ are integrable on Ω , we can differentiate under the integral sign, and a straightforward calculation shows that

$$h'(s) = - \int_{\Omega} \left(\varphi(x) - \int_{\Omega} \varphi(y) f_s(y) d\lambda(y) \right)^2 f_s(x) d\lambda(x),$$

so $h'(s) < 0$. We also have that

$$\lim_{s \rightarrow \infty} h(s) = \min_{x \in \Omega} \varphi(x),$$

which is strictly less than ε , and $\lim_{s \rightarrow 0} h(s)$ equals $\frac{1}{\lambda(\Omega)} \int_{\Omega} \varphi$ if $\lambda(\Omega) < \infty$ and infinity if $\lambda(\Omega) = \infty$, which are both greater than or equal to ε . Hence the existence of s is assured by the monotonicity of h . Moreover, since φ is nonconstant, s is unique with this property.

Let

$$g \in \left\{ f : f \geq 0, \int_{\Omega} f d\lambda = 1, \int_{\Omega} f \varphi d\lambda = \varepsilon \right\}.$$

We may assume that $g(x) < \infty$ for all $x \in \Omega$. If $I(g) = -\infty$, then it is obvious that $I(f_s) > I(g)$. So suppose that $I(g) > -\infty$. Consider the function $\eta : [0, 1] \rightarrow \mathbb{R}$ defined by $\eta(t) = I((1-t)f_0 + tg)$. This function is continuous on $[0, 1]$, differentiable on $(0, 1)$, and twice differentiable on $(0, 1)$. Moreover, $\eta'(t) = - \int_{\Omega} (g - f_s) \log((1-t)f_s + tg) + 1$, and $\eta''(t) = - \int_{\Omega} \frac{(g-f_s)^2}{(1-t)f_s + tg}$. (This follows immediately from the dominated convergence theorem. Indeed,

$$|(g - f_s) \log((1-t)f_s + tg) + 1| \leq g + f_s + |(g - f_s) \log((1-t)f_s + tg)|$$

for all $t \in [0, 1)$, and for any compact interval $K \subset [0, 1)$, $\sup_{t \in K} |(g - f_s) \log((1-t)f_s + tg) + 1|$ is integrable on Ω . Also,

$$\left| \frac{(g - f_s)^2}{(1-t)f_s + tg} \right| \leq \frac{1}{t} g + \frac{1}{1-t} f_s$$

for all $t \in (0, 1)$, and, of course, $\sup_{t \in K} |\frac{1}{t} g + \frac{1}{1-t} f_s|$ is integrable on Ω for any compact interval $K \subset (0, 1)$.)

By direct computation, we obtain that $\eta'(0) = 0$. Also, η' is continuous on $[0, 1)$, and $\eta''(t) \leq 0$ for all $t \in (0, 1)$, so the point $t = 0$ is a global maximum for η . This proves that $I(f_s) = M$ and concludes the proof. \square

Remark 2.2. To see how this lemma extends Shannon's inequality, set $\Omega = \mathbb{R}^n$, and $\varphi(x) = \|x\|_2^2$. In this case the extremal density f is $G_{\varepsilon}^{(p)}$ and $I(G_{\varepsilon}^{(p)}) = 2^{-1} p \log(2\pi e p^{-1} \varepsilon)$. Also note that for $\lambda(\Omega) < +\infty$ and $\varphi = \varepsilon$, the extremal density is $\frac{1}{\lambda(\Omega)} \chi_{\Omega}$ and $I(\frac{1}{\lambda(\Omega)} \chi_{\Omega}) = \log \lambda(\Omega)$.

Now we can prove the main result of the paper.

Theorem 2.1. *Let X_1, \dots, X_n be selfadjoint random variables in a tracial W^* -probability space (\mathcal{M}, τ) . Then*

$$\chi_\infty(X_1, \dots, X_n) = \chi_\infty^{(1)}(X_1, \dots, X_n).$$

Proof. As noted in the introduction, $\chi_\infty(X_1, \dots, X_n) \leq \chi_\infty^{(1)}(X_1, \dots, X_n)$, so it is enough to prove the opposite inequality. Let $m \geq 2$ and $\varepsilon \in (0, 1/2)$ be fixed. Let $V_0 = \{(A_1, \dots, A_n) \in (\mathcal{M}_k^{sa})^n : \Psi_{m,k}(A_1, \dots, A_n) \leq \varepsilon\}$, and $V_1 = (\mathcal{M}_k^{sa})^n \setminus V_0$. Consider $f_0 \in \mathcal{D}_{1,\infty}(X_1, \dots, X_n; m, k, \varepsilon^2)$ such that $I(f_0) = \max\{I(f) : f \in \Gamma_\infty^{(1)}(X_1, \dots, X_n; m, k, \varepsilon^2)\}$, and define $f_0^{(i)} = c_i^{-1} f_0 \chi_{V_i}$, where $c_i = \int_{V_i} f$, $i = 0, 1$ depend on X_1, \dots, X_n, m, k and ε . Then we have:

$$\varepsilon^2 \geq \int f_0 \Psi_{m,k} \geq \int_{V_1} f_0 \Psi_{m,k} \geq \varepsilon \int_{V_1} f_0 = \varepsilon c_1,$$

so $c_1 \leq \varepsilon$, and $c_0 \geq 1 - \varepsilon$.

Since $f_0 \in \mathcal{D}_{1,\infty}(X_1, \dots, X_n; m, k, \varepsilon^2)$,

$$\int f_0^{(1)} \Psi_{m,k} = \left(\int_{V_1} f_0 \right)^{-1} \int_{V_1} f_0 \Psi_{m,k} < \frac{\varepsilon^2}{c_1},$$

which implies that $f_0^{(1)} \in \mathcal{D}_{1,\infty}(X_1, \dots, X_n; m, k, \frac{\varepsilon^2}{c_1})$.

Now,

$$\begin{aligned} I(f_0) &= c_0 I(f_0^{(0)}) + c_1 I(f_0^{(1)}) - c_0 \log c_0 - c_1 \log c_1 \\ &\leq c_0 I\left(\frac{1}{|V_0|} \chi_{V_0}\right) + c_1 I(G_{k(C^2 + n\varepsilon^2/c_1)}^{(nk^2)}) - c_0 \log c_0 - c_1 \log c_1 \\ &= c_0 \chi_\infty(X_1, \dots, X_n; m, k, \varepsilon) + c_1 \frac{nk^2}{2} \left(\log \frac{2\pi e(C^2 + n\varepsilon^2/c_1)}{n} - \log k \right) \\ (1) \quad &- c_0 \log c_0 - c_1 \log c_1, \end{aligned}$$

where $C^2 = \tau(X_1^2 + \dots + X_n^2)$. Here we used Lemma 2.1 in the last inequality.

Dividing by k^2 and adding $\frac{n}{2} \log k$ in (1), we obtain:

$$\begin{aligned} \frac{1}{k^2} I(f_0) + \frac{n}{2} \log k &\leq c_0 \left(\frac{1}{k^2} \chi_\infty(X_1, \dots, X_n; m, k, \varepsilon) + \frac{n}{2} \log k \right) \\ &\quad + c_1 \frac{n}{2} \log \frac{2\pi e(C^2 + n\varepsilon^2/c_1)}{n} - \frac{1}{k^2} (c_0 \log c_0 + c_1 \log c_1). \end{aligned}$$

As $k \rightarrow \infty$, this yields

$$\begin{aligned} \chi_\infty^{(1)}(X_1, \dots, X_n; m, \varepsilon^2) &\leq \limsup_{k \rightarrow \infty} c_0 \left(\frac{1}{k^2} \chi_\infty(X_1, \dots, X_n; m, k, \varepsilon) + \frac{n}{2} \log k \right) \\ &\quad + \limsup_{k \rightarrow \infty} c_1 \frac{n}{2} \log \frac{2\pi e(C^2 + n\varepsilon^2/c_1)}{n} \end{aligned}$$

Since $\lim_{t \rightarrow 0, t > 0} t \log \frac{2\pi e(C^2 + n\varepsilon^2/t)}{n} = 0$, the function

$$s(\varepsilon) := \sup \left\{ t \frac{n}{2} \log \frac{2\pi e(C^2 + n\varepsilon^2/t)}{n} : t \in (0, \varepsilon] \right\}$$

does not depend on k and m and tends to zero as $\varepsilon \rightarrow 0$.

Hence

$$\chi_\infty^{(1)}(X_1, \dots, X_n; m, \varepsilon^2) \leq \limsup_{k \rightarrow \infty} c_0 \left(\frac{1}{k^2} \chi_\infty(X_1, \dots, X_n; m, k, \varepsilon) + \frac{n}{2} \log k \right) + s(\varepsilon)$$

On the other hand, since $1 \geq c_0 \geq 1 - \varepsilon$, we have that

$$\begin{aligned} \limsup_{k \rightarrow \infty} c_0 \left(\frac{1}{k^2} \chi_\infty(X_1, \dots, X_n; m, k, \varepsilon) + \frac{n}{2} \log k \right) \leq \\ \max \{ (1 - \varepsilon) \chi_\infty(X_1, \dots, X_n; m, \varepsilon), \chi_\infty(X_1, \dots, X_n; m, \varepsilon) \}. \end{aligned}$$

This allows us to conclude that

$$\begin{aligned} \chi_\infty^{(1)}(X_1, \dots, X_n; m, \varepsilon^2) \leq \\ \max \{ (1 - \varepsilon) \chi_\infty(X_1, \dots, X_n; m, \varepsilon), \chi_\infty(X_1, \dots, X_n; m, \varepsilon) \} + s(\varepsilon). \end{aligned}$$

Taking limits when $m \rightarrow \infty$ and $\varepsilon \rightarrow 0$ in the previous inequality, we obtain

$$\chi_\infty^{(1)}(X_1, \dots, X_n) = \chi_\infty(X_1, \dots, X_n).$$

□

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MATHEMATICS DEPARTMENT, INDIANA UNIVERSITY, BLOOMINGTON, INDIANA 47405, U.S.A.
E-mail address: sbelinsc@indiana.edu

INSTITUTE OF MATHEMATICS, ROMANIAN ACADEMY, P.O. BOX 1-764, BUCHAREST, RO-70700, ROMANIA.