

Quadrant marked mesh patterns in alternating permutations II

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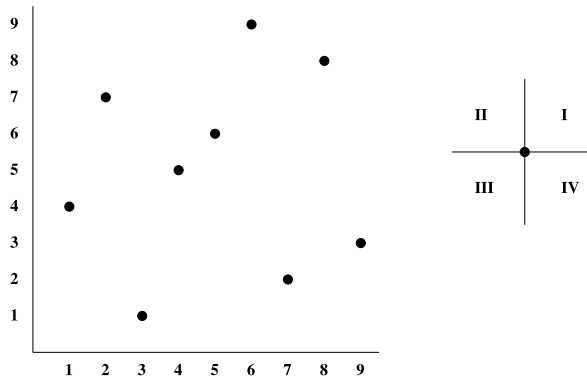
This paper is a continuation of the systematic study of distribution of quadrant marked mesh patterns initiated in [7]. We study quadrant marked mesh patterns on up-down and down-up permutations.

1. Introduction

Let $\sigma = \sigma_1 \cdots \sigma_n$ be a permutation in the symmetric group S_n written in one-line notation. Then we will consider the graph of σ , $G(\sigma)$, to be the set of points (i, σ_i) for $i = 1, \dots, n$. For example, the graph of the permutation $\sigma = 471569283$ is pictured in Figure 1. Then if we draw a coordinate system centered at a point (i, σ_i) , we will be interested in the points that lie in the four quadrants I, II, III, and IV of that coordinate system as pictured in Figure 1.

For any $a, b, c, d \in \mathbb{N}$ where $\mathbb{N} = \{0, 1, 2, \dots\}$ is the set of natural numbers and any $\sigma = \sigma_1 \cdots \sigma_n \in S_n$, we say that σ_i matches the *quadrant marked mesh pattern* $MMP(a, b, c, d)$ in σ if in $G(\sigma)$ relative to the coordinate system which has the point (i, σ_i) as its origin, there are $\geq a$ points in quadrant I, $\geq b$ points in quadrant II, $\geq c$ points in quadrant III, and $\geq d$ points in quadrant IV. For example, if $\sigma = 471569283$, the point $\sigma_4 = 5$ matches the quadrant marked mesh pattern $MMP(2, 1, 2, 1)$ since relative to the coordinate system with origin $(4, 5)$, there are 3 points in $G(\sigma)$ in quadrant I, 1 point in $G(\sigma)$ in quadrant II, 2 points in $G(\sigma)$ in quadrant III, and 2 points in $G(\sigma)$ in quadrant IV.

Note that if a coordinate in $MMP(a, b, c, d)$ is 0, then there is no condition imposed on the points in the corresponding quadrant. In addition, we shall consider patterns $MMP(a, b, c, d)$ where $a, b, c, d \in \mathbb{N} \cup \{\emptyset\}$. Here when one of the parameters a, b, c , or d in $MMP(a, b, c, d)$ is the empty set, then for σ_i to match $MMP(a, b, c, d)$ in $\sigma = \sigma_1 \cdots \sigma_n \in S_n$, it must be the case that there are no points in $G(\sigma)$ relative to coordinate system with

Figure 1: The graph of $\sigma = 471569283$.

origin (i, σ_i) in the corresponding quadrant. For example, if $\sigma = 471569283$, the point $\sigma_3 = 1$ matches the marked mesh pattern $MMP(4, 2, \emptyset, \emptyset)$ since relative to the coordinate system with origin $(3, 1)$, there are 6 points in $G(\sigma)$ in quadrant I, 2 points in $G(\sigma)$ in quadrant II, no points in $G(\sigma)$ in quadrant III, and no points in $G(\sigma)$ in quadrant IV. We let $\text{mmp}^{(a,b,c,d)}(\sigma)$ denote the number of i such that σ_i matches the marked mesh pattern $MMP(a, b, c, d)$ in σ .

The notion of *mesh patterns* was introduced by Brändén and Claesson [4] to provide explicit expansions for certain permutation statistics as, possibly infinite, linear combinations of (classical) permutation patterns (see [6] for a comprehensive introduction to the theory of permutation patterns). This notion was further studied in [3, 5, 7–11, 14].

Note how the (two-dimensional) notation of Úlfarsson [14] for *marked mesh patterns* corresponds to our (one-line) notation for quadrant marked mesh patterns. For example,

$$\begin{aligned}
 MMP(0, 0, k, 0) &= \begin{array}{c} | \\ \bullet \\ | \\ \boxed{k} \end{array}, & MMP(k, 0, 0, 0) &= \begin{array}{c} \boxed{k} \\ | \\ \bullet \\ | \end{array}, \\
 MMP(0, a, b, c) &= \begin{array}{c} \boxed{a} \\ | \\ \bullet \\ | \\ \boxed{b} \quad \boxed{c} \end{array} & \text{and} & MMP(0, 0, \emptyset, k) &= \begin{array}{c} | \\ \bullet \\ | \\ \text{shaded} \quad \boxed{k} \end{array}.
 \end{aligned}$$

Kitaev and Remmel [7] studied the distribution of quadrant marked mesh patterns in the symmetric group S_n , and Kitaev, Remmel, and Tiefenbruck [9, 10] studied the distribution of quadrant marked mesh patterns

in 132-avoiding permutations in S_n . In [8], Kitaev and Remmel studied the distribution of the statistics $\text{mmp}^{(1,0,0,0)}$, $\text{mmp}^{(0,1,0,0)}$, $\text{mmp}^{(0,0,1,0)}$, and $\text{mmp}^{(0,0,0,1)}$ in the set of up-down and down-up permutations. The main goal of this paper is to study the distribution of the statistics $\text{mmp}^{(1,0,\emptyset,0)}$, $\text{mmp}^{(0,1,0,\emptyset)}$, $\text{mmp}^{(0,\emptyset,0,1)}$, and $\text{mmp}^{(\emptyset,0,1,0)}$ in the set of up-down and down-up permutations. Given a permutation $\sigma = \sigma_1 \cdots \sigma_n \in S_n$, we let $\text{Des}(\sigma) = \{i : \sigma_i > \sigma_{i+1}\}$. Then we say that σ is an *up-down permutation* if $\text{Des}(\sigma)$ is the set of all even numbers less than or equal to n and a *down-up permutation* if $\text{Des}(\sigma)$ is the set of all odd numbers less than or equal to n . That is, σ is an up-down permutation if

$$\sigma_1 < \sigma_2 > \sigma_3 < \sigma_4 > \sigma_5 < \cdots$$

and σ is a down-up permutation if

$$\sigma_1 > \sigma_2 < \sigma_3 > \sigma_4 < \sigma_5 > \cdots.$$

Let UD_n denote the set of all up-down permutations in S_n and DU_n denote the set of all down-up permutations in S_n . Given a permutation $\sigma = \sigma_1 \cdots \sigma_n \in S_n$, we define the reverse of σ , σ^r , to be $\sigma_n \sigma_{n-1} \cdots \sigma_2 \sigma_1$ and the complement of σ , σ^c , to be $(n+1-\sigma_1) \cdots (n+1-\sigma_n)$.

For $n \geq 1$, we let

$$\begin{aligned} A_{2n}^{(a,b,c,d)}(x) &= \sum_{\sigma \in UD_{2n}} x^{\text{mmp}^{(a,b,c,d)}(\sigma)}, & B_{2n-1}^{(a,b,c,d)}(x) &= \sum_{\sigma \in UD_{2n-1}} x^{\text{mmp}^{(a,b,c,d)}(\sigma)}, \\ C_{2n}^{(a,b,c,d)}(x) &= \sum_{\sigma \in DU_{2n}} x^{\text{mmp}^{(a,b,c,d)}(\sigma)}, & D_{2n-1}^{(a,b,c,d)}(x) &= \sum_{\sigma \in DU_{2n-1}} x^{\text{mmp}^{(a,b,c,d)}(\sigma)}. \end{aligned}$$

We then have the following simple proposition.

Proposition 1. For all $n \geq 1$,

- (1) $A_{2n}^{(a,b,c,d)}(x) = C_{2n}^{(b,a,d,c)}(x) = C_{2n}^{(d,c,b,a)}(x) = A_{2n}^{(c,d,a,b)}(x)$,
- (2) $C_{2n}^{(a,b,c,d)}(x) = A_{2n}^{(b,a,d,c)}(x) = A_{2n}^{(d,c,b,a)}(x) = C_{2n}^{(c,d,a,b)}(x)$,
- (3) $B_{2n-1}^{(a,b,c,d)}(x) = B_{2n-1}^{(b,a,d,c)}(x) = D_{2n-1}^{(d,c,b,a)}(x) = D_{2n-1}^{(c,d,a,b)}(x)$, and
- (4) $D_{2n-1}^{(a,b,c,d)}(x) = D_{2n-1}^{(b,a,d,c)}(x) = B_{2n-1}^{(d,c,b,a)}(x) = B_{2n-1}^{(c,d,a,b)}(x)$.

Proof. It is easy to see that for any $\sigma \in S_n$,

$$\text{mmp}^{(a,b,c,d)}(\sigma) = \text{mmp}^{(b,a,d,c)}(\sigma^r) = \text{mmp}^{(d,c,b,a)}(\sigma^c) = \text{mmp}^{(c,d,a,b)}((\sigma^r)^c).$$

Then part 1 easily follows since

$$\sigma \in UD_{2n} \iff \sigma^r \in DU_{2n} \iff \sigma^c \in DU_{2n} \iff (\sigma^r)^c \in UD_{2n}.$$

Parts 2, 3, and 4 are proved in a similar manner. \square

In [8], we studied the distribution of the statistics $\text{mmp}^{(1,0,0,0)}$, $\text{mmp}^{(0,1,0,0)}$, $\text{mmp}^{(0,0,1,0)}$, and $\text{mmp}^{(0,0,0,1)}$ in the set of up-down and down-up permutations. It follows from Proposition 1 that the study of the distribution of the statistics $\text{mmp}^{(1,0,0,0)}$, $\text{mmp}^{(0,1,0,0)}$, $\text{mmp}^{(0,0,1,0)}$, and $\text{mmp}^{(0,0,0,1)}$ in the set of up-down and down-up permutations can be reduced to the study of the following generating functions:

$$\begin{aligned} A^{(1,0,0,0)}(t, x) &= 1 + \sum_{n \geq 1} A_{2n}^{(1,0,0,0)}(x) \frac{t^{2n}}{(2n)!}, \\ B^{(1,0,0,0)}(t, x) &= \sum_{n \geq 1} B_{2n-1}^{(1,0,0,0)}(x) \frac{t^{2n-1}}{(2n-1)!}, \\ C^{(1,0,0,0)}(t, x) &= 1 + \sum_{n \geq 1} C_{2n}^{(1,0,0,0)}(x) \frac{t^{2n}}{(2n)!}, \text{ and} \\ D^{(1,0,0,0)}(t, x) &= \sum_{n \geq 1} D_{2n-1}^{(1,0,0,0)}(x) \frac{t^{2n-1}}{(2n-1)!}. \end{aligned}$$

In the case when $x = 1$, these generating functions are well known. That is, for any (a, b, c, d) , let $A_{2n}(1) = A_{2n}^{(a,b,c,d)}(1)$, $B_{2n+1}(1) = B_{2n+1}^{(a,b,c,d)}(1)$, $C_{2n}(1) = C_{2n}^{(a,b,c,d)}(1)$, and $D_{2n}(1) = D_{2n}^{(a,b,c,d)}(1)$. The operation of complementation shows that $A_{2n}(1) = C_{2n}(1)$ and $B_{2n-1}(1) = D_{2n-1}(1)$ for all $n \geq 1$. André [1, 2] proved that

$$(1) \quad 1 + \sum_{n \geq 0} A_{2n}(1) \frac{t^{2n}}{(2n)!} = \sec(t)$$

and

$$(2) \quad \sum_{n \geq 1} B_{2n-1}(1) \frac{t^{2n+1}}{(2n+1)!} = \tan(t).$$

In [8], we proved the following which can be viewed as a refinement of André's results.

Theorem 1.

$$(3) \quad A^{(1,0,0,0)}(t, x) = (\sec(xt))^{1/x},$$

$$(4) \quad B^{(1,0,0,0)}(t, x) = (\sec(xt))^{1/x} \int_0^t (\sec(xz))^{-1/x} dz,$$

$$(5) \quad C^{(1,0,0,0)}(t, x) = 1 + \int_0^t (\sec(xy))^{1+\frac{1}{x}} \int_0^y (\sec(xz))^{1/x} dz dy, \text{ and}$$

$$(6) \quad vD^{(1,0,0,0)}(t, x) = \int_0^t (\sec(xz))^{1+\frac{1}{x}} dz.$$

In this paper, we prove a different refinement of André's results by studying the distribution of the statistics $\text{mmp}^{(1,0,\emptyset,0)}$, $\text{mmp}^{(\emptyset,0,1,0)}$, $\text{mmp}^{(0,1,0,\emptyset)}$, and $\text{mmp}^{(0,\emptyset,0,1)}$ in the set of up-down and down-up permutations. It follows from Proposition 1 that the study of the distribution of the statistics $\text{mmp}^{(1,0,\emptyset,0)}$, $\text{mmp}^{(\emptyset,0,1,0)}$, $\text{mmp}^{(0,1,0,\emptyset)}$, and $\text{mmp}^{(0,\emptyset,0,1)}$ in the set of up-down and down-up permutations can be reduced to the study of the following generating functions:

$$\begin{aligned} A^{(1,0,\emptyset,0)}(t, x) &= 1 + \sum_{n \geq 1} A_{2n}^{(1,0,\emptyset,0)}(x) \frac{t^{2n}}{(2n)!}, \\ B^{(1,0,\emptyset,0)}(t, x) &= \sum_{n \geq 1} B_{2n-1}^{(1,0,\emptyset,0)}(x) \frac{t^{2n-1}}{(2n-1)!}, \\ C^{(1,0,\emptyset,0)}(t, x) &= 1 + \sum_{n \geq 1} C_{2n}^{(1,0,\emptyset,0)}(x) \frac{t^{2n}}{(2n)!}, \text{ and} \\ D^{(1,0,\emptyset,0)}(t, x) &= \sum_{n \geq 1} D_{2n-1}^{(1,0,\emptyset,0)}(x) \frac{t^{2n-1}}{(2n-1)!}. \end{aligned}$$

The main goal of this paper is prove the following theorem.

Theorem 2.

$$\begin{aligned} A^{(1,0,\emptyset,0)}(t, x) &= (\sec(t))^x, \\ B^{(1,0,\emptyset,0)}(t, x) &= \frac{\sin(t) \cos(t)(1 - x + x \sec(t))}{x + (1 - x) \cos(t)} \\ &\quad \times \left((1 - x) {}_2F_1 \left(\frac{1}{2}, \frac{1+x}{2}; \frac{3}{2}; (\sin(t))^2 \right) \right. \\ &\quad \left. + x {}_2F_1 \left(\frac{1}{2}, \frac{2+x}{2}; \frac{3}{2}; (\sin(t))^2 \right) \right), \end{aligned}$$

$$D^{(1,0,\emptyset,0)}(t, x) = x(\sec(t))^x \int_0^t (\cos(z))^x dz + (1-x) \int_0^t (\sec(z))^x dz, \text{ and}$$

$$C^{(1,0,\emptyset,0)}(t, x) = 1 + \int_0^t \left[x(\sec(z))^x (1-x+x\sec(z)) \int_0^z \cos(y) dy \right] dz$$

$$+ (1-x) \int_0^t B^{(1,0,\emptyset,0)}(t, z) dz.$$

Here ${}_2F_1(a, b; c; z) = \sum_{n=0}^{\infty} \frac{(a)_n (b)_n}{(c)_n} \frac{z^n}{n!}$ where $(x)_n = x(x-1)\cdots(x-n+1)$ if $n \geq 1$ and $(x)_0 = 1$.

One can use these generating functions to find some initial values of the polynomials $A_{2n}^{(1,0,\emptyset,0)}(x)$, $B_{2n-1}^{(1,0,\emptyset,0)}(x)$, $C_{2n}^{(1,0,\emptyset,0)}(x)$, and $D_{2n-1}^{(1,0,\emptyset,0)}(x)$. For example, we have used Mathematica to compute the following tables.

n	$A_{2n}^{(1,0,\emptyset,0)}(x)$
0	1
1	x
2	$x(2+3x)$
3	$x(16+30x+15x^2)$
4	$x(272+588x+420x^2+105x^3)$
5	$x(7936+18960x+16380x^2+6300x^3+945x^4)$
6	$x(353792+911328x+893640x^2+429660x^3+103950x^4+10395x^5)$

n	$B_{2n+1}^{(1,0,\emptyset,0)}(x)$
0	1
1	$2x$
2	$x(7+9x)$
3	$x(77+135x+60x^2)$
4	$x(1657+3444x+2310x^2+525x^3)$
5	$x(58457+135945x+112770x^2+40950x^3+5670x^4)$
6	$x(3056557+7715664x+7347945x^2+3395700x^3+777625x^4+72765x^5)$

n	$C_{2n}^{(1,0,\emptyset,0)}(x)$
0	1
1	1
2	$x(2+3x)$
3	$x(7+35x+19x^2)$
4	$x(77+581x+571x^2+156x^3)$
5	$x(1657+16428x+21066x^2+9738x^3+1587x^4)$
6	$x(58457+712579x+1079747x^2+652452x^3+180240x^4+19290x^5)$

n	$D_{2n+1}^{(1,0,\emptyset,0)}(x)$
0	1
1	$x(1+x)$
2	$x(2+9x+5x^2)$
3	$x(16+110x+113x^2+33x^3)$
4	$x(272+2492x+3288x^2+1605x^3+279x^4)$
5	$x(7936+90384x+139756x^2+87456x^3+25365x^4+2895x^5)$
6	$x(353792+4803040x+8323816x^2+6110100x^3+2297778x^4+444045x^5+35685x^6)$

The outline of this paper is as follows. In Section 2, we shall prove Theorem 2. Then in Section 3, we shall show how several of the entries of the tables above can be explained. In particular, we will derive formulas for the coefficient of the highest and lowest coefficient of x in the polynomials the polynomials $A_{2n}^{(1,0,\emptyset,0)}(x)$, $B_{2n+1}^{(1,0,\emptyset,0)}(x)$, $C_{2n}^{(1,0,\emptyset,0)}(x)$, and $D_{2n+1}^{(1,0,\emptyset,0)}(x)$, as well as formulas for the second highest and second lowest coefficient of x in these polynomials. Finally, in Section 4, we shall discuss some connections with our previous work [8] on quadrant marked mesh patterns in alternating permutations, as well as some directions for further research.

In this paper, we need the following notation. Given a sequence $\sigma = \sigma_1 \cdots \sigma_n$ of distinct integers, we let $\text{red}(\sigma)$ be the permutation found by replacing the i th largest integer that appears in σ by i . For example, if $\sigma = 2854$, then $\text{red}(\sigma) = 1432$.

2. Proof of Theorem 2

The proof of all parts of Theorem 2 proceed in the same manner. That is, there are simple recursions satisfied by the polynomials $A_{2n}^{(1,0,\emptyset,0)}(x)$, $B_{2n+1}^{(1,0,\emptyset,0)}(x)$, $C_{2n}^{(1,0,\emptyset,0)}(x)$, and $D_{2n+1}^{(1,0,\emptyset,0)}(x)$ based on the possible positions of 1 in an up-down or a down-up permutation.

2.1. The generating function $A^{(1,0,\emptyset,0)}(t, x)$

If $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$, then 1 must occur in one of the positions $1, 3, \dots, 2n-1$. Let $UD_{2n}^{(2k+1)}$ denote the set of permutations $\sigma \in UD_{2n}$ such that $\sigma_{2k+1} = 1$. A schematic diagram of an element in $UD_{2n}^{(2k+1)}$ is pictured in Figure 2.

Consider a $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}^{(2k+1)}$ where $0 \leq k \leq n-1$. Note that there are $\binom{2n-1}{2k}$ ways to pick the elements which occur to the right of posi-

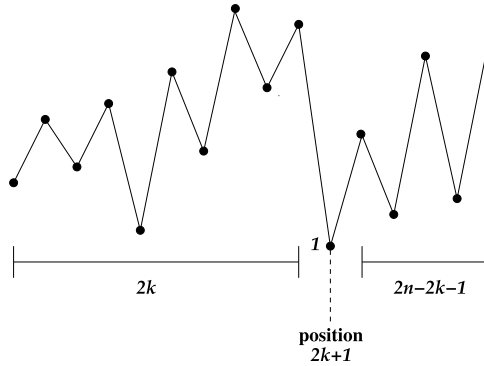


Figure 2: The graph of a $\sigma \in UD_{2n}^{(2k+1)}$.

tion $2k+1$ in such a σ and there are $D_{2n-2k-1}(1) = B_{2n-2k-1}(1)$ ways to order them since the elements to the right of position $2k+1$ must form a down-up permutation of length $2n-2k-1$. The fact that $\sigma_{2k+1} = 1$ implies that σ_{2k+1} matches $MMP(1, 0, \emptyset, 0)$ in σ and that none of the elements to the right of position $2k+1$ match $MMP(1, 0, \emptyset, 0)$ in σ . Thus, the contribution of the elements to the right of position $2k+1$ in $\sum_{\sigma \in UD_{2n}^{(2k+1)}} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)}$ is $B_{2n-2k-1}(1)$. Now the only possible elements of $\sigma_1, \dots, \sigma_{2k}$ that can contribute to $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$ are $\sigma_1, \sigma_3, \dots, \sigma_{2k-1}$. Since each of the elements have an element to its right in $\sigma_1 \cdots \sigma_{2k}$ which is larger than that element, it follows that the elements to the right of position $2k+1$ have no effect on whether $\sigma_1, \dots, \sigma_{2k}$ can contribute to $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$. Hence, the contribution of the elements to the left of position $2k+1$ in $\sum_{\sigma \in UD_{2n}^{(2k+1)}} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)}$ is $A_{2k}^{(1,0,\emptyset,0)}(x)$. It thus follows that for $n \geq 1$,

$$A_{2n}^{(1,0,\emptyset,0)}(x) = x \sum_{k=0}^{n-1} \binom{2n-1}{2k} B_{2n-2k-1}(1) A_{2k}^{(1,0,\emptyset,0)}(x)$$

or, equivalently,

$$(7) \quad \frac{A_{2n}^{(1,0,\emptyset,0)}(x)}{(2n-1)!} = x \sum_{k=0}^{n-1} \frac{B_{2n-2k-1}(1)}{(2n-2k-1)!} \frac{A_{2k}^{(1,0,\emptyset,0)}(x)}{(2k)!}.$$

Multiplying both sides of (7) by t^{2n-1} and summing for $n \geq 1$, we see that

$$\sum_{n \geq 1} \frac{A_{2n}^{(1,0,\emptyset,0)}(x) t^{2n-1}}{(2n-1)!} = x \left(\sum_{k \geq 1} \frac{B_{2k-1}(1) t^{2k-1}}{(2k-1)!} \right) \left(\sum_{k \geq 0} \frac{A_{2k}^{(1,0,\emptyset,0)}(x) t^{2k}}{(2k)!} \right).$$

By (2),

$$\sum_{k \geq 1} \frac{B_{2k-1}(1)t^{2k-1}}{(2k-1)!} = \tan(t)$$

so that

$$(8) \quad \frac{\partial}{\partial t} A^{(1,0,\emptyset,0)}(t, x) = x \tan(t) A^{(1,0,\emptyset,0)}(t, x).$$

Our initial condition is that $A^{(1,0,\emptyset,0)}(0, x) = 1$. It is easy to check that the solution to this differential equation is

$$A^{(1,0,\emptyset,0)}(t, x) = (\sec(t))^x.$$

2.2. The generating function $B^{(1,0,\emptyset,0)}(t, x)$

If $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in UD_{2n+1}$, then 1 must occur in one of the positions $1, 3, \dots, 2n+1$. Let $UD_{2n+1}^{(2k+1)}$ denote the set of permutations $\sigma \in UD_{2n+1}$ such that $\sigma_{2k+1} = 1$. A schematic diagram of an element in $UD_{2n}^{(2k+1)}$ is pictured in Figure 3.

A permutation $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in UD_{2n+1}^{(2k+1)}$ ends with 1 so that $\sigma_{2n+1} = 1$ does not match $MMP(1, 0, \emptyset, 0)$ in σ . Moreover, $\text{red}(\sigma_1 \cdots \sigma_{2n}) \in UD_{2n}$ and $\sigma_{2n+1} = 1$ cannot affect whether any of the other elements in σ match $MMP(1, 0, \emptyset, 0)$. Thus,

$$\sum_{\sigma \in UD_{2n+1}^{(2k+1)}} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} = A_{2n}^{(1,0,\emptyset,0)}(x).$$

Next, consider $UD_{2n+1}^{(2k+1)}$ where $0 \leq k \leq n-1$. Note that there are $\binom{2n}{2k}$ ways to pick the elements which occur to the right of position $2k+1$ in such a σ and there are $C_{2n-2k}(1) = A_{2n-2k}(1)$ ways to order them since the elements to the right of position $2k+1$ form a down-up permutation of length $2n-2k$. That is, the fact that $\sigma_{2k+1} = 1$ implies that σ_{2k+1} matches $MMP(1, 0, \emptyset, 0)$ in σ and that none of the elements to the right of position $2k+1$ matches $MMP(1, 0, \emptyset, 0)$ in σ . Thus, the contribution of the elements to the right of position $2k+1$ in $\sum_{\sigma \in UD_{2n+1}^{(2k+1)}} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)}$ is $C_{2n-2k}(1) = A_{2n-2k}(1)$ since the elements to the right of position $2k+1$ must form a down-up permutation of length $2n-2k$. As we proved above, the elements to the right of position $2k+1$ have no effect on whether $\sigma_1, \dots, \sigma_{2k}$ can contribute to $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$. It follows that the contribution of the elements to the

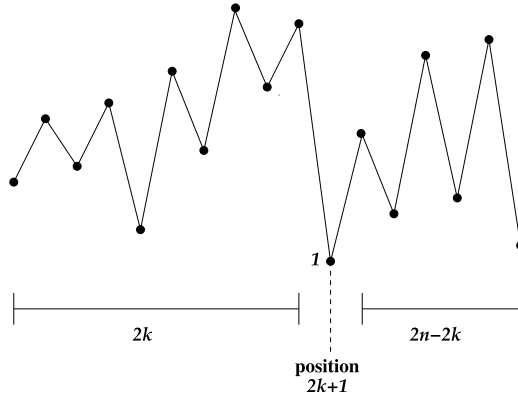


Figure 3: The graph of a $\sigma \in UD_{2n+1}^{(2k+1)}$.

left of position $2k + 1$ in $\sum_{\sigma \in UD_{2n+1}^{(2k+1)}} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)}$ is $A_{2k}^{(1,0,\emptyset,0)}(x)$. It thus follows that for $n \geq 1$,

$$B_{2n+1}^{(1,0,\emptyset,0)}(x) = A_{2n}^{(1,0,\emptyset,0)}(x) + x \sum_{k=0}^{n-1} \binom{2n}{2k} A_{2n-2k}(1) A_{2k}^{(1,0,\emptyset,0)}(x)$$

or, equivalently,

$$(9) \quad \frac{B_{2n+1}^{(1,0,\emptyset,0)}(x)}{(2n)!} = \frac{A_{2n}^{(1,0,\emptyset,0)}(x)}{(2n)!} + x \sum_{k=0}^{n-1} \frac{A_{2n-2k}(1)}{(2n-2k)!} \frac{A_{2k}^{(1,0,\emptyset,0)}(x)}{(2k)!}.$$

Note that $B_1^{(1,0,\emptyset,0)}(x) = 1$. Multiplying both sides of (9) by t^{2n} and summing for $n \geq 1$, we see that

$$\begin{aligned} & \sum_{n \geq 0} \frac{B_{2n+1}^{(1,0,\emptyset,0)}(x) t^{2n}}{(2n)!} \\ &= \sum_{n \geq 0} \frac{A_{2n}^{(1,0,\emptyset,0)}(x) t^{2n}}{(2n)!} + x \left(\sum_{k \geq 1} \frac{A_{2k}(1) t^{2k}}{(2k)!} \right) \left(\sum_{k \geq 0} \frac{A_{2k}^{(1,0,\emptyset,0)}(x) t^{2k}}{(2k)!} \right). \end{aligned}$$

By (1),

$$\sum_{k \geq 0} \frac{A_{2k}(1) t^{2k}}{(2k)!} = \sec(t)$$

so that

$$\frac{\partial}{\partial t} B^{(1,0,\emptyset,0)}(t, x) = (\sec(t))^x + x(\sec(t))^x(\sec(t) - 1).$$

Thus,

$$(10) \quad \frac{\partial}{\partial t} B^{(1,0,\emptyset,0)}(t, x) = (\sec(t))^x(1 - x + x \sec(t)).$$

Our initial condition is that $B^{(1,0,\emptyset,0)}(0, x) = 0$. We used Mathematica to solve this differential equation which gave the following formula for $B^{(1,0,\emptyset,0)}(t, x)$:

$$\begin{aligned} B^{(1,0,\emptyset,0)}(t, x) &= \frac{\sin(t) \cos(t)(1 - x + x \sec(t))}{x + (1 - x) \cos(t)} \\ &\quad \times \left((1 - x) {}_2F_1 \left(\frac{1}{2}, \frac{1+x}{2}; \frac{3}{2}; (\sin(t))^2 \right) \right. \\ &\quad \left. + x {}_2F_1 \left(\frac{1}{2}, \frac{2+x}{2}; \frac{3}{2}; (\sin(t))^2 \right) \right). \end{aligned}$$

2.3. The generating function $D^{(1,0,\emptyset,0)}(t, x)$

If $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}$, then 1 must occur in one of the positions $2, 4, \dots, 2n$. Let $DU_{2n+1}^{(2k)}$ denote the set of permutations $\sigma \in DU_{2n+1}$ such that $\sigma_{2k} = 1$. A schematic diagram of an element in $DU_{2n+1}^{(2k)}$ is pictured in Figure 4.

Let

$$D_{2n+1}^{(1,0,\emptyset,0)}(x, y) = \sum_{\sigma \in DU_{2n+1}} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} y^{\chi(\sigma_1=2n+1)}.$$

First, we want to study the polynomial $\overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x) = D_{2n+1}^{(1,0,\emptyset,0)}(x, x)$. Suppose that $\sigma = \sigma_1 \cdots \sigma_{2n+1}$ is an element of DU_{2n+1} . If $\sigma_1 < 2n+1$, then σ_1 will automatically contribute to $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$. However, if $\sigma_1 = 2n+1$, then σ_1 will not contribute to $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$. Thus, the difference between $\overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x)$ and $D_{2n+1}^{(1,0,\emptyset,0)}(x)$ is that σ_1 always contributes a factor of x to $x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} x^{\chi(\sigma_1=2n+1)}$.

First, we shall prove a simple recursion for $\overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x)$. That is, consider a $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}^{(2k)}$ where $1 \leq k \leq n$. Note that there are $\binom{2n}{2k-1}$ ways to pick the elements which occur to the right of position

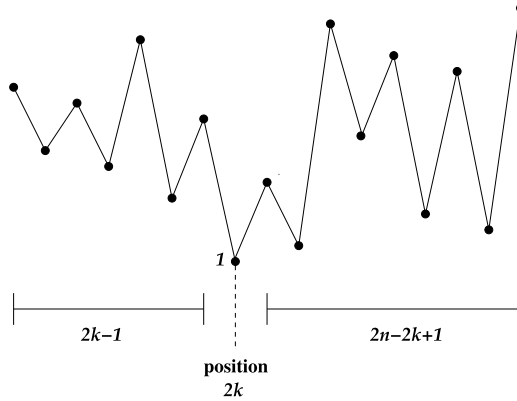


Figure 4: The graph of a $\sigma \in DU_{2n+1}^{(2k)}$.

$2k$ in such a σ and there are $D_{2n-2k+1}(1) = B_{2n-2k-1}(1)$ ways to order them since the elements to the right of position $2k$ form a down-up permutation of length $2n - 2k + 1$. For a $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}^{(2k)}$, none of the elements σ_i for $i > 2k$ matches $MMP(1, 0, \emptyset, 0)$ in σ and $\sigma_{2k} = 1$ always matches $MMP(1, 0, \emptyset, 0)$ in σ . Thus, the only other elements of σ that can possibly contribute to $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)x^{\chi(\sigma_1=2n+1)}$ are the elements $\sigma_1, \sigma_2, \sigma_4, \dots, \sigma_{2k-2}$. Since in $\overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x)$, σ_1 always contributes to $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)x^{\chi(\sigma_1=2n+1)}$ and the elements to the right of position $2k$ have no effect on whether $\sigma_2, \dots, \sigma_{2k-2}$ contribute to $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$, it follows that the contribution of the elements to the left of position $2k$ to

$$\sum_{\sigma \in DU_{2n+1}^{(2k)}} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} x^{\chi(\sigma_1=2n+1)}$$

is $\overline{D}_{2k-1}(x)$. Hence, for $n \geq 1$,

$$\overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x) = x \sum_{k=1}^n \binom{2n}{2k-1} \overline{D}_{2k-1}^{(1,0,\emptyset,0)}(x) B_{2n-2k+1}(1)$$

or, equivalently,

$$(11) \quad \frac{\overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x)}{(2n)!} = x \sum_{k=1}^n \frac{\overline{D}_{2k-1}^{(1,0,\emptyset,0)}(x)}{(2k-1)!} \frac{B_{2n-2k+1}(1)}{(2n-2k+1)!}.$$

Note that $\overline{D}_1^{(1,0,\emptyset,0)}(x) = x$. Multiplying both sides of (11) by t^{2n} and summing for $n \geq 1$, we see that

$$\sum_{n \geq 1} \frac{\overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x)t^{2n}}{(2n)!} - x = x \left(\sum_{k \geq 0} \frac{\overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x)t^{2k}}{(2k)!} \right) \left(\sum_{k \geq 0} \frac{B_{2k+1}^{(1,0,\emptyset,0)}(1)t^{2k+1}}{(2k+1)!} \right).$$

By (2),

$$\sum_{k \geq 1} \frac{B_{2k-1}(1)t^{2k-1}}{(2k-1)!} = \tan(t)$$

so that

$$\frac{\partial}{\partial t} \overline{D}^{(1,0,\emptyset,0)}(t, x) = x + x \tan(x) \overline{D}^{(1,0,\emptyset,0)}(t, x).$$

Our initial condition is that $\overline{D}^{(1,0,\emptyset,0)}(0, x) = 0$. One can easily check that the solution to this differential equation is

$$(12) \quad \overline{D}^{(1,0,\emptyset,0)}(t, x) = x(\sec(t))^x \int_0^t (\cos(z))^x dz.$$

As observed above, the difference between $D_{2n+1}^{(1,0,\emptyset,0)}(x)$ and $\overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x)$ is that the permutations $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}$ such that $\sigma_1 = 2n+1$ are weighted differently in that such permutations are weighted with an extra power of x in $\overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x)$ than they are in $D_{2n+1}^{(1,0,\emptyset,0)}(x)$. That is,

$$x \sum_{\sigma \in DU_{2n+1}, \sigma_1=2n+1} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} = \sum_{\sigma \in DU_{2n+1}, \sigma_1=2n+1} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} x^{\chi(\sigma_1=2n+1)}.$$

It is easy to see that

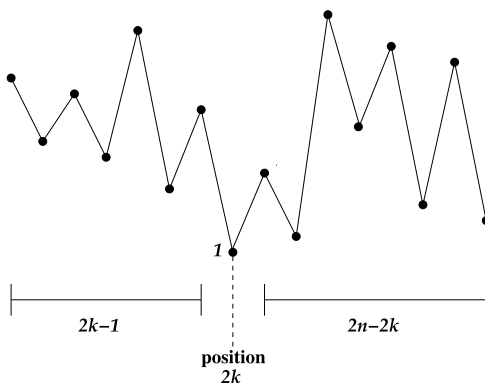
$$\sum_{\sigma \in DU_{2n+1}, \sigma_1=2n+1} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} = A_{2n}^{(1,0,\emptyset,0)}(x).$$

Thus, it follows that

$$(13) \quad D_{2n+1}^{(1,0,\emptyset,0)}(x) = \overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x) + (1-x)A_{2n}^{(1,0,\emptyset,0)}(x).$$

Multiplying both sides of (13) by $\frac{t^{2n+1}}{(2n+1)!}$ and summing for $n \geq 0$, we see that

$$D^{(1,0,\emptyset,0)}(t, x) = \overline{D}^{(1,0,\emptyset,0)}(t, x) + (1-x) \int_0^t A^{(1,0,\emptyset,0)}(z, x) dz.$$

Figure 5: The graph of a $\sigma \in DU_{2n}^{(2k)}$.

Hence,

$$D^{(1,0,\emptyset,0)}(t, x) = x(\sec(t))^x \int_0^t (\cos(z))^x dz + (1-x) \int_0^t (\sec(z))^x dz.$$

2.4. The generating function $C^{(1,0,\emptyset,0)}(t, x)$

If $\sigma = \sigma_1 \cdots \sigma_{2n} \in DU_{2n}$, then 1 must occur in one of the positions $2, 4, \dots, 2n$. Let $DU_{2n}^{(2k)}$ denote the set of permutations $\sigma \in DU_{2n}$ such that $\sigma_{2k} = 1$. A schematic diagram of an element in $DU_{2n}^{(2k)}$ is pictured in Figure 5.

Let

$$C_{2n}^{(1,0,\emptyset,0)}(x, y) = \sum_{\sigma \in DU_{2n}} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} y^{\chi(\sigma_1=2n)}.$$

First, we want to study the polynomial $\overline{C}_{2n}^{(1,0,\emptyset,0)}(x) = C_{2n+1}^{(1,0,\emptyset,0)}(x, x)$. As was the case with $D_{2n+1}^{(1,0,\emptyset,0)}(x)$, if $\sigma = \sigma_1 \cdots \sigma_{2n} \in DU_{2n}$ and $\sigma_1 < 2n$, then σ_1 will automatically contribute to $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$. However, if $\sigma_1 = 2n$, then σ_1 will not contribute to $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$. Thus, the difference between $\overline{C}_{2n}^{(1,0,\emptyset,0)}(x)$ and $C_{2n}^{(1,0,\emptyset,0)}(x)$ is that σ_1 always contributes a factor of x to $x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)}$.

First, consider $\sigma = \sigma_1 \cdots \sigma_{2n} \in DU_{2n}^{(2n)}$. Since $\sigma_{2n} = 1$, it is easy to see that

$$\sum_{\sigma \in DU_{2n}^{(2n)}} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} x^{\chi(\sigma_1=2n)} = \overline{D}_{2n-1}^{(1,0,\emptyset,0)}(x).$$

Next, consider a $\sigma = \sigma_1 \cdots \sigma_{2n} \in DU_{2n}^{(2k)}$ where $1 \leq k < n$. Note that there are $\binom{2n-1}{2k-1}$ ways to pick the elements which occur to the right of position $2k$ in such a σ and there are $C_{2n-2k}(1) = A_{2n-2k}(1)$ ways to order them since the elements to the right of position $2k$ form a down-up permutation of length $2n - 2k$. For a $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n}^{(2k)}$, none of the elements σ_i for $i > 2k$ matches $MMP(1, 0, \emptyset, 0)$ in σ and $\sigma_{2k} = 1$ always matches $MMP(1, 0, \emptyset, 0)$ in σ . Thus, the only other elements of σ that can possibly contribute to $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)x^{\chi(\sigma_1=2n)}$ are the elements $\sigma_1, \sigma_2, \sigma_4, \dots, \sigma_{2k-2}$. Since in $\overline{C}_{2n+1}^{(1,0,\emptyset,0)}(x)$, σ_1 always contributes to $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)x^{\chi(\sigma_1=2n)}$ and the elements to the right of position $2k$ have no effect on whether $\sigma_2, \dots, \sigma_{2k-2}$ contribute to $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$, it follows that the contribution of the elements to the left of position $2k$ to $\sum_{\sigma \in DU_{2n}^{(2k)}} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} x^{\chi(\sigma_1=2n)}$ is $\overline{D}_{2k-1}(x)$. Hence, for $n \geq 1$,

$$\overline{C}_{2n}^{(1,0,\emptyset,0)}(x) = \overline{D}_{2n-1}^{(1,0,\emptyset,0)}(x) + x \sum_{k=1}^{n-1} \binom{2n-1}{2k-1} \overline{D}_{2k-1}^{(1,0,\emptyset,0)}(x) A_{2n-2k}(1)$$

or, equivalently,

$$(14) \quad \frac{\overline{C}_{2n}^{(1,0,\emptyset,0)}(x)}{(2n-1)!} = \frac{\overline{D}_{2n-1}^{(1,0,\emptyset,0)}(x)}{(2n-1)!} + x \sum_{k=1}^{n-1} \frac{\overline{D}_{2k-1}^{(1,0,\emptyset,0)}(x)}{(2k-1)!} \frac{A_{2n-2k}(1)}{(2n-2k)!}.$$

Multiplying both sides of (14) by t^{2n-1} and summing for $n \geq 1$, we see that

$$\begin{aligned} \sum_{n \geq 1} \frac{\overline{C}_{2n}^{(1,0,\emptyset,0)}(x) t^{2n-1}}{(2n-1)!} &= \sum_{n \geq 1} \frac{\overline{D}_{2n-1}^{(1,0,\emptyset,0)}(x) t^{2n-1}}{(2n-1)!} \\ &+ x \left(\sum_{k \geq 0} \frac{\overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x) t^{2n+1}}{(2n+1)!} \right) \left(\sum_{k \geq 1} \frac{A_{2k}(1) t^{2k}}{(2k)!} \right). \end{aligned}$$

By (1),

$$\sum_{k \geq 0} \frac{A_{2k}(1) t^{2k}}{(2k)!} = \sec(t)$$

so that

$$\begin{aligned} \frac{\partial}{\partial t} \overline{C}^{(1,0,\emptyset,0)}(t, x) &= \overline{D}^{(1,0,\emptyset,0)}(t, x) + x \overline{D}^{(1,0,\emptyset,0)}(t, x) (\sec(t) - 1) \\ &= \overline{D}^{(1,0,\emptyset,0)}(t, x) (1 - x + x \sec(t)) \\ &= x (\sec(t))^x (1 - x + x \sec(t)) \int_0^t (\cos(z))^x dz. \end{aligned}$$

Our initial condition is that $\overline{C}^{(1,0,\emptyset,0)}(0, x) = 1$. Maple will give a solution to this differential equation, but it is a complicated expression which is not particularly useful for our purposes so that we will simply record the solution to this differential equation as

(15)

$$\overline{C}^{(1,0,\emptyset,0)}(t, x) = 1 + \int_0^t \left[x (\sec(z))^x (1 - x + x \sec(z)) \int_0^z (\cos(y))^x dy \right] dz.$$

As observed above, the difference between $C_{2n}^{(1,0,\emptyset,0)}(x)$ and $\overline{C}_{2n}^{(1,0,\emptyset,0)}(x)$ is that the permutations $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n}$ such that $\sigma_1 = 2n$ are weighted differently in that such permutations are weighted with an extra power of x in $\overline{C}_{2n}^{(1,0,\emptyset,0)}(x)$ than they are in $C_{2n}^{(1,0,\emptyset,0)}(x)$. That is,

$$x \sum_{\sigma \in DU_{2n}, \sigma_1=2n} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} = \sum_{\sigma \in DU_{2n}, \sigma_1=2n} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} x^{\chi(\sigma_1=2n)}.$$

It is easy to see that

$$\sum_{\sigma \in DU_{2n}, \sigma_1=2n} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} = B_{2n-1}^{(1,0,\emptyset,0)}(x).$$

Thus, it follows that

$$(16) \quad C_{2n}^{(1,0,\emptyset,0)}(x) = \overline{C}_{2n}^{(1,0,\emptyset,0)}(x) + (1-x) B_{2n-1}^{(1,0,\emptyset,0)}(x).$$

Multiplying both sides of (16) by $\frac{t^{2n}}{(2n)!}$ and summing for $n \geq 0$, we see that

$$C^{(1,0,\emptyset,0)}(t, x) = \overline{C}^{(1,0,\emptyset,0)}(t, x) + (1-x) \int_0^t B^{(1,0,\emptyset,0)}(z, x) dz.$$

Hence,

$$C^{(1,0,\emptyset,0)}(t, x) = 1 + \int_0^t \left[x(\sec(z))^x (1 - x + x \sec(z)) \int_0^z (\cos(y))^x dy \right] dz \\ + (1 - x) \int_0^t B^{(1,0,\emptyset,0)}(z, x) dz.$$

3. The coefficients of the polynomials $A_{2n}^{(1,0,\emptyset,0)}(x)$, $B_{2n+1}^{(1,0,\emptyset,0)}(x)$, $C_{2n}^{(1,0,\emptyset,0)}(x)$, and $D_{2n+1}^{(1,0,\emptyset,0)}(x)$.

The main goal of this section is to explain several of the coefficients of the polynomials $A_{2n}^{(1,0,\emptyset,0)}(x)$, $B_{2n+1}^{(1,0,\emptyset,0)}(x)$, $C_{2n}^{(1,0,\emptyset,0)}(x)$, and $D_{2n+1}^{(1,0,\emptyset,0)}(x)$. For $n \geq 1$, let $(2n)!! = \prod_{i=1}^n 2i$ and $(2n-1)!! = \prod_{i=1}^n (2i-1)$. First, it is easy to understand the coefficients of the lowest power of x in each of these polynomials. That is, we have the following theorem.

Theorem 3.

(1) For all $n \geq 1$,

$$A_{2n}^{(1,0,\emptyset,0)}(x)|_x = B_{2n-1}(1).$$

(2) For all $n \geq 1$,

$$B_{2n+1}^{(1,0,\emptyset,0)}(x)|_x = A_{2n}(1) + B_{2n-1}(1).$$

(3) For all $n \geq 2$,

$$C_{2n}^{(1,0,\emptyset,0)}(x)|_x = A_{2n-2}(1) + B_{2n-3}(1).$$

(4) For all $n \geq 1$,

$$D_{2n+1}^{(1,0,\emptyset,0)}(x)|_x = B_{2n-1}(1).$$

Proof. For (1), note that if $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$ where $n \geq 1$, then σ_1 always matches $MMP(1, 0, \emptyset, 0)$ in σ . Moreover if $\sigma_1 \neq 1$, then $\sigma_{2k+1} = 1$ for some $k \geq 1$ in which case σ_{2k+1} will also match $MMP(1, 0, \emptyset, 0)$ in σ . Thus, the only possible way to have $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = 1$ is if $\sigma_1 = 1$ in which case none of $\sigma_2, \dots, \sigma_{2n}$ will match $MMP(1, 0, \emptyset, 0)$ in σ . Clearly, in such a situation, $\text{red}(\sigma_2 \cdots \sigma_{2n}) \in DU_{2n-1}$ so that we have $D_{2n-1}(1) = B_{2n-1}(1)$ ways to choose $\sigma_2 \cdots \sigma_{2n}$. It follows that $A_{2n}^{(1,0,\emptyset,0)}(x)|_x = B_{2n-1}(1)$ for $n \geq 1$.

For (2), note that if $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in UD_{2n+1}$ where $n \geq 1$, then again σ_1 always matches $MMP(1, 0, \emptyset, 0)$ in σ . However, in this case, if

$\sigma_1 \neq 1$, then $\sigma_{2k+1} = 1$ for some $k \geq 1$ in which case σ_{2k+1} will also match $MMP(1, 0, \emptyset, 0)$ in σ if $1 \leq k < n$, but it will not match $MMP(1, 0, \emptyset, 0)$ in σ if $k = n$. Thus, if $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in UD_{2n+1}$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = 1$, it must be the case that $\sigma_1 = 1$ or $\sigma_{2n+1} = 1$. Now if $\sigma_1 = 1$, then none of $\sigma_2, \dots, \sigma_{2n+1}$ will match $MMP(1, 0, \emptyset, 0)$ in σ . Clearly, in such a situation, $\text{red}(\sigma_2 \cdots \sigma_{2n+1}) \in DU_{2n}$ so that we have $C_{2n}(1) = A_{2n}(1)$ ways to choose $\sigma_2 \cdots \sigma_{2n+1}$. If $\sigma_{2n+1} = 1$, then it must be the case that $\sigma_1 = 2$ since otherwise $\sigma_{2k+1} = 2$ for some $k \geq 1$ in which case σ_1 and $\sigma_{2k+1} = 2$ will match $MMP(1, 0, \emptyset, 0)$ in σ . But then $\text{red}(\sigma_2 \cdots \sigma_{2n}) \in DU_{2n-1}$ so that we have $D_{2n-1}(1) = B_{2n-1}(1)$ ways to choose $\sigma_2 \cdots \sigma_{2n}$. Hence, $B_{2n+1}^{(1,0,\emptyset,0)}(x)|_x = A_{2n}(1) + B_{2n-1}(1)$ for $n \geq 1$.

For (3), note that if $\sigma = \sigma_1 \cdots \sigma_{2n} \in DU_{2n}$ where $n \geq 2$, then σ_2 always matches $MMP(1, 0, \emptyset, 0)$ in σ . Moreover, if $\sigma_2 \neq 1$, then $\sigma_{2k} = 1$ for some $k \geq 2$ in which case σ_{2k} will also match $MMP(1, 0, \emptyset, 0)$ in σ for $1 < k < n$ but will not match $MMP(1, 0, \emptyset, 0)$ in σ if $k = n$. In addition, σ_1 will match $MMP(1, 0, \emptyset, 0)$ in σ unless $\sigma_1 = 2n$. Thus, if $\sigma = \sigma_1 \cdots \sigma_{2n} \in DU_{2n}$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = 1$, then we must have $\sigma_1 = 2n$ and either $\sigma_2 = 1$ or $\sigma_{2n} = 1$. If $\sigma_2 = 1$, then none of $\sigma_3, \dots, \sigma_{2n}$ will match $MMP(1, 0, \emptyset, 0)$ in σ . Clearly, in such a situation, $\text{red}(\sigma_3 \cdots \sigma_{2n}) \in DU_{2n-2}$ so that we have $C_{2n-2}(1) = A_{2n-2}(1)$ ways to choose $\sigma_3 \cdots \sigma_{2n}$. If $\sigma_{2n} = 1$, then it must be the case that $\sigma_2 = 2$ since otherwise $\sigma_{2k} = 2$ for some $k \geq 2$ in which case σ_2 and $\sigma_{2k} = 2$ will match $MMP(1, 0, \emptyset, 0)$ in σ . But then $\text{red}(\sigma_3 \cdots \sigma_{2n-1}) \in DU_{2n-3}$ so that we have $D_{2n-3}(1) = B_{2n-3}(1)$ ways to choose $\sigma_3 \cdots \sigma_{2n-1}$. Hence, $C_{2n}^{(1,0,\emptyset,0)}(x)|_x = A_{2n-2}(1) + B_{2n-3}(1)$ for $n \geq 2$.

For (4), note that if $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}$ where $n \geq 1$, then σ_2 always matches $MMP(1, 0, \emptyset, 0)$ in σ . Moreover if $\sigma_2 \neq 1$, then $\sigma_{2k} = 1$ for some $k \geq 2$ in which case σ_{2k} will also match $MMP(1, 0, \emptyset, 0)$ in σ . Finally, σ_1 will also match $MMP(1, 0, \emptyset, 0)$ in σ unless $\sigma_1 = 2n + 1$. Thus, if $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = 1$, it must be the case that $\sigma_1 = 2n + 1$ and $\sigma_2 = 1$ in which case none of $\sigma_3, \dots, \sigma_{2n+1}$ will match $MMP(1, 0, \emptyset, 0)$ in σ . Clearly, in such a situation, $\text{red}(\sigma_3 \cdots \sigma_{2n}) \in DU_{2n-1}$ so that we have $D_{2n-1}(1) = B_{2n-1}(1)$ ways to choose $\sigma_3 \cdots \sigma_{2n+1}$. It follows that $D_{2n+1}^{(1,0,\emptyset,0)}(x)|_x = B_{2n-1}(1)$ for $n \geq 1$. \square

We can also explain the coefficients of the highest power of x in each of the polynomials $A_{2n}(x)$, $B_{2n+1}(x)$, and $D_{2n+1}(x)$. That is, we have the following theorem.

Theorem 4.

- (1) For all $n \geq 1$, the highest power of x that appears in $A_{2n}^{(1,0,\emptyset,0)}(x)$ is x^n which appears with coefficient $(2n - 1)!!$.

- (2) For all $n \geq 1$, the highest power of x that appears in $B_{2n+1}^{(1,0,\emptyset,0)}(x)$ is x^n which appears with coefficient $(n+1)((2n-1)!!)$.
- (3) For all $n \geq 2$, the highest power of x that appears in $C_{2n}^{(1,0,\emptyset,0)}(x)$ is x^n which appears with coefficient $(2n^2 - n - 1)((2n-4)!!) - n((2n-3)!!)$.
- (4) For all $n \geq 2$, the highest power of x that appears in $D_{2n+1}^{(1,0,\emptyset,0)}(x)$ is x^{n+1} which appears with coefficient $(2n)!! - (2n-1)!!$.

Proof. For (1), we proceed by induction on n . Clearly, the formula holds for $n = 1$ since $A_2^{(1,0,\emptyset,0)}(x) = x$. Thus, assume that $n > 1$, and that by induction, we know that $A_{2n-2}^{(1,0,\emptyset,0)}(x)|_{x^{n-1}} = (2n-3)!!$. It is easy to see that the maximum that $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$ can be is n since for any $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$ only $\sigma_1, \sigma_3, \dots, \sigma_{2n-1}$ can match $MMP(1, 0, \emptyset, 0)$ in σ . If $\sigma_{2k+1} = 1$ for some $k < n-1$, then $\sigma_{2k+3}, \dots, \sigma_{2n-1}$ will not match $MMP(1, 0, \emptyset, 0)$ in σ . Thus, if $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$ is such that $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = n$, then $\sigma_{2n-1} = 1$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma_1 \cdots \sigma_{2n-2}) = n-1$. We then have $(2n-1)$ ways to choose the value σ_{2n} , and once we have chosen the value of σ_{2n} , we have $(2n-3)!!$ ways to choose $\sigma_1 \cdots \sigma_{2n-2}$. Hence, $A^{(1,0,\emptyset,0)}(x)|_{x^n} = (2n-1)!!$.

For (2), it is easy to see that our formula holds for $n = 1$ and $n = 2$ since $B_3^{(1,0,\emptyset,0)}(x)|_x = 2$ and $B_5^{(1,0,\emptyset,0)}(x)|_{x^2} = 9 = 3(3!!)$. So assume that $n \geq 3$ and suppose that $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in UD_{2n+1}$. Then only $\sigma_1, \sigma_3, \dots, \sigma_{2n-1}$ can match the $MMP(1, 0, \emptyset, 0)$ in σ . Thus, the maximum that $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$ can be is n . Note that if $\sigma_{2k+1} = 1$ where $0 \leq k < n-1$, then none of σ_{2j+1} for $j > k$ will match the $MMP(1, 0, \emptyset, 0)$ in σ . It follows that if $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = n$, then it must be the case that $\sigma_{2n+1} = 1$ or $\sigma_{2n-1} = 1$. Now if $\sigma_{2n-1} = 1$, then we have $\binom{2n}{2}$ ways to choose the values of σ_{2n} and σ_{2n+1} , and it must be the case that $\text{red}(\sigma_1 \cdots \sigma_{2n-2}) = \tau$ where $\tau \in UD_{2n-2}$ and $\text{mmp}^{(1,0,\emptyset,0)}(\tau) = n-1$. It then follows from part (1) that we have $(2n-3)!!$ ways to choose $\sigma_1 \cdots \sigma_{2n-2}$ so that the set of permutations $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in UD_{2n+1}$ with $\sigma_{2n-1} = 1$ contributes $\binom{2n}{2}(2n-3)!! = n((2n-1)!!)$ to $B_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^n}$. If $\sigma_{2n+1} = 1$, then it must be the case that $\sigma_{2n-1} = 2$ and $\text{red}(\sigma_1 \cdots \sigma_{2n-2}) = \tau$ where $\tau \in UD_{2n-2}$ and $\text{mmp}^{(1,0,\emptyset,0)}(\tau) = n-1$. Thus, we have $2n-1$ choices for the value of σ_{2n} , and then as before we have $(2n-3)!!$ to choose $\sigma_1 \cdots \sigma_{2n-2}$. Thus, the set of permutations $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in UD_{2n+1}$ with $\sigma_{2n+1} = 1$ contributes $(2n-1)!!$ to $B_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^n}$. Hence, $B_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^n} = (n+1)(2n-1)!!$.

For (4), it is easy to see that if $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}$, then only $\sigma_1, \sigma_2, \sigma_4, \dots, \sigma_{2n}$ can match $MMP(1, 0, \emptyset, 0)$ in σ . Thus, $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$ is at most $n+1$. It is also easy to see that if $\sigma_{2k} = 1$ for $k < n$, then $\sigma_{2k+2}, \dots, \sigma_{2n}$ will not match $MMP(1, 0, \emptyset, 0)$ in σ so that if $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) =$

$n + 1$, then it must be the case that $\sigma_{2n} = 1$. Thus, assume that $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}$ is such that $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = n + 1$. Since $\sigma_{2n} = 1$, we have two cases.

Case 1. $\sigma_{2n+1} = 2n + 1$. In this case, we know that σ_1 will always match $MMP(1, 0, \emptyset, 0)$ in σ . Thus, as far as $\sigma_1 \cdots \sigma_{2n-1}$ is concerned, we are dealing with the polynomial $\overline{D}_{2n-1}^{(1,0,\emptyset,0)}(x)$. We then have the following lemma.

Lemma 5. *For $n \geq 1$, the highest power of x which occurs in $\overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x)$ is x^{n+1} which occurs with a coefficient of $(2n)!!$.*

Proof. We proceed by induction on n . Our theorem holds for $n = 1$ since $\overline{D}_3^{(1,0,\emptyset,0)}(x) = 2x^2$. Now assume that $n > 1$ and the lemma holds for $n - 1$. As in our discussion for $D_{2n+1}(x)$, if $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}$ is such that $\chi(\sigma_1 = 2n + 1) + \text{mmp}^{(1,0,\emptyset,0)}(\sigma) = n + 1$, then it must be the case that $\sigma_{2n} = 1$. But then we have $2n$ choices for σ_{2n+1} , and once we have chosen σ_{2n+1} , then $\tau = \text{red}(\sigma_1 \cdots \sigma_{2n-1})$ must be an element of DU_{2n-1} such that $\chi(\tau_1 = 2n - 1) + \text{mmp}^{(1,0,\emptyset,0)}(\tau) = n$. By induction, we have $(2(n - 1))!!$ ways to pick $\sigma_1 \cdots \sigma_{2n-1}$. Thus, $\overline{D}_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^n} = (2n)!!$. \square

It follows that in Case 1, we have $(2n - 2)!!$ ways to pick $\sigma_1 \cdots \sigma_{2n-1}$ so that the permutations such that $\sigma_{2n} = 1$ and $\sigma_{2n+1} = 2n + 1$ contribute $(2n - 2)!!$ to $D_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^{n+1}}$.

Case 2. $\sigma_{2n+1} < 2n + 1$. In this case, $\tau = \text{red}(\sigma_1 \cdots \sigma_{2n-1})$ must be an element of DU_{2n-1} such that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) = n$. It then follows by induction that we have $(2n - 1)$ ways to pick σ_{2n+1} , and once we have chosen σ_{2n+1} , we have $(2n - 2)!! - (2n - 3)!!$ ways to pick $\sigma_1 \cdots \sigma_{2n-1}$. Hence, the permutations such that $\sigma_{2n} = 1$ and $\sigma_{2n+1} < 2n + 1$ contribute $(2n - 1)((2n - 2)!! - (2n - 3)!!) = (2n - 1)((2n - 2)!!) - (2n - 1)!!$ to $D_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^{n+1}}$.

Thus,

$$\begin{aligned} D_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^{n+1}} &= (2n - 2)!! + (2n - 1)((2n - 2)!!) - (2n - 1)!! \\ &= (2n)!! - (2n - 1)!! \end{aligned}$$

For (3), observe that if $\sigma = \sigma_1 \cdots \sigma_{2n} \in DU_{2n}$, then only $\sigma_1, \sigma_2, \sigma_4, \dots, \sigma_{2n-2}$ can match $MMP(1, 0, \emptyset, 0)$ so that $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$ is at most n . It is also easy to see that if $\sigma_{2k} = 1$ for $k < n - 1$, then $\sigma_{2k+2}, \dots, \sigma_{2n-2}$ will not match $MMP(1, 0, \emptyset, 0)$ in σ so that if $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = n$, then it must be the case that $\sigma_{2n} = 1$ or $\sigma_{2n-2} = 1$. Suppose that $\sigma = \sigma_1 \cdots \sigma_{2n} \in DU_{2n}$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = n$. We then have three cases.

Case I. $\sigma_{2n} = 1$. In this case, it must be that $\tau = \text{red}(\sigma_1 \cdots \sigma_{2n-1}) \in DU_{2n-1}$ and $\text{mmp}^{(1,0,\emptyset,0)}(\tau) = n$. Thus, by part (3), we have $(2n-2)!! - (2n-3)!!$ choices for $\sigma_1 \cdots \sigma_{2n}$.

Case II. $\sigma_{2n-2} = 1$ and $\sigma_{2n-1} = 2n$. In this case, we have $(2n-2)$ choices for σ_{2n} . The fact that $\sigma_{2n-1} = 2n$ implies that σ_1 will always match $MMP(1, 0, \emptyset, 0)$ so that $\gamma = \text{red}(\sigma_1 \cdots \sigma_{2n-3})$ is a permutation in DU_{2n-3} such that $\chi(\sigma_1 = 2n-3) + \text{mmp}^{(1,0,\emptyset,0)}(\gamma) = n-1$. By Lemma 5, we will have $(2n-4)!!$ choices for $\sigma_1 \cdots \sigma_{2n-3}$ once we have chosen σ_{2n} . Thus, the permutations in Case II will contribute $(2n-2)!!$ to $C_{2n}^{(1,0,\emptyset,0)}(x)|_{x^n}$.

Case III. $\sigma_{2n-2} = 1$ and $\sigma_{2n-1} < 2n$. In this case, $\tau = \text{red}(\sigma_1 \cdots \sigma_{2n-3})$ must be an element of DU_{2n-3} such that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) = n-1$. Then we have $\binom{2n-2}{2}$ ways to pick σ_{2n-1} and σ_{2n} , and once we have chosen σ_{2n-1} and σ_{2n} , we have $(2n-4)!! - (2n-5)!!$ ways to pick $\sigma_1 \cdots \sigma_{2n-3}$ by part (4). It follows that the permutations in Case III contribute $\binom{2n-2}{2}((2n-4)!! - (2n-5)!!)$ to $C_{2n}^{(1,0,\emptyset,0)}(x)|_{x^n}$.

Thus,

$$\begin{aligned} C_{2n}^{(1,0,\emptyset,0)}(x)|_{x^n} &= (2n-2)!! - (2n-3)!! + (2n-2)!! \\ &\quad + \binom{2n-2}{2}((2n-4)!! - (2n-5)!!) \\ &= 2((2n-2)!!) - (2n-3)!! \\ &\quad + (n-1)(2n-3)((2n-4)!! - (2n-5)!!) \\ &= (2(2n-2) + (n-1)(2n-3))((2n-4)!!) - n((2n-3)!!) \\ &= (2n^2 - n - 1)((2n-4)!!) - n((2n-3)!!). \quad \square \end{aligned}$$

Next, we give formulas for the coefficient of x^2 in the polynomials $A_{2n}^{(1,0,\emptyset,0)}(x)$, $B_{2n+1}^{(1,0,\emptyset,0)}(x)$, $C_{2n}^{(1,0,\emptyset,0)}(x)$, and $B_{2n+1}^{(1,0,\emptyset,0)}(x)$. None of the corresponding sequences had previously appeared in the OEIS [12].

Theorem 6.

(1) For $n \geq 2$,

$$A_{2n}^{(1,0,\emptyset,0)}(x)|_{x^2} = \sum_{k=1}^{n-1} \binom{2n-1}{2k} B_{2k-1}(1) B_{2n-2k-1}(1).$$

(2) For $n \geq 3$,

$$B_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^2} = A_{2n}^{(1,0,\emptyset,0)}(x)|_{x^2} + \sum_{k=1}^{n-1} \binom{2n}{2k} B_{2k-1}(1) A_{2n-2k}(1).$$

(3) For $n \geq 2$,

$$D_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^2} = (2n-1)B_{2n-1}(1) + \sum_{k=2}^{n-1} \binom{2n-1}{2k-2} B_{2k-3}(1) B_{2n-2k+1}(1).$$

(4) For $n \geq 2$,

$$\begin{aligned} C_{2n}^{(1,0,\emptyset,0)}(x)|_{x^2} &= D_{2n-1}^{(1,0,\emptyset,0)}(x)|_{x^2} + (2n-2)A_{2n-2}(1) \\ &\quad + \sum_{k=2}^{n-1} \binom{2n-2}{2k-2} B_{2k-3}(1) A_{2n-2k}(1). \end{aligned}$$

Proof. For (1), suppose that $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = 2$. Then it cannot be that $\sigma_1 = 1$ since that would force that $\sigma_2, \dots, \sigma_{2n}$ do not match $MMP(1, 0, \emptyset, 0)$ in σ . Thus, $1 \in \{\sigma_{2k+1} : k = 1, \dots, n-1\}$. Now suppose that $\sigma_{2k+1} = 1$ where $1 \leq k \leq n-1$. Then σ_{2k+1} will match $MMP(1, 0, \emptyset, 0)$ in σ and $\sigma_{2k+2}, \dots, \sigma_{2n}$ will not match $MMP(1, 0, \emptyset, 0)$ in σ . Hence, it must be the case that $\tau = \text{red}(\sigma_1 \cdots \sigma_{2k})$ is a permutation in UD_{2k} such that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) = 1$. Thus, we have $\binom{2n-1}{2k}$ ways to choose the set of elements for $\sigma_1, \dots, \sigma_{2k}$, and by Theorem 3, we have $B_{2k-1}(1)$ ways to order them. We also have $B_{2n-2k-1}(1)$ ways to order $\sigma_{2k+2} \cdots \sigma_{2n}$. Hence,

$$A_{2n}^{(1,0,\emptyset,0)}(x)|_{x^2} = \sum_{k=1}^{n-1} \binom{2n-1}{2k} B_{2k-1}(1) B_{2n-2k-1}(1).$$

The argument for (2) is similar. That is, suppose $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in UD_{2n+1}$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = 2$. Then again we cannot have $\sigma_1 = 1$. Thus, $1 \in \{\sigma_{2k+1} : k = 1, \dots, n\}$. Now suppose $\sigma_{2k+1} = 1$ where $1 \leq k \leq n-1$. Then σ_{2k+1} will match $MMP(1, 0, \emptyset, 0)$ in σ and $\sigma_{2k+2}, \dots, \sigma_{2n+1}$ will not match $MMP(1, 0, \emptyset, 0)$ in σ . Hence, it must be the case that $\tau = \text{red}(\sigma_1 \cdots \sigma_{2k})$ is a permutation in UD_{2k} such that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) = 1$. Thus, we have $\binom{2n}{2k}$ ways to choose the set of elements for $\sigma_1, \dots, \sigma_{2k}$, and by Theorem 3, we have $B_{2k-1}(1)$ ways to order them. We also have $A_{2n-2k}(1)$ ways to order $\sigma_{2k+2} \cdots \sigma_{2n+1}$. However if $\sigma_{2n+1} = 1$, then σ_{2n+1} does not match

$MMP(1, 0, \emptyset, 0)$ in σ so that it must be the case that $\alpha = \text{red}(\sigma_1 \cdots \sigma_{2n})$ is an element of UD_{2n} such that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) = 2$. It follows that in this case, we have $A_{2n}^{(1,0,\emptyset,0)}(x)|_{x^2}$ ways to choose $\sigma_1 \cdots \sigma_{2n}$. Hence,

$$B_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^2} = A_{2n}^{(1,0,\emptyset,0)}(x)|_{x^2} + \sum_{k=1}^{n-1} \binom{2n}{2k} B_{2k-1}(1) A_{2n-2k}(1).$$

For part (3), suppose that $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = 2$. Then $1 \in \{\sigma_2, \sigma_4, \dots, \sigma_{2n}\}$. Now if $\sigma_2 = 1$, then we cannot have $\sigma_1 = 2n + 1$ because that would force $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = 1$. Thus, if $\sigma_2 = 1$, then $2 \leq \sigma_1 \leq 2n$ in which case σ_1 and σ_2 will be the only two elements of σ to match $MMP(1, 0, \emptyset, 0)$ in σ . We then have $D_{2n-1}(1) = B_{2n-1}(1)$ ways to pick $\sigma_3 \cdots \sigma_{2n+1}$ as $\text{red}(\sigma_3 \cdots \sigma_{2n+1}) \in DU_{2n-1}$. Thus, the number of $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}$ such that $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = 2$ and $\sigma_2 = 1$ is $(2n - 1)B_{2n-1}(1)$. Next, assume that $\sigma_{2k} = 1$ where $2 \leq k \leq n$. Then σ_{2k} matches $MMP(1, 0, \emptyset, 0)$ in σ . It follows that we cannot have $2n + 1 \in \{\sigma_3, \dots, \sigma_{2n+1}\}$ since otherwise σ_1 and σ_2 would also match $MMP(1, 0, \emptyset, 0)$ in σ which would force $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) \geq 3$. Thus, it must be the case that $\sigma_1 = 2n + 1$. Moreover, if $s = \min(\{\sigma_2, \dots, \sigma_{2k-1}\})$, then it must be the case that $\sigma_2 = s$ since otherwise $s = \sigma_{2j}$ for some $2 \leq j \leq k-1$ in which case both σ_2 and σ_{2j} would match $MMP(1, 0, \emptyset, 0)$ in σ which would mean $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) \geq 3$. Thus, we have $\binom{2n-1}{2k-2}$ ways to choose the elements $\sigma_3, \dots, \sigma_{2k-1}$ and then we have $B_{2k-3}(1)$ ways to order $\sigma_3, \dots, \sigma_{2k-1}$ since $\text{red}(\sigma_3 \cdots \sigma_{2k-1})$ must be an element of DU_{2k-3} , and we have $B_{2n-2k+1}(1)$ ways to order $\sigma_{2k+1} \cdots \sigma_{2n+1}$ since $\text{red}(\sigma_{2k+1} \cdots \sigma_{2n+1})$ must be an element of $DU_{2n-2k+1}$. Hence,

$$D_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^2} = (2n - 1)B_{2n-1}(1) + \sum_{k=2}^{n-1} \binom{2n-1}{2k-2} B_{2k-3}(1) B_{2n-2k+1}(1).$$

For part (4), suppose that $\sigma = \sigma_1 \cdots \sigma_{2n} \in DU_{2n}$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = 2$. Then $1 \in \{\sigma_2, \sigma_4, \dots, \sigma_{2n}\}$. We then have three cases.

Case 1. $\sigma_{2n} = 1$. In this case, σ_{2n} does not match $MMP(1, 0, \emptyset, 0)$ in σ so that it must be the case that if $\tau = \text{red}(\sigma_1 \cdots \sigma_{2n-1})$, then τ is a permutation in DU_{2n-1} such that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) = 2$. Thus, we have $D_{2n-1}^{(1,0,\emptyset,0)}(x)|_{x^2}$ choices for $\sigma_1 \cdots \sigma_{2n-1}$.

Case 2. $\sigma_2 = 1$. In this case, we cannot have $\sigma_1 = 2n$ because that would force $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = 1$. Thus, if $\sigma_2 = 1$, then $2 \leq \sigma_1 \leq 2n - 1$

in which case σ_1 and σ_2 will be the only two elements of σ to match $MMP(1, 0, \emptyset, 0)$ in σ . We then have $C_{2n-2}(1) = A_{2n-2}(1)$ ways to pick $\sigma_3 \cdots \sigma_{2n}$ as $\text{red}(\sigma_3 \cdots \sigma_{2n}) \in DU_{2n-2}$. Thus, the number of $\sigma = \sigma_1 \cdots \sigma_{2n} \in DU_{2n}$ such that $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = 2$ and $\sigma_2 = 1$ is $(2n-2)A_{2n-2}(1)$.

Case 3. $\sigma_{2k} = 1$ where $2 \leq k \leq n-1$. Then σ_{2k} matches $MMP(1, 0, \emptyset, 0)$ in σ . It follows that we cannot have $2n+1 \in \{\sigma_3, \dots, \sigma_{2n}\}$ since otherwise σ_1 and σ_2 would also match $MMP(1, 0, \emptyset, 0)$ in σ which would force $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) \geq 3$. Thus, it must be the case that $\sigma_1 = 2n$. Again, if $s = \min(\{\sigma_2, \dots, \sigma_{2k-1}\})$, then it must be the case that $\sigma_2 = s$ since otherwise $s = \sigma_{2j}$ for some $2 \leq j \leq k-1$ in which case both σ_2 and σ_{2j} would match $MMP(1, 0, \emptyset, 0)$ in σ which would force $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) \geq 3$. It follows that we have $\binom{2n-3}{2k-2}$ ways to choose the elements $\sigma_2, \dots, \sigma_{2k-1}$ and then we have $B_{2k-3}(1)$ ways to order $\sigma_3, \dots, \sigma_{2k-1}$ since $\text{red}(\sigma_3 \cdots \sigma_{2k-1})$ must be an element of DU_{2n-3} and $A_{2n-2k}(1)$ ways to order $\sigma_{2k+1} \cdots \sigma_{2n}$ since $\text{red}(\sigma_{2k+1} \cdots \sigma_{2n})$ must be an element of DU_{2n-2k} . Hence, the elements in Case 3 contribute $\sum_{k=2}^{n-1} \binom{2n-3}{2k-2} B_{2k-3}(1) A_{2n-2k}(1)$ to $C_{2n}^{(1,0,\emptyset,0)}(x)|_{x^2}$. Hence,

$$\begin{aligned} C_{2n}^{(1,0,\emptyset,0)}(x)|_{x^2} &= D_{2n-1}^{(1,0,\emptyset,0)}(x)|_{x^2} + (2n-2)A_{2n-2}(1) \\ &\quad + \sum_{k=2}^{n-1} \binom{2n-2}{2k-2} B_{2k-3}(1) A_{2n-2k}(1). \quad \square \end{aligned}$$

Finally, we have the following theorem which gives formulas for the second highest coefficient in $A_{2n}^{(1,0,\emptyset,0)}(x)$, $B_{2n+1}^{(1,0,\emptyset,0)}(x)$, $C_{2n}^{(1,0,\emptyset,0)}(x)$, and $D_{2n+1}^{(1,0,\emptyset,0)}(x)$. None of the corresponding sequences had previously appeared in the OEIS [12].

Theorem 7. (1) For all $n \geq 2$,

$$(17) \quad A_{2n}^{(1,0,\emptyset,0)}(x)|_{x^{n-1}} = \frac{2}{3} \binom{n}{2} ((2n-1)!!).$$

(2) For all $n \geq 2$,

$$(18) \quad B_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^{n-1}} = \left(\frac{7}{3} \binom{n}{2} + 2 \binom{n}{3} \right) ((2n-1)!!).$$

(3) For all $n \geq 1$,

$$(19) \quad D_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^n} = \left(\sum_{k=1}^n \frac{(5k-4)k}{3} ((2k-2)!) \prod_{i=k+1}^n (2i-1) \right) - \frac{2}{3} \left(\binom{n}{2} - 1 \right) ((2n-1)!!).$$

(4) For all $n \geq 3$,

$$(20) \quad C_{2n}^{(1,0,\emptyset,0)}(x)|_{x^{n-1}} = D_{2n-1}^{(1,0,\emptyset,0)}(x)|_{x^{n-1}} + \binom{2n-2}{2} D_{2n-3}^{(1,0,\emptyset,0)}(x)|_{x^{n-2}} + \frac{28n^2 - 72n + 39}{24} ((2n-2)!!) - \frac{5}{3} \binom{n-1}{2} ((2n-3)!!).$$

Proof. For (1), we proceed by induction on n . Now (17) holds for $n = 2$ since $A_4^{(1,0,\emptyset,0)}(x)|_x = 2$. Now suppose that $n > 2$, $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = n - 1$. Only $\sigma_1, \sigma_3, \dots, \sigma_{2n-1}$ can match $MMP(1, 0, \emptyset, 0)$ in σ . Now it cannot be that $\sigma_{2k+1} = 1$ where $k < n - 2$ since then $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) \leq k + 1 < n - 1$. Thus, it must be the case that $\sigma_{2n-3} = 1$ or $\sigma_{2n-1} = 1$. Now if $\sigma_{2n-3} = 1$, then σ_{2n-3} matches $MMP(1, 0, \emptyset, 0)$ in σ and σ_{2n-1} does not match $MMP(1, 0, \emptyset, 0)$ in σ . Then we have $\binom{2n-1}{3}$ ways to choose the values of $\sigma_{2n-2}, \sigma_{2n-1}$, and σ_{2n} , and we have two ways to order them. In addition, we must have that $\text{mmp}^{(1,0,\emptyset,0)}(\text{red}(\sigma_1, \dots, \sigma_{2n-4})) = n - 2$. But then by Theorem 4, we have $(2n-5)!!$ ways to choose $\sigma_1 \cdots \sigma_{2n-4}$ so that the number of $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$ such that $\sigma_{2n-3} = 1$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = n - 1$ is

$$2 \binom{2n-1}{3} (2n-5)!! = \frac{2}{3} (n-1) ((2n-1)!!).$$

Now if $\sigma_{2n-1} = 1$, then we have $2n - 1$ ways to pick the value of σ_{2n} and we must have that $\text{mmp}^{(1,0,\emptyset,0)}(\text{red}(\sigma_1, \dots, \sigma_{2n-2})) = n - 2$. Thus, we have $A_{2(n-1)}(x)|_{x^{n-2}} = \frac{2}{3} \binom{n-1}{2} (2n-3)!!$ ways to choose $\sigma_1 \cdots \sigma_{2n-2}$. Thus, the number of $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$ such that $\sigma_{2n-1} = 1$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = n - 1$ is $\frac{2}{3} \binom{n-1}{2} (2n-1)!!$. Hence,

$$A_{2n}(x)|_{x^{n-1}} = \frac{2}{3} (n-1) ((2n-1)!!) + \frac{2}{3} \binom{n-1}{2} ((2n-1)!!) = \frac{2}{3} \binom{n}{2} ((2n-1)!!).$$

Part (2) can be proved by induction in a similar manner. Now (18) holds for $n = 2$ since $B_5^{(1,0,\emptyset,0)}(x)|_x = 7$. Now suppose that $n > 2$, $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in UD_{2n+1}$, and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = n - 1$. Only $\sigma_1, \sigma_3, \dots, \sigma_{2n-1}$ can match $MMP(1, 0, \emptyset, 0)$ in σ . Again, it cannot be that $\sigma_{2k+1} = 1$ where $k < n - 2$ since then $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) \leq k + 1 < n - 1$. Thus, it must be the case that $\sigma_{2n-3} = 1$, $\sigma_{2n-1} = 1$, and $\sigma_{2n+1} = 1$. Thus, we have three cases.

Case A. $\sigma_{2n-3} = 1$. Then σ_{2n-3} matches $MMP(1, 0, \emptyset, 0)$ in σ and σ_{2n-1} does not match $MMP(1, 0, \emptyset, 0)$ in σ . Then we have $\binom{2n}{4}$ ways to choose the values of σ_{2n-2} , σ_{2n-1} , σ_{2n} , and σ_{2n+1} , and we have 5 ways to order them. In addition, we must have that $\text{mmp}^{(1,0,\emptyset,0)}(\text{red}(\sigma_1, \dots, \sigma_{2n-4})) = n - 2$. But then by Theorem 4, we have $(2n - 5)!!$ ways to pick $\sigma_1 \cdots \sigma_{2n-4}$ so that the number of $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$ such that $\sigma_{2n-3} = 1$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = n - 1$ is

$$5 \binom{2n}{4} (2n - 5)!! = \frac{5}{6} n(n - 1) ((2n - 1)!!) = \frac{5}{3} \binom{n}{2} ((2n - 1)!!).$$

Case B. $\sigma_{2n-1} = 1$. Then we have $\binom{2n}{2}$ ways to pick the values of σ_{2n} , and σ_{2n+1} and we must have that $\text{mmp}^{(1,0,\emptyset,0)}(\text{red}(\sigma_1, \dots, \sigma_{2n-2})) = n - 2$. Thus, we have $A_{2(n-1)}(x)|_{x^{n-2}} = \frac{2}{3} \binom{n-1}{2} (2n - 3)!!$ ways to pick $\sigma_1 \cdots \sigma_{2n-2}$. Thus, number of $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$ such that $\sigma_{2n-1} = 1$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = n - 1$ is

$$\binom{2n}{2} \frac{2}{3} \binom{n-1}{2} ((2n - 3)!!) = \frac{2}{3} n \binom{n-1}{2} ((2n - 1)!!) = 2 \binom{n}{3} ((2n - 1)!!).$$

Case C. $\sigma_{2n+1} = 1$. In this case, σ_{2n+1} does not match $MMP(1, 0, \emptyset, 0)$ in σ so that we must have that $\text{mmp}^{(1,0,\emptyset,0)}(\text{red}(\sigma_1, \dots, \sigma_{2n})) = n - 1$. By part (1), we have $\frac{2}{3} \binom{n}{2} ((2n - 1)!!)$ ways to choose $\sigma_1 \cdots \sigma_{2n}$ in this case.

Thus, it follows that $B_{2n+1}(x)|_{x^{n-1}} = (\frac{7}{3} \binom{n}{2} + 2 \binom{n}{3}) ((2n - 1)!!)$.

Before we can prove part (3), we first need to establish the following lemma.

Lemma 8. For $n \geq 1$,

$$(21) \quad \overline{D}_{2n+1}(x)|_{x^n} = \frac{1}{3} (n^2 - 1) ((2n)!!).$$

Proof. We proceed by induction on n . The lemma holds for $n = 1$ since $\overline{D}_3^{(1,0,\emptyset,0)}(x) = 2x^2$ so that $\overline{D}_3(x)|_x = 0$. Now assume that $n \geq 2$ and $\sigma =$

$\sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}$, and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) + \chi(\sigma_1 = 2n + 1) = n$. Only $\sigma_1, \sigma_2, \sigma_4, \dots, \sigma_{2n}$ can match $MMP(1, 0, \emptyset, 0)$ in σ . Now it cannot be that $\sigma_{2k} = 1$ where $k \leq n - 2$ since then $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) + \chi(\sigma_1 = 2n + 1) \leq k + 1 < n$. Thus, it must be the case that $\sigma_{2n-2} = 1$ or $\sigma_{2n} = 1$. Now if $\sigma_{2n-2} = 1$, then σ_{2n-2} matches $MMP(1, 0, \emptyset, 0)$ in σ and σ_{2n} does not match $MMP(1, 0, \emptyset, 0)$ in σ . Then we have $\binom{2n}{3}$ ways to choose the values of $\sigma_{2n-1}, \sigma_{2n}$, and σ_{2n+1} , and we have two ways to order them. In addition, we must have that if $\tau = \text{red}(\sigma_1, \dots, \sigma_{2n-3})$, then $\text{mmp}^{(1,0,\emptyset,0)}(\tau) + \chi(\tau_1 = 2n - 3) = n - 1$. By Lemma 5, we then have $(2n - 4)!!$ ways to choose $\sigma_1 \cdots \sigma_{2n-3}$ so that the number of $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}$ such that $\sigma_{2n-2} = 1$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) + \chi(\sigma_1 = 2n + 1) = n$ is

$$2 \binom{2n}{3} (2n - 4)!! = \frac{2n - 1}{3} ((2n)!!).$$

Now if $\sigma_{2n} = 1$, then we have $2n$ ways to pick the value of σ_{2n+1} , and if $\alpha = \text{red}(\sigma_1, \dots, \sigma_{2n-1})$, then $\text{mmp}^{(1,0,\emptyset,0)}(\alpha) + \chi(\alpha_1 = 2n - 1) = n - 1$. Then we have $\overline{D}_{2(n-1)+1}(x)|_{x^{n-1}} = \frac{1}{3}((n-1)^2 - 1)((2n-2)!!)$ ways to choose $\sigma_1 \cdots \sigma_{2n-1}$. Thus, number of $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$ such that $\sigma_{2n-1} = 1$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) + \chi(\sigma_1 = 2n + 1) = n$ is $\frac{1}{3}((n-1)^2 - 1)(2n)!!$. Hence,

$$\begin{aligned} \overline{D}_{2n+1}(x)|_{x^n} &= \frac{2n-1}{3}((2n)!!) + \frac{1}{3}((n-1)^2 - 1)((2n)!!) = \frac{1}{3}(n^2 - 1)((2n)!!). \quad \square \end{aligned}$$

We prove part (3) by induction. We have that (19) holds for $n = 2$ since $D_5^{(1,0,\emptyset,0)}(x)|_x = 9$. Now suppose that $n > 2$, $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in DU_{2n+1}$, and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = n$. Only $\sigma_1, \sigma_2, \sigma_4, \dots, \sigma_{2n}$ can match $MMP(1, 0, \emptyset, 0)$ in σ . It cannot be that $\sigma_{2k} = 1$ where $k \leq n - 2$ since then $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) \leq k + 1 < n$. Hence, it must be the case that $\sigma_{2n-2} = 1$ or $\sigma_{2n} = 1$. Thus, we have two cases.

Case I. $\sigma_{2n-2} = 1$. Then σ_{2n-2} matches $MMP(1, 0, \emptyset, 0)$ in σ and σ_{2n} does not match $MMP(1, 0, \emptyset, 0)$ in σ . Then we have two subcases.

Subcase I.a. $2n + 1 \in \{\sigma_{2n-1}, \sigma_{2n}, \sigma_{2n+1}\}$. In this case, we have $\binom{2n-1}{2}$ ways to choose the values of the other 2 elements in the set $\{\sigma_{2n-1}, \sigma_{2n}, \sigma_{2n+1}\}$ and then we have 2 ways to order $\sigma_{2n-1}\sigma_{2n}\sigma_{2n+1}$. Then since $2n + 1 \in \{\sigma_{2n-1}, \sigma_{2n}, \sigma_{2n+1}\}$, we are guaranteed that σ_1 matches $MMP(1, 0, \emptyset, 0)$ in σ . Thus, when we consider $\tau = \text{red}(\sigma_1 \cdots \sigma_{2n-3})$, we must have that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) + \chi(\tau_1 = 2n - 3) = n - 1$. It follows from Lemma 5 that

we have $(2n - 4)!!$ ways to pick $\sigma_1 \cdots \sigma_{2n-3}$. Hence, the $\sigma \in DU_{2n+1}$ in this subcase contribute $2 \binom{2n-1}{2} (2n-4)!! = (2n-1)((2n-2)!!)$ to $D_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^n}$.

Subcase I.b. $2n + 1 \notin \{\sigma_{2n-1}, \sigma_{2n}, \sigma_{2n+1}\}$. We then have $\binom{2n-1}{3}$ ways to choose the values of the elements of $\{\sigma_{2n-1}, \sigma_{2n}, \sigma_{2n+1}\}$ and 2 ways to order them. Because $2n + 1 \notin \{\sigma_{2n-1}, \sigma_{2n}, \sigma_{2n+1}\}$, we are not guaranteed that σ_1 matches $MMP(1, 0, \emptyset, 0)$ in σ . Thus, when we consider $\tau = \text{red}(\sigma_1 \cdots \sigma_{2n-3})$, we must have that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) = n-1$. It follows from Theorem 4 that we have $(2n-4)!! - (2n-5)!!$ ways to pick $\sigma_1 \cdots \sigma_{2n-3}$. Thus, the $\sigma \in DU_{2n+1}$ in this subcase contribute

$$\begin{aligned} & 2 \binom{2n-1}{3} ((2n-4)!! - (2n-5)!!) \\ &= \frac{(2n-1)(2n-3)}{3} ((2n-2)!!) - \frac{2}{3} (n-1)((2n-1)!!) \end{aligned}$$

to $D_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^n}$.

Case II. $\sigma_{2n} = 1$. In this case σ_{2n} matches $MMP(1, 0, \emptyset, 0)$ in σ . Then again, we have two subcases.

Subcase II.a. $\sigma_{2n+1} = 2n + 1$. Because $2n + 1 = \sigma_{2n+1}$, we are guaranteed that σ_1 matches $MMP(1, 0, \emptyset, 0)$ in σ . Thus, when we consider $\tau = \text{red}(\sigma_1 \cdots \sigma_{2n-1})$, we must have that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) + \chi(\tau_1 = 2n-1) = n-1$. It follows from Lemma 8 that we have $\frac{1}{3}((n-1)^2 - 1)((2n-2)!!)$ ways to choose $\sigma_1 \cdots \sigma_{2n-3}$. Hence, the $\sigma \in DU_{2n+1}$ in this subcase contribute $\frac{1}{3}((n-1)^2 - 1)((2n-2)!!)$ to $D_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^n}$.

Subcase II.b. $2n + 1 \neq \sigma_{2n+1}$. We then have $(2n-1)$ ways to choose the value of σ_{2n+1} . Because $2n + 1 \neq \sigma_{2n+1}$, then we are not guaranteed that σ_1 matches $MMP(1, 0, \emptyset, 0)$ in σ . Thus, when we consider $\tau = \text{red}(\sigma_1 \cdots \sigma_{2n-1})$, we must have that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) = n-1$. It then follows by induction that the permutations $\sigma \in DU_{2n+1}$ in this subcase contribute

$$\begin{aligned} & (2n-1) \left(\left(\sum_{k=1}^{n-1} \frac{(5k-4)k}{3} ((2k-2)!!) \prod_{i=k+1}^{n-1} (2i-1) \right) \right. \\ & \quad \left. - \frac{2}{3} \left(\binom{n-1}{2} - 1 \right) ((2n-3)!!) \right) \\ &= \left(\sum_{k=1}^{n-1} \frac{(5k-4)k}{3} ((2k-2)!!) \prod_{i=k+1}^n (2i-1) \right) \end{aligned}$$

$$- \frac{2}{3} \left(\binom{n-1}{2} - 1 \right) ((2n-1)!!)$$

to $D_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^n}$.

It follows that

$$\begin{aligned} D_{2n+1}^{(1,0,\emptyset,0)}(x)|_{x^n} &= (2n-1)((2n-2)!!) + \frac{1}{3}((n-1)^2-1)((2n-2)!!) \\ &\quad + \frac{(2n-1)(2n-3)}{3}((2n-2)!!) - \frac{2n-2}{3}((2n-1)!!) \\ &\quad + \left(\sum_{k=1}^{n-1} \frac{(5k-4)k}{3}((2k-2)!!) \prod_{i=k+1}^n (2i-1) \right) \\ &\quad - \frac{2}{3} \left(\binom{n-1}{2} - 1 \right) ((2n-1)!!) \\ &= ((2n-2)!!) \left((2n-1) + \frac{(2n-1)(2n-3)}{3} + \frac{(n-1)^2-1}{3} \right) \\ &\quad - ((2n-1)!!) \left(\frac{2}{3}(n-1) + \frac{2}{3} \left(\binom{n-1}{2} - 1 \right) \right) \\ &\quad + \left(\sum_{k=1}^{n-1} \frac{(5k-4)k}{3}((2k-2)!!) \prod_{i=k+1}^n (2i-1) \right) \\ &= \frac{(5n-4)n}{3}((2n-2)!!) - \frac{2}{3} \left(\binom{n}{2} - 1 \right) ((2n-1)!!) \\ &\quad + \left(\sum_{k=1}^{n-1} \frac{(5k-4)k}{3}((2k-2)!!) \prod_{i=k+1}^n (2i-1) \right) \\ &= \left(\sum_{k=1}^n \frac{(5k-4)k}{3}((2k-2)!!) \prod_{i=k+1}^n (2i-1) \right) \\ &\quad - \frac{2}{3} \left(\binom{n}{2} - 1 \right) ((2n-1)!!). \end{aligned}$$

For part (4), suppose that $\sigma = \sigma_1 \cdots \sigma_{2n} \in DU_{2n}$. Then only $\sigma_1, \sigma_2, \dots, \sigma_{2n-2}$ can match $MMP(1, 0, \emptyset, 0)$ in σ . Thus, it cannot be the case that $\sigma_{2k} = 1$ where $k < n-2$ since then $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) \leq k+1 < n-1$. Thus, we must have $1 \in \{\sigma_{2n-4}, \sigma_{2n-2}, \sigma_{2n}\}$. We then have three cases.

Case 1. $\sigma_{2n} = 1$. In this case, σ_{2n} does not match $MMP(1, 0, \emptyset, 0)$ in σ . Hence, it must be the case that $\text{mmp}^{(1,0,\emptyset,0)}(\text{red}(\sigma_1 \cdots \sigma_{2n-1})) = n-1$ so that by part (3), we have $D_{2n-1}^{(1,0,\emptyset,0)}(x)|_{x^{n-1}}$ ways to choose $\sigma_1 \cdots \sigma_{2n-1}$.

Case 2. $\sigma_{2n-2} = 1$. In this case, σ_{2n-2} matches $MMP(1, 0, \emptyset, 0)$ in σ . We then have two subcases.

Subcase 2.1. $\sigma_{2n-1} = 2n$. In this case, we are guaranteed that σ_1 will match $MMP(1, 0, \emptyset, 0)$ in σ . Thus, if $\tau = \text{red}(\sigma_1 \cdots \sigma_{2n-3})$, then we must have that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) + \chi(\tau_1 = 2n - 3) = n - 2$. We then have $(2n - 2)$ ways to choose σ_{2n} , and once we have chosen σ_{2n} , we have $\frac{2}{3}((n - 2)^2 - 1)((2n - 4)!!)$ ways to choose $\sigma_1 \cdots \sigma_{2n-3}$ by Lemma 8. Thus, the permutations $\sigma \in DU_{2n}$ in this case contribute $\frac{2}{3}((n - 2)^2 - 1)((2n - 2)!!)$ to $C_{2n}(x)|_{x^{n-1}}$.

Subcase 2.2. $\sigma_{2n-1} \neq 2n$. In this case, we are not guaranteed that σ_1 will match $MMP(1, 0, \emptyset, 0)$ in σ . Thus, if $\tau = \text{red}(\sigma_1 \cdots \sigma_{2n-3})$, then we must have that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) = n - 2$. We then have $\binom{2n-2}{2}$ ways to choose σ_{2n-1} and σ_{2n} , and once we have chosen σ_{2n-1} and σ_{2n} , we have $D_{2n-3}(x)|_{x^{n-2}}$ ways to choose $\sigma_1 \cdots \sigma_{2n-3}$. Thus, the permutations $\sigma \in DU_{2n}$ in this case contribute $\binom{2n-2}{2} D_{2n-3}(x)|_{x^{n-2}}$ to $C_{2n}(x)|_{x^{n-1}}$.

Case 3. $\sigma_{2n-4} = 1$. In this case, σ_{2n-4} matches $MMP(1, 0, \emptyset, 0)$ in σ , but σ_{2n-2} does not match $MMP(1, 0, \emptyset, 0)$ in σ . Again, we have two subcases.

Subcase 3.1. $2n \in \{\sigma_{2n-3}, \sigma_{2n-2}, \sigma_{2n-1}, \sigma_{2n}\}$. In this case, we are guaranteed that σ_1 will match $MMP(1, 0, \emptyset, 0)$ in σ . Thus, if $\tau = \text{red}(\sigma_1 \cdots \sigma_{2n-5})$, then we must have that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) + \chi(\tau_1 = 2n - 5) = n - 2$. We then have $\binom{2n-2}{3}$ ways to choose the remaining elements for $\{\sigma_{2n-3}, \sigma_{2n-2}, \sigma_{2n-1}, \sigma_{2n}\}$. Once we have chosen the remaining elements for $\{\sigma_{2n-3}, \sigma_{2n-2}, \sigma_{2n-1}, \sigma_{2n}\}$, we have 5 ways to order them and we have $\overline{D}_{2n-5}(x)|_{x^{n-2}}$ ways to choose $\sigma_1 \cdots \sigma_{2n-5}$. By Lemma 5, $\overline{D}_{2n-5}(x)|_{x^{n-2}} = (2n - 6)!!$. Thus, the permutations $\sigma \in DU_{2n}$ in this case contribute $5 \binom{2n-2}{3} ((2n - 6)!!) = \frac{5}{6}(2n - 3)((2n - 2)!!)$ to $C_{2n}(x)|_{x^{n-1}}$.

Subcase 3.2. $2n \notin \{\sigma_{2n-3}, \sigma_{2n-2}, \sigma_{2n-1}, \sigma_{2n}\}$. In this case, we are not guaranteed that σ_1 will match $MMP(1, 0, \emptyset, 0)$ in σ . Thus, if $\tau = \text{red}(\sigma_1 \cdots \sigma_{2n-5})$, then we must have that $\text{mmp}^{(1,0,\emptyset,0)}(\tau) = n - 2$. We then have $\binom{2n-2}{4}$ ways to choose the set $\{\sigma_{2n-3}, \sigma_{2n-2}, \sigma_{2n-1}, \sigma_{2n}\}$. Once we have chosen $\{\sigma_{2n-3}, \sigma_{2n-2}, \sigma_{2n-1}, \sigma_{2n}\}$, we have 5 ways to order them and we have $D_{2n-5}(x)|_{x^{n-2}}$ ways to choose $\sigma_1 \cdots \sigma_{2n-5}$. By Theorem 4, $D_{2n-5}(x)|_{x^{n-2}} = (2n - 6)!! - (2n - 7)!!$. Thus, the permutations $\sigma \in DU_{2n}$ in this case contribute

$$\begin{aligned} & 5 \binom{2n-2}{4} ((2n - 6)!! - (2n - 7)!!) \\ & = \frac{5}{24}(2n - 3)(2n - 5)((2n - 2)!!) - \frac{5}{3} \binom{n-1}{2} ((2n - 3)!!) \end{aligned}$$

to $C_{2n}(x)|_{x^{n-1}}$.

It follows that

$$\begin{aligned}
 C_{2n}(x)|_{x^{n-1}} &= D_{2n-1}(x)|_{x^{n-1}} + \binom{2n-2}{2} D_{2n-3}(x)|_{x^{n-2}} \\
 &\quad + \frac{2}{3}((n-2)^2 - 1)((2n-2)!!) + \frac{5}{6}(2n-3)((2n-2)!!) \\
 &\quad \times \frac{5}{24}(2n-3)(2n-5)((2n-2)!!) - \frac{5}{3} \binom{n-1}{2} ((2n-3)!!) \\
 &= D_{2n-1}(x)|_{x^{n-1}} + \binom{2n-2}{2} D_{2n-3}(x)|_{x^{n-2}} \\
 &\quad + \left(\frac{28n^2 - 72n + 39}{24} \right) ((2n-2)!!) - \frac{5}{3} \binom{n-1}{2} ((2n-3)!!).
 \end{aligned}$$

□

4. Conclusions

As pointed out in [8], the simple type of recursions for the distribution of $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$ for σ in UD_n or DU_n proved in this paper no longer holds for the distribution of $\text{mmp}^{(k,0,0,0)}(\sigma)$ for σ in UD_n or DU_n if $k \geq 2$. For example, suppose that we try to develop a recursion for $A_{2n}^{(2,0,\emptyset,0)}(x) = \sum_{\sigma \in UD_{2n}} x^{\text{mmp}^{(2,0,\emptyset,0)}(\sigma)}$. Then if we consider the permutations $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$ such that $\sigma_{2k+1} = 1$, we still have $\binom{2n-1}{2k}$ ways to pick the elements for $\sigma_1 \cdots \sigma_{2k}$. However, in this case the question of whether some σ_i with $i \leq 2k$ matches $MMP(2, 0, \emptyset, 0)$ in σ is dependent on what values occur in $\sigma_{2k+2} \cdots \sigma_{2n}$. For example, if $2n \in \{\sigma_{2k+2}, \dots, \sigma_{2n}\}$, then every σ_i with $i \leq k$ will match $MMP(2, 0, \emptyset, 0)$ in σ . However, if $2n \in \{\sigma_1, \dots, \sigma_{2k-1}\}$, this will not be the case. Thus, we cannot develop a simple recursion for $A_{2n}^{(2,0,\emptyset,0)}(x)$.

However, one can develop recursions similar to the ones used in this paper to study the distribution in up-down and down-up permutations of other quadrant marked meshed patterns $MMP(a, b, c, d)$ in the case where $a, b, c, d \in \{\emptyset, 1\}$. Indeed, in some cases, there are simple relations between such distributions beyond those given in Proposition 1. For example, consider the statistics $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$ and $\text{mmp}^{(0,0,\emptyset,0)}(\sigma)$ over UD_{2n} . Clearly, for any $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$, σ_{2i} can never match $MMP(1, 0, \emptyset, 0)$ or $MMP(0, 0, \emptyset, 0)$ since $(2i-1, \sigma_{2i-1})$ will always be an element of $G(\sigma)$ that lies in the third quadrant with respect to the coordinate system centered

at $(2i, \sigma_{2i})$. On the other hand, elements of the form σ_{2i-1} for $i = 1, \dots, n$, always have an element $G(\sigma)$ in the first quadrant relative to the coordinate system centered at $(2i-1, \sigma_{2i-1})$, namely $(2i, \sigma_{2i})$. Thus, if $\sigma \in UD_{2n}$, then $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = \text{mmp}^{(0,0,\emptyset,0)}(\sigma)$. Therefore, for all $n \geq 1$,

$$A_{2n}^{(1,0,\emptyset,0)}(x) = A_{2n}^{(0,0,\emptyset,0)}(x).$$

It is not true that $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = \text{mmp}^{(0,0,\emptyset,0)}(\sigma)$ for all $\sigma \in UD_{2n+1}$ since if $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in UD_{2n+1}$ and $\sigma_{2n+1} = 1$, then σ_{2n+1} matches $MMP(0, 0, \emptyset, 0)$ in σ but does not match $MMP(1, 0, \emptyset, 0)$ in σ . However, this is the only case where $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$ and $\text{mmp}^{(0,0,\emptyset,0)}(\sigma)$ differ. That is, if $\sigma = \sigma_1 \cdots \sigma_{2n+1} \in UD_{2n+1}$ and $\sigma_{2n+1} \neq 1$, then σ_{2n+1} does not match $MMP(0, 0, \emptyset, 0)$ in σ and we can argue as above that $\text{mmp}^{(1,0,\emptyset,0)}(\sigma) = \text{mmp}^{(0,0,\emptyset,0)}(\sigma)$. However if $\sigma_{2n+1} = 1$, then σ_{2n+1} matches $MMP(0, 0, \emptyset, 0)$ in σ but does not match $MMP(1, 0, \emptyset, 0)$ in σ . Thus, if $\sigma \in UD_{2n+1}^{(2n+1)}$, $1 + \text{mmp}^{(1,0,\emptyset,0)}(\sigma) = \text{mmp}^{(0,0,\emptyset,0)}(\sigma)$. It is easy to see that

$$\sum_{\sigma \in UD_{2n+1}^{(2n+1)}} x^{\text{mmp}^{(1,0,\emptyset,0)}(\sigma)} = A_{2n}^{(1,0,\emptyset,0)}(x)$$

so that for all $n \geq 1$,

$$B_{2n+1}^{(0,0,\emptyset,0)}(x) + (1-x)A_{2n}^{(1,0,\emptyset,0)}(x) = B_{2n+1}^{(1,0,\emptyset,0)}(x).$$

A slightly more subtle relation holds between the distribution of $\text{mmp}^{(1,0,0,0)}(\sigma)$ and $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$ for $\sigma \in UD_{2n}$. For example, in [8], the authors computed the following table for $A_{2n}^{(1,0,0,0)}(x)$.

n	$A_{2n}^{(1,0,0,0)}(x)$
0	1
1	x
2	$x^2(3+2x)$
3	$x^3(15+30x+16x^2)$
4	$x^4(105+420x+588x^2+272x^3)$
5	$x^5(945+6300x+16380x^2+18960x^3+7936x^4)$
6	$x^6(10395+103950x+429660x^2+893640x^3+911328x^4+353792x^5)$

Comparing the tables for $A_{2n}^{(1,0,0,0)}(x)$ and $A_{2n}^{(1,0,\emptyset,0)}(x)$, one is naturally led to conjecture that for all $n \geq 1$ and $1 \leq k \leq n$,

$$(22) \quad A_{2n}^{(1,0,\emptyset,0)}(x)|_{x^k} = A_{2n}^{(1,0,0,0)}(x)|_{x^{2n-k}}.$$

This follows from comparing $\sigma = \sigma_1 \cdots \sigma_{2n} \in UD_{2n}$ with its reverse complement $(\sigma^r)^c = (2n + 1 - \sigma_{2n})(2n + 1 - \sigma_{2n-1}) \cdots (2n + 1 - \sigma_1)$ which is also in UD_{2n} . That is, suppose that σ_i matches $MMP(1, 0, \emptyset, 0)$ in σ . Then i must be odd, i.e. $i = 2k + 1$ for some $0 \leq k \leq n - 1$, and there must be no elements in $\sigma_1 \cdots \sigma_{2k}$ which are less than σ_{2k+1} . This means that in $(\sigma^r)^c$, $(2n + 1 - \sigma_{2k+1})$ has no elements to its right which are greater than $(2n + 1 - \sigma_{2k+1})$ so that $(2n + 1 - \sigma_{2k+1})$ will not match $MMP(1, 0, \emptyset, 0)$ in $(\sigma^r)^c$. Vice versa, if σ_{2k+1} does not match $MMP(1, 0, \emptyset, 0)$ in σ , there is an element in $\sigma_1 \cdots \sigma_{2k}$ which is less than σ_{2k+1} . This means that in $(\sigma^r)^c$, $(2n + 1 - \sigma_{2k+1})$ has an element to its right which is greater than $(2n + 1 - \sigma_{2k+1})$ so that $(2n + 1 - \sigma_{2k+1})$ will match $MMP(1, 0, \emptyset, 0)$ in $(\sigma^r)^c$. Similarly, in σ , none of $\sigma_2, \sigma_4, \dots, \sigma_{2n}$ will match $MMP(1, 0, \emptyset, 0)$ while in $(\sigma^r)^c$, each of $(2n + 1 - \sigma_2), \dots, (2n + 1 - \sigma_{2n})$ will match $MMP(1, 0, \emptyset, 0)$ in $(\sigma^r)^c$. Thus, it follows that for all $\sigma \in UD_{2n}$,

$$n + (n - \text{mmp}^{(1,0,\emptyset,0)}(\sigma)) = \text{mmp}^{(1,0,\emptyset,0)}((\sigma^r)^c).$$

This shows that (22) holds.

There is no such simple relation between the distribution of $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$ and the distribution of $\text{mmp}^{(1,0,\emptyset,0)}(\sigma)$ for UD_{2n+1} , DU_{2n} or DU_{2n+1} as can be seen from the following tables computed in [8].

Based on these tables, we conjectured in [8] that the polynomials $A_{2n}^{(1,0,\emptyset,0)}(x)$, $B_{2n+1}^{(1,0,\emptyset,0)}(x)$, $C_{2n}^{(1,0,\emptyset,0)}(x)$, and $D_{2n+1}^{(1,0,\emptyset,0)}(x)$ are unimodal for all $n \geq 1$. We also conjecture that $A_{2n}^{(1,0,\emptyset,0)}(x)$, $B_{2n+1}^{(1,0,\emptyset,0)}(x)$, $C_{2n}^{(1,0,\emptyset,0)}(x)$, and $D_{2n+1}^{(1,0,\emptyset,0)}(x)$ are unimodal for all $n \geq 1$. Of course, by (22), the conjectures for $A_{2n}^{(1,0,\emptyset,0)}(x)$ and $A_{2n}^{(1,0,\emptyset,0)}(x)$ are equivalent.

Finally, we suggest that it should be interesting to study the distribution of quadrant marked mesh patterns on other classes of pattern-restricted permutations such as *2-stack-sortable permutations* or *vexillary permutations* (see [6] for definitions of these) and many other permutation classes having nice properties.

n	$B_{2n-1}^{(1,0,\emptyset,0)}(x)$
1	1
2	$2x$
3	$8x^2(1 + x)$
4	$16x^3(3 + 8x + 6x^2)$
5	$128x^4(3 + 15x + 27x^2 + 17x^3)$
6	$256x^5(15 + 120x + 381x^2 + 556x^3 + 310x^4)$
7	$1024x^6(45 + 525x + 2562x^2 + 6420x^3 + 8146x^4 + 4146x^5)$

n	$C_{2n}^{(1,0,0,0)}(x)$
0	1
1	1
2	$x(2 + 3x)$
3	$x^2(8 + 28x + 25x^2)$
4	$x^3(48 + 296x + 614x^2 + 427x^3)$
5	$x^4(384 + 3648x + 13104x^2 + 20920x^3 + 12465x^4)$
6	$x^5(3840 + 51840x + 282336x^2 + 769072x^3 + 1039946x^4 + 555731x^5)$

n	$D_{2n-1}^{(1,0,0,0)}(x)$
1	1
2	$x(1 + x)$
3	$x^2(3 + 8x + 5x^2)$
4	$x^3(15 + 75x + 121x^2 + 61x^3)$
5	$x^4(105 + 840x + 2478x^2 + 3128x^3 + 1385x^4)$
6	$x^5(945 + 11025x + 51030x^2 + 115350x^3 + 124921x^4 + 50521x^5)$
7	$x^6(10395 + 166320x + 1105335x^2 + 3859680x^3 + 7365633x^4 + 7158128x^5 + 2702765x^6)$

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