

## CRITICAL SETS OF RANDOM SMOOTH FUNCTIONS ON COMPACT MANIFOLDS\*

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**Abstract.** Given a compact, connected Riemann manifold without boundary  $(M, g)$  of dimension  $m$  and a large positive constant  $L$  we denote by  $\mathbf{U}_L$  the subspace of  $C^\infty(M)$  spanned by eigenfunctions of the Laplacian corresponding to eigenvalues  $\leq L$ . We equip  $\mathbf{U}_L$  with the standard Gaussian probability measure induced by the  $L^2$ -metric on  $\mathbf{U}_L$ , and we denote by  $\mathcal{N}_L$  the expected number of critical points of a random function in  $\mathbf{U}_L$ . We prove that  $\mathcal{N}_L \sim C_m \dim \mathbf{U}_L$  as  $L \rightarrow \infty$ , where  $C_m$  is an explicit positive constant that depends only on the dimension  $m$  of  $M$  and satisfying the asymptotic estimate  $\log C_m \sim \frac{m}{2} \log m$  as  $m \rightarrow \infty$ .

**Key words.** Random Morse functions, critical points, Kac-Rice formula, gaussian random processes, spectral function, random matrices.

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**1. Introduction.** Suppose that  $(M, g)$  is a smooth, compact, connected Riemann manifold of dimension  $m > 1$ . We denote by  $|dV_g|$  the volume density on  $M$  induced by  $g$ . Throughout the paper we assume that the volume is normalized

$$\int_M |dV_g(x)| = 1.$$

For any  $\mathbf{u}, \mathbf{v} \in C^\infty(M)$  we denote by  $(\mathbf{u}, \mathbf{v})_g$  their  $L^2$  inner product,

$$(\mathbf{u}, \mathbf{v})_g := \int_M \mathbf{u}(x)\mathbf{v}(x) |dV_g(x)|.$$

The  $L^2$ -norm of a smooth function  $u$  is then

$$\|\mathbf{u}\| := \sqrt{(\mathbf{u}, \mathbf{u})_g}.$$

Let  $\Delta_g : C^\infty(M) \rightarrow C^\infty(M)$  denote the scalar Laplacian defined by the metric  $g$ . For  $L > 0$  we set

$$\mathbf{U}_L = \mathbf{U}_L(M, g) := \bigoplus_{\lambda \in [0, L]} \ker(\lambda - \Delta_g), \quad d(L) := \dim \mathbf{U}_L.$$

We equip  $\mathbf{U}_L$  with the Gaussian probability measure.

$$d\gamma_L(\mathbf{u}) := (2\pi)^{-\frac{d(L)}{2}} e^{-\frac{\|\mathbf{u}\|^2}{2}} |d\mathbf{u}|.$$

For any  $\mathbf{u} \in \mathbf{U}_L$  we denote by  $\mathcal{N}_L(\mathbf{u})$  the number of critical points of  $\mathbf{u}$ . If  $L$  is sufficiently large, then  $\mathcal{N}_L(\mathbf{u})$  is finite with probability 1. We obtain in this fashion a random variable  $\mathcal{N}_L = \mathcal{N}_{L, M, g}$ , and we denote by  $\mathbf{E}(\mathcal{N}_L)$  its expectation

$$\mathbf{E}(\mathcal{N}_L) := \int_{\mathbf{U}_L} \mathcal{N}_L(\mathbf{u}) d\gamma_L(\mathbf{u}).$$

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In this paper we investigate the behavior of  $\mathbf{E}(\mathcal{N}_L)$  as  $L \rightarrow \infty$ . More precisely, we will prove the following result.

**THEOREM 1.1.** *For any  $m > 1$  there exists a positive constant  $C = C(m)$  such that **for any** compact, connected,  $m$ -dimensional Riemannian manifold  $M$  we have*

$$(1.1) \quad \mathbf{E}(\mathcal{N}_{L,M,g}) \sim C(m) \dim \mathbf{U}_L(M, g) \text{ as } L \rightarrow \infty.$$

The constant  $C(m)$  can be expressed in terms of certain statistics on the space  $\mathcal{S}_m$  the space of symmetric  $m \times m$  matrices . We denote  $d\gamma_*$  the Gaussian measure<sup>1</sup> on  $\mathcal{S}_m$  given by

$$(1.2) \quad d\gamma_*(X) = \frac{1}{(2\pi)^{\frac{m(m+1)}{4}} \sqrt{\mu_m}} \cdot e^{-\frac{1}{4} \left( \text{tr } X^2 - \frac{1}{m+2} (\text{tr } X)^2 \right)} 2^{\frac{1}{2} \binom{m}{2}} \prod_{i \leq j} dx_{ij},$$

$$\mu_m = 2^{\binom{m}{2} + 1} (m + 2)^{m-1}.$$

Then

$$(1.3) \quad C(m) = \left( \frac{4\pi}{m+4} \right)^{\frac{m}{2}} \Gamma\left(1 + \frac{m}{2}\right) \underbrace{\int_{\mathcal{S}_m} |\det X| d\gamma_*(X)}_{=: I_m}.$$

A similar result holds in the case  $m = 1$ . In this case  $M = S^1$  and  $\mathbf{U}_L$  is the space of trigonometric polynomials of degree  $\leq L$ . One can show (see [34])

$$\mathbf{E}(\mathcal{N}_{L,S^1}) \sim \sqrt{\frac{3}{5}} \dim \mathbf{U}_L \text{ as } L \rightarrow \infty.$$

We can say something about the behavior of  $C(m)$  as  $m \rightarrow \infty$ .

**THEOREM 1.2.**

$$(1.4) \quad \log C(m) \sim \log I_m \sim \frac{m}{2} \log m \text{ as } m \rightarrow \infty.$$

The proof of (1.1) is based on Kac-Rice’s integral formula [1, 4, 8, 14, 15] which expresses the expected number of critical points of a function in  $\mathbf{U}_L$  as an integral

$$(*) \quad \mathbf{E}(\mathcal{N}_L) = \int_M \rho_L(\mathbf{x}) |dV_g(\mathbf{x})|.$$

The above equality was given a geometric interpretation by Chern and Lashof [11]. More precisely, they showed that the integral in the right-hand side of the above equality is the the total curvature of the immersion given by the evaluation map

$$(1.5) \quad \mathbf{ev} : M \rightarrow \text{Hom}(\mathbf{U}_L, \mathbb{R}), \quad \mathbf{p} \mapsto \mathbf{ev}_{\mathbf{p}},$$

where  $\mathbf{ev}_{\mathbf{p}}(\mathbf{u}) = \mathbf{u}(\mathbf{p}), \forall \mathbf{u} \in \mathbf{U}_L$ .

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<sup>1</sup>We refer to Appendix B for a detailed description of a 3-parameter family Gaussian measures  $d\Gamma_{a,b,c}$  on  $\mathcal{S}_m$  that includes  $d\gamma_*$  as  $d\Gamma_{3,1,1}$ .

For our purposes the probabilistic description of the integrand  $\rho_L(\mathbf{x})$  is more useful. To formulate it let us denote by  $\text{Hess}_{\mathbf{x}}(\mathbf{u}, g)$  the Hessian at  $\mathbf{x}$  of the random function  $\mathbf{u} \in \mathbf{U}_L$  computed using the Levi-Civita connection of the metric  $g$ . Using the metric we identify  $\text{Hess}_{\mathbf{x}}(\mathbf{u}, g)$  with a symmetric linear operator  $T_{\mathbf{x}}M \rightarrow T_{\mathbf{x}}M$ . Then

$$(1.6) \quad \rho_L(\mathbf{x}) = \frac{1}{\sqrt{\det 2\pi S_{d\mathbf{u}(x)}}} \mathbf{E} \left( |\det \text{Hess}_{\mathbf{x}}(\mathbf{u}, g)| \mid d\mathbf{u}(x) = 0 \right).$$

Above,  $S_{d\mathbf{u}(x)}$  denotes covariance matrix of the Gaussian vector  $\mathbf{U}_L \ni \mathbf{u} \mapsto d\mathbf{u}(x) \in T_{\mathbf{x}}^*M$ , while the quantity

$$\mathbf{E} \left( |\det \text{Hess}_{\mathbf{x}}(\mathbf{u}, g)| \mid d\mathbf{u}(x) = 0 \right)$$

is the conditional expectation of the random variable  $\mathbf{u} \mapsto |\det \text{Hess}_{\mathbf{x}}(\mathbf{u}, g)|$  given that  $d\mathbf{u}(x) = 0$ .

Using the *regression formula* (see [4, Prop. 1.2] or (A.2)) we express this conditional expectation as the unconditional expectation of a new random variable  $|\det A_L(\mathbf{x})|$ , where  $A_L(\mathbf{x})$  denotes a random, Gaussian symmetric  $m \times m$  matrix whose covariance takes into account the correlations between the Gaussian variables  $\mathbf{u} \mapsto \text{Hess}_{\mathbf{x}}(\mathbf{u}, g)$  and  $\mathbf{u} \mapsto d\mathbf{u}(x)$ .

Next, we reduce the large  $L$  asymptotics of the Gaussian random vector  $d\mathbf{u}(x)$  and matrix  $A_L(\mathbf{x})$  to questions concerning the asymptotics of the spectral function  $\mathcal{E}_L$  of the Laplacian, i.e., the Schwartz kernel of the orthogonal projection onto  $\mathbf{U}_L$ . These issues were addressed in the pioneering work of L. Hörmander [21].

We actually prove a bit more. We show that

$$(1.7) \quad \lim_{L \rightarrow \infty} L^{-\frac{m}{2}} \rho_L(\mathbf{x}) = \frac{C(m)\omega_m}{(2\pi)^m}, \quad \text{uniformly in } \mathbf{x} \in M,$$

where  $\omega_m$  denotes the volume of the unit ball in  $\mathbb{R}^m$ . Using the classical Weyl estimates (3.2) we see that (1.7) implies (1.1).

The equality (1.7) has an interesting interpretation. We can think of  $\rho_L(\mathbf{x})|dV_g(\mathbf{x})|$  as the expected number of critical points of a random function in  $\mathbf{U}_L$  inside an infinitesimal region of volume  $|dV_g(\mathbf{x})|$  around the point  $\mathbf{x}$ . From this point of view we see that (1.7) states that *for large  $L$  we expect the critical points of a random function in  $\mathbf{U}_L$  to be approximatively uniformly distributed.*

We are inclined to believe that as  $L \rightarrow \infty$  the ratio

$$q_L = \frac{\text{var}(\mathcal{N}_L)}{\mathbf{E}(\mathcal{N}_L)}$$

has a finite limit  $q(M, g)$ . Such a result would show that  $\mathcal{N}_L$  is highly concentrated near its mean value as  $L \rightarrow \infty$ . In [34] we proved that this is the case when  $M = S^1$  and moreover,  $q(S^1) \approx 0.4518\dots$ . In [35] we proved that a closely related concentration result is valid for all flat tori.

For a holomorphic counterpart of such an estimate we refer to [42].

We obtain the asymptotics of  $C(m)$  by relying on a trick used by Y.V. Fyodorov [17] in a related context. This reduces the asymptotics of the integral  $I_m$  to known asymptotics of the 1-point correlation function in random matrix theory, more precisely, Wigner’s semi-circle law.

Philosophically, the universality result contained in Theorem 1.1 is a consequence of a universal behavior of the spectral function  $\mathcal{E}_L$  along the diagonal. Roughly speaking, if we rescale the metric  $g$  so that in the limit it becomes flatter, and flatter, then the corresponding spectral function begins to resemble the spectral function of the Laplacian on the Euclidean space  $\mathbb{R}^m$ . For a precise formulation of this universal rescaling phenomenon we refer to [25, 33].

A related problem was considered by M. Douglas, B. Shiffman, S. Zelditch, [14, 15] where they investigate the number of critical points of a random holomorphic section of a large power  $N$  of a positive holomorphic line bundle  $\mathcal{L}$  over a Kähler manifold  $X$ . In these papers the role of our  $\mathcal{U}_L$  is played by the space of holomorphic sections  $H^0(X, \mathcal{L}^N)$ , and the large  $L$  asymptotics is replaced by large  $N$  asymptotics. The large  $N$  asymptotics ultimately follow from the refined asymptotics of the Szegő kernels obtained by S. Zelditch in [44]. These refined asymptotics then lead to a complete asymptotic expansion as  $N \rightarrow \infty$  for the expected number of critical points of a random holomorphic section of  $\mathcal{L}^N$ .

The proof of Theorem 1.1 reveals several additional interesting universal rescaling phenomena. We identify  $\mathcal{U}_L$  with  $\mathcal{U}_L^\vee = \text{Hom}(\mathcal{U}_L, \mathbb{R})$  using the  $L^2$ -metric. We can thus view the evaluation map in (1.5) as a map  $\mathbf{ev} : M \rightarrow \mathcal{U}_L$ . For large  $L$  this map is an embedding, and we denote by  $\sigma_L$  the pullback to  $M$  via  $\mathbf{ev}$  of the  $L^2$ -metric on  $\mathcal{U}_L$ . Equivalently, if  $(\psi_k)$  is an orthonormal basis of  $\mathcal{U}_L$ , then

$$\sigma_L = \sum_k d\psi_k \otimes d\psi_k.$$

The equality (3.9) in the proof of Theorem 1.1 shows that the rescaled metric  $g(L) := L^{-\frac{m+2}{2}} \sigma_L$  converges in the  $C^0$  topology to  $K_m g$ , where  $g$  is the original metric on  $M$  and  $K_m$  is a certain, explicit constant that depends only on  $m$ ; see (3.5). This was also observed by S. Zelditch, [45, Prop. 2.3]. A closely related result was proved in [5, Thm.5].

To obtain the convergence of  $g(L)$  in stronger topologies we would need bounds on the sectional curvature of  $g(L)$ . We show that these bounds are equivalent to some refined asymptotic estimates satisfied by certain linear combinations of fourth order derivatives of the spectral function, (3.20).

A related embedding can be constructed in the holomorphic case and S. Zelditch [44] has proved that the resulting sequence of suitably rescaled metrics  $g_N$  converges  $C^\infty$  to the original Kähler metric. The main reason for such a stronger form of convergence is the better behavior of the Szegő kernels. Such a regular behavior is not to be expected for the spectral function  $\mathcal{E}_L$ .

The present paper is structured as follows. Section 2 contains the formulation and the proof of the key integral formula (1.6), including several reformulations in the language of random processes. In this section we also present a simple application of this formula to the number of critical points of random spherical harmonics of large degree on  $S^2$ . This sheds additional light on a recent result of Nazarov and Sodin [30] on the number of nodal domains of random spherical harmonics. More precisely, the inequality (2.40) shows that the expected number  $\delta_n$  of zonal domain on  $S^2$  of a random harmonic polynomial of large degree  $n$  satisfies the upper bound  $\delta_n < 0.29n^2$ .

Section 3 contains the proof of the asymptotic estimate (1.1) and Section 4 contains the proof of the estimate (1.3).

In our experience, many basic probabilistic technologies are not that familiar to an audience with a more geometric background. With this audience in mind we

decided to include in Appendix A a coordinate-free, brief survey of several facts about Gaussian measures and Gaussian processes in a form adapted to the applications in this paper. Appendix B contains a detailed description of a 3-parameter family of Gaussian measures on the space  $\mathcal{S}_m$  of real, symmetric  $m \times m$  matrices. These measures play a central role in the proof of (1.1) and we could not find an appropriate reference for the mostly elementary facts discussed in this appendix. Appendix C contains the computations of a Gaussian integral involving random  $2 \times 2$  matrices.

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**Notations.**

- i. For any random variable  $\xi$  we denote by  $E(\xi)$  and respectively  $\mathit{var}(\xi)$  its expectation and respectively its variance.
- ii.  $\mathcal{S}_m$  denotes the space of symmetric  $m \times m$  real matrices.
- iii. For any finite dimensional real vector space  $\mathbf{V}$  we denote by  $\mathbf{V}^\vee$  its dual,  $\mathbf{V}^\vee := \text{Hom}(\mathbf{V}, \mathbb{R})$ .
- iv. For any Euclidean space  $\mathbf{V}$ , we denote by  $S(\mathbf{V})$  the unit sphere in  $\mathbf{V}$  centered at the origin and by  $B(\mathbf{V})$  the unit ball in  $\mathbf{V}$  centered at the origin.
- v. We will denote by  $\sigma_n$  the “area” of the round  $n$ -dimensional sphere  $S^n$  of radius 1, and by  $\omega_n$  the “volume” of the unit ball in  $\mathbb{R}^n$ . These quantities are uniquely determined by the equalities (see [31, Ex. 9.1.11])

$$(σ) \quad \sigma_{n-1} = n\omega_n = 2 \frac{\pi^{\frac{n}{2}}}{\Gamma(\frac{n}{2})} = \frac{n\pi^{\frac{n}{2}}}{\Gamma(1 + \frac{n}{2})}, \quad \Gamma\left(\frac{1}{2}\right) = \sqrt{\pi},$$

where  $\Gamma$  is Euler’s Gamma function.

- vi. If  $\mathbf{V}_0$  and  $\mathbf{V}_1$  are two Euclidean spaces of dimensions  $n_0, n_1 < \infty$  and  $A : \mathbf{V}_0 \rightarrow \mathbf{V}_1$  is a linear map, then the *Jacobian* of  $A$  is the nonnegative scalar  $J(A)$  defined as the norm of the linear map

$$\Lambda^k A : \Lambda^k \mathbf{V}_0 \rightarrow \Lambda^k \mathbf{V}_1, \quad k := \min(n_0, n_1).$$

More concretely, if  $n_0 \leq n_1$ , and  $\{e_1, \dots, e_{n_0}\}$  is an orthonormal basis of  $\mathbf{V}_0$ , then

$$(J_-) \quad J(A) = (\det G(A))^{1/2},$$

where  $G(A)$  is the  $n_0 \times n_0$  Gramm matrix with entries

$$G_{ij} = (Ae_i, Ae_j)_{\mathbf{V}_1}.$$

If  $n_1 \geq n_0$  then

$$(J_+) \quad J(A) = J(A^\dagger) = (\det G(A^\dagger))^{1/2},$$

where  $A^\dagger$  denotes the adjoint (transpose) of  $A$ . Equivalently, if  $d\text{Vol}_i \in \Lambda^{n_i} \mathbf{V}_i^*$  denotes the metric volume form on  $\mathbf{V}_1$ , and  $d\text{Vol}_A$  denotes the metric volume form on  $\ker A$ , then  $J(A)$  is the positive number such that

$$(J'_+) \quad d\text{Vol}_0 = \pm d\text{Vol}_A \wedge A^* d\text{Vol}_1.$$

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<sup>2</sup>He suddenly and untimely passed away in June 2011. I will miss his generosity and expertise.

**2. A Kac-Rice type formula.**

**2.1. The key integral formula.** As we mentioned in the introduction, a key component in the proof of Theorem 1.1 is an integral formula that describes the expected number of critical points as an integral over the background manifold  $M$ . The literature on random fields contains many formulæ of this type, and their proofs follow the strategy pioneered by M. Kac and S. Rice, [1, 4, 24, 40].

We believe that it would greatly benefit a reader less fluent in the probabilistic language to first see the geometric origins of these formulae. For this reason we decided to include a complete proof of these formulae in our special case. Not surprisingly, the ubiquitous double-fibration trick in integral geometry, [2, 19, 31] will carry the day. As a matter of fact, our main integral formula (2.2) contains as special cases the integral formulæ of Chern-Lashof, [11] and Milnor, [27].

Suppose that  $M$  is a smooth, compact, connected manifold without boundary. Set  $m := \dim M$ .

DEFINITION 2.1. (a) For any nonnegative integer  $k$ , any point  $\mathbf{p} \in M$  and any  $f \in C^\infty(M)$  we will denote by  $j_k(f, \mathbf{p})$  the  $k$ -th jet of  $f$  at  $\mathbf{p}$ .  
 (b) Suppose that  $\mathbf{U} \subset C^\infty(M)$  is a linear subspace. If  $k$  is nonnegative integer, then we say that  $\mathbf{U}$  is  $k$ -ample if for any  $\mathbf{p} \in M$  and any  $f \in C^\infty(M)$  there exists  $\mathbf{u} \in \mathbf{U}$  such that

$$j_k(\mathbf{u}, \mathbf{p}) = j_k(f, \mathbf{p}). \quad \square$$

Fix a finite dimensional vector space  $\mathbf{U} \subset C^\infty(M)$  and set  $N := \dim \mathbf{U}$ . We have an evaluation map

$$\mathbf{ev} = \mathbf{ev}^{\mathbf{U}} : M \rightarrow \mathbf{U}^\vee := \text{Hom}(\mathbf{U}, \mathbb{R}), \quad \mathbf{p} \mapsto \mathbf{ev}_{\mathbf{p}},$$

where for any  $\mathbf{p} \in M$  the linear map  $\mathbf{ev}_{\mathbf{p}} : \mathbf{U} \rightarrow \mathbb{R}$  is given by

$$\mathbf{ev}_{\mathbf{p}}(\mathbf{u}) = \mathbf{u}(\mathbf{p}), \quad \forall \mathbf{u} \in \mathbf{U}.$$

For any  $\mathbf{u} \in C^\infty(M)$  we denote by  $\mathcal{N}(\mathbf{u})$  the number of critical points of  $\mathbf{u}$ . In the remainder of this section we will assume that  $\mathbf{U}$  is 1-ample. This implies that the evaluation map  $\mathbf{ev}^{\mathbf{U}}$  is an immersion. Moreover, as explained in [32, Cor. 1.26], the 1-ampleness condition also implies that almost all functions  $\mathbf{u} \in \mathbf{U}$  are Morse functions and thus  $\mathcal{N}(\mathbf{u}) < \infty$  for almost all  $\mathbf{u} \in \mathbf{U}$ .

We fix an inner product  $h = (-, -)_h$  on  $\mathbf{U}$  and we denote by  $|\cdot|_h$  the resulting Euclidean norm. Using the metric  $h$  we identify  $\mathbf{U}$  with its dual and thus we can regard the evaluation map as a smooth map  $\mathbf{ev} : M \rightarrow \mathbf{U}$ . We define the expected number of critical points of a function in  $\mathbf{U}$  to be the quantity

$$(2.1) \quad \mathcal{N}(\mathbf{U}, h) := \frac{1}{\sigma_{N-1}} \int_{S(\mathbf{U})} \mathcal{N}(\mathbf{u}) |dA_h(\mathbf{u})| = \int_{\mathbf{U}} \mathcal{N}(\mathbf{u}) \underbrace{\frac{e^{-\frac{|\mathbf{u}|_h^2}{2}}}{(2\pi)^{\frac{N}{2}}} |dV_h(\mathbf{u})|}_{=: d\gamma_h(\mathbf{u})},$$

where  $\sigma_{n-1}$  denotes the "area" of the unit sphere in  $\mathbb{R}^n$ ,  $|dA_h|$  denotes the "area" density on the unit sphere  $S(\mathbf{U})$ , and  $|dV_h(\mathbf{u})|$  denotes the volume density on  $\mathbf{U}$  determined by the metric  $h$ . A priori, the expected number of critical points could be

infinite, but in any case, it is independent of any choice of metric on  $M$ . The space  $\mathbf{U}$  equipped with the Gaussian probability measure  $d\gamma_h$  is a probability space. We denote by  $\mathcal{N}_U$  the random variable  $\mathbf{U} \ni \nu \mapsto \mathcal{N}(\mathbf{u}) \in \mathbb{Z}$  so that

$$\mathcal{N}(\mathbf{U}, h) = \mathbf{E}(\mathcal{N}_U, d\gamma_h),$$

where  $\mathbf{E}(-, d\gamma_h)$  denotes the expectation computed with respect to the probability measure  $d\gamma_h$ . We will refer to the pair  $(\mathbf{U}, h)$  as the *sample space*.

Fix a metric  $g$  on  $M$ . We will express  $\mathcal{N}(\mathbf{U}, h)$  as an integral

$$\int_M \rho_g(\mathbf{p}) |dV_g(\mathbf{p})|.$$

The function  $\rho_g$  does depend on  $g$ , but the density  $\rho_g(\mathbf{p}) |dV_g(\mathbf{p})|$  is independent of  $g$ . The concrete description of  $\rho_g(\mathbf{p})$  relies on several fundamental objects naturally associated to the triplet  $(\mathbf{U}, h, g)$ .

For any  $\mathbf{p} \in M$  we set

$$U_{\mathbf{p}}^0 := \{\mathbf{u} \in \mathbf{U}; d\mathbf{u}(\mathbf{p}) = 0\}.$$

The 1-ampleness assumption on  $\mathbf{U}$  implies that for any  $\mathbf{p} \in M$  the subspace  $U_{\mathbf{p}}^0$  has codimension  $m$  in  $\mathbf{U}$  so that  $\dim U_{\mathbf{p}}^0 = N - m$ . Denote by  $dA_{S(U_{\mathbf{p}}^0)}$  the area density along the unit sphere  $S(U_{\mathbf{p}}^0) \subset U_{\mathbf{p}}^0$ .

The differential of the evaluation map at  $\mathbf{p}$  is a linear map  $\mathcal{A}_{\mathbf{p}} : T_{\mathbf{p}}M \rightarrow \mathbf{U}$ . We will refer to  $\mathcal{A}_{\mathbf{p}}$  as the *adjunction map* and we will denote by  $J_g(\mathbf{p}) = J_g(\mathbf{p}, \mathbf{U})$  its Jacobian. More precisely, if  $(\mathbf{e}_1, \dots, \mathbf{e}_m)$  is a  $g$ -orthonormal basis of  $T_{\mathbf{p}}M$ , then

$$J_g(\mathbf{p})^2 = \det \left[ (\mathcal{A}_{\mathbf{p}}\mathbf{e}_i, \mathcal{A}_{\mathbf{p}}\mathbf{e}_j)_h \right]_{1 \leq i, j \leq m}.$$

Since  $\mathbf{ev}^U$  is an immersion we have  $J_g(\mathbf{p}) \neq 0, \forall \mathbf{p} \in M$ .

For any  $\mathbf{p} \in M$  and any  $\mathbf{u} \in U_{\mathbf{p}}^0$ , the Hessian of  $\mathbf{u}$  at  $\mathbf{p}$  is a well defined symmetric bilinear form on  $T_{\mathbf{p}}M$  that can be identified via the metric  $g$  with a symmetric endomorphism  $\text{Hess}_{\mathbf{p}}(\mathbf{u}, g)$  of  $T_{\mathbf{p}}M$ . We denote this symmetric endomorphism by  $\text{Hess}_{\mathbf{p}}(\mathbf{u}, g)$ .

**THEOREM 2.2.** *If  $(\mathbf{U}, h)$  is an 1-ample sample space on  $M$ , then*

$$\begin{aligned} (2.2) \quad \mathcal{N}(\mathbf{U}, h) &= \frac{1}{\sigma_{N-1}} \int_M \frac{1}{J_g(\mathbf{p})} \left( \int_{S(U_{\mathbf{p}}^0)} |\det \text{Hess}_{\mathbf{p}}(\mathbf{v}, g)| |dA_{S(U_{\mathbf{p}}^0)}(\mathbf{v})| \right) |dV_g(\mathbf{p})| \\ &= (2\pi)^{-\frac{m}{2}} \int_M \frac{1}{J_g(\mathbf{p})} \underbrace{\left( \int_{U_{\mathbf{p}}^0} |\det \text{Hess}_{\mathbf{p}}(\mathbf{u}, g)| \frac{e^{-\frac{|\mathbf{u}|_h^2}{2}}}{(2\pi)^{\frac{N-m}{2}}} |dV_h(\mathbf{u})| \right)}_{=: I_{\mathbf{p}}} |dV_g(\mathbf{p})|. \end{aligned}$$

*Proof.* Denote by  $\underline{U}_M$  the trivial vector bundle over  $M$  with fiber  $\mathbf{U}$ ,  $\underline{U}_M := (\mathbf{U} \times M \rightarrow M)$ . For any  $\mathbf{p} \in M$  we denote by  $\mathbf{K}_{\mathbf{p}}$  the orthogonal complement of  $U_{\mathbf{p}}^0$  in  $\mathbf{U}$ .

**LEMMA 2.3.** *The subspace  $\mathbf{K}_{\mathbf{p}}$  coincides with the range of the adjunction map  $\mathcal{A}_{\mathbf{p}}$ .*

*Proof.* Indeed, if  $(\Psi_n)_{1 \leq n \leq N}$  is an orthonormal basis of  $(U, h)$ , then

$$e\mathbf{v}_p = \sum_n \Psi_n(p)\Psi_n \in U.$$

and for any vector field  $X$  on  $M$  we have

$$\mathcal{A}_p X_p = \sum_n (X\Psi_n)_p \Psi_n.$$

Thus, the function  $\mathbf{u} = \sum_n u_n \Psi_n \in U$ ,  $u_n \in \mathbb{R}$ , belongs to  $\mathbf{K}_p^\perp$  if and only if for any vector field  $X$  on  $M$  we have

$$0 = \sum_n u_n (X\Psi_n)_p = X \cdot \mathbf{u}(p) \iff \mathbf{u} \in U_p^0.$$

□

This proves that the collection  $(\mathbf{K}_p)$  defines a subbundle  $\mathbf{K}$  of  $\underline{U}_M$  and the adjunction map induces an isomorphism of vector bundle  $\mathcal{A} : TM \rightarrow \mathbf{K}$ . We deduce that the collection of spaces  $(U_p^0)_{p \in M}$  also forms a vector subbundle  $U^0$  of the trivial bundle  $\underline{U}_M$  and we have an orthogonal direct sum decomposition

$$\underline{U}_M = U^0 \oplus \mathbf{K}.$$

For any section  $u$  of  $\underline{U}_M$  we denote by  $u^0$  its  $U^0$ -component.

The bundle  $\underline{U}_M$  is equipped with a canonical trivial connection  $D$ . More precisely, if we regard a section  $u$  of  $\underline{U}_M$  as a smooth map  $u : M \rightarrow U$ , then for any vector field  $X$  on  $M$  we define  $D_X u$  as the smooth function  $M \rightarrow U$  obtained by derivating  $u$  along  $X$ . The *shape operator* of the subbundle  $\mathbf{K}$  is the bundle morphism  $\Xi : TM \otimes \mathbf{K} \rightarrow U^0$  defined by the equality

$$\Xi(X, \mathbf{u}) := (D_X \mathbf{u})^0, \quad \forall X \in C^\infty(TM), \quad \mathbf{u} \in C^\infty(\mathbf{K}).$$

For every  $p \in M$ , we denote by  $\Xi_p$  the induced linear map  $\Xi_p : T_p M \otimes \mathbf{K}_p \rightarrow U^0$ . If we denote by  $\mathbf{Gr}_m(U)$  the Grassmannian of  $m$ -dimensional subspaces of  $U$ , then we have a Gauss map

$$M \ni p \xrightarrow{\mathcal{G}} \mathcal{G}(p) := \mathbf{K}_p \in \mathbf{Gr}_m(U).$$

The shape operator  $\Xi_p$  can be viewed as a linear map

$$\Xi_p : T_p M \rightarrow \text{Hom}(\mathbf{K}_p, U_p^0) = T_{\mathbf{K}_p} \mathbf{Gr}_m(U),$$

and, as such, it can be identified with the differential of  $\mathcal{G}$  at  $p$ , [31, §9.1.2]. Any  $\mathbf{v} \in U_p^0$  determines a bilinear map

$$\Xi_p \cdot \mathbf{v} : T_p M \otimes \mathbf{K}_p \rightarrow \mathbb{R}, \quad \Xi_p \cdot \mathbf{v}(e, \mathbf{u}) := (\Xi_p(e, \mathbf{u}), \mathbf{v})_h.$$

By choosing orthonormal bases  $(e_i)$  in  $T_p M$  and  $(u_j)$  of  $\mathbf{K}_p$  we can identify this bilinear form with an  $m \times m$ -matrix. This matrix depends on the choices of bases, but the absolute value of its determinant is independent of these bases. It is thus an invariant of the pair  $(\Xi_p, \mathbf{v})$  that we will denote by  $|\det \Xi_p \cdot \mathbf{v}|$ .



LEMMA 2.4.

$$(2.3) \quad \mathcal{N}(\mathbf{U}, h) = \frac{1}{\sigma_{N-1}} \int_M \left( \int_{S(\mathbf{U}_{\mathbf{p}}^0)} |\det \Xi_{\mathbf{p}} \cdot \mathbf{v}| |dA_{S(\mathbf{U}_{\mathbf{p}}^0)}(\mathbf{v})| \right) |dV_g(\mathbf{p})|.$$

*Proof.* Consider the incidence variety

$$\mathcal{J} := \{(\mathbf{p}, \mathbf{v}) \in M \times S(\mathbf{U}); d\mathbf{v}(\mathbf{p}) = 0\} = \{(\mathbf{x}, \mathbf{v}) \in M \times S(\mathbf{U}); \mathbf{v} \in S(\mathbf{U}_{\mathbf{p}}^0)\}.$$

We have a natural double “fibration”

$$M \xleftarrow{\lambda} \mathcal{J} \xrightarrow{\rho} S(\mathbf{U}),$$

where the left/right projections  $\lambda, \rho$  are the canonical projections. The left projection  $\lambda : \mathcal{J} \rightarrow M$  describes  $\mathcal{J}$  as the unit sphere bundle associated to the metric vector bundle  $\mathbf{U}^0$ . In particular, this shows that  $\mathcal{J}$  is a compact, smooth manifold of dimension  $(N - 1)$ . For generic  $\mathbf{v} \in S(\mathbf{U})$  the fiber  $\rho^{-1}(\mathbf{v})$  is finite and can be identified with the set of critical points of  $\mathbf{v} : M \rightarrow \mathbb{R}$ . We deduce

$$(2.4) \quad \mathcal{N}(\mathbf{U}, h) = \frac{1}{\text{area}(S(\mathbf{U}))} \int_{S(\mathbf{U})} \#\rho^{-1}(\mathbf{v}) |dA_h(\mathbf{u})|.$$

Denote by  $g_{\mathcal{J}}$  the metric on  $\mathcal{J}$  induced by the metric on  $M \times S(\mathbf{U})$  and by  $|dV_{\mathcal{J}}|$  the induced volume density. The coarea formula, [10, §13.4], implies that

$$(2.5) \quad \int_{S(\mathbf{U})} \#\rho^{-1}(\mathbf{v}) |dA_h(\mathbf{v})| = \int_{\mathcal{J}} J_{\rho}(\mathbf{p}, \mathbf{v}) |dV_{\mathcal{J}}(\mathbf{p}, \mathbf{v})|,$$

where the nonnegative function  $J_{\rho}$  is the Jacobian of  $\rho$  defined by the equality

$$\rho^* |dA_h| = J_{\rho} \cdot |dV_{\mathcal{J}}|.$$

To compute the integral in the right-hand side of (2.5) we need a more explicit description of the geometry of the incidence variety  $\mathcal{J}$ .

Fix a local orthonormal frame  $(e_1, \dots, e_m)$  of  $TM$  defined in a neighborhood  $\mathcal{O}$  in  $M$  of a given point  $\mathbf{p}_0 \in M$ . We denote by  $(e^1, \dots, e^m)$  the dual co-frame of  $T^*M$ . Set

$$\mathbf{f}_i(\mathbf{p}) := A_{\mathbf{p}} e_i(\mathbf{p}) \in \mathbf{U}, \quad i = 1, \dots, m, \quad \mathbf{p} \in \mathcal{O}.$$

More explicitly,  $\mathbf{f}_i(\mathbf{u})$  is defined by the equality

$$(2.6) \quad (\mathbf{f}_i(\mathbf{p}), \mathbf{v})_h = \partial_{e_i} \mathbf{u}(\mathbf{p}), \quad \forall \mathbf{u} \in \mathbf{U}.$$

Fix a neighborhood  $\mathcal{U} \subset \lambda^{-1}(\mathcal{O})$  in  $M \times S(\mathbf{U})$  of the point  $(\mathbf{p}_0, \mathbf{v}_0)$ , and a local orthonormal frame  $\mathbf{u}_1(\mathbf{p}, \mathbf{v}), \dots, \mathbf{u}_{N-1}(\mathbf{p}, \mathbf{v})$  over  $\mathcal{U}$  of the bundle  $\rho^*TS(\mathbf{U}) \rightarrow M \times S(\mathbf{U})$  such that the following hold.

- The vectors  $\mathbf{u}_1(\mathbf{p}, \mathbf{v}), \dots, \mathbf{u}_m(\mathbf{p}, \mathbf{v})$  are independent of the variable  $\mathbf{v}$  and form an orthonormal basis of  $K_{\mathbf{x}}^{\perp}$ . (E.g., we can obtain such vectors from the vectors  $\mathbf{f}_1(\mathbf{p}), \dots, \mathbf{f}_m(\mathbf{p})$  via the Gramm-Schmidt process.)
- For  $(\mathbf{p}, \mathbf{v}) \in \mathcal{U}$ , the space  $T_{\mathbf{p}}E_{\mathbf{x}}$  is spanned by the vectors  $\mathbf{u}_{m+1}(\mathbf{p}, \mathbf{v}), \dots, \mathbf{u}_{N-1}(\mathbf{p}, \mathbf{v})$ .

The collection  $\mathbf{u}_1(\mathbf{p}), \dots, \mathbf{u}_m(\mathbf{p})$  is a collection of smooth sections of  $\underline{U}_M$  over  $\mathcal{O}$ . For any  $\mathbf{p} \in \mathcal{O}$  and any  $\mathbf{e} \in T_{\mathbf{p}}M$ , we obtain the vectors (functions).

$$D_{\mathbf{e}}\mathbf{u}_1(\mathbf{p}), \dots, D_{\mathbf{p}}\mathbf{u}_m(\mathbf{x}) \in \mathbf{U},$$

where we recall that  $D$  denotes the trivial connection on  $\underline{U}_M$ . Observe that

$$(2.7) \quad \mathcal{J} \cap \mathcal{U} = \{(\mathbf{p}, \mathbf{v}) \in \mathcal{U}; U_i(\mathbf{p}, \mathbf{v}) = 0, \forall i = 1, \dots, m\},$$

where  $U_i$  is the function  $U_i : \mathcal{O} \times \mathbf{U} \rightarrow \mathbb{R}$  given by

$$U_i(\mathbf{p}, \mathbf{v}) := (\mathbf{u}_i(\mathbf{p}), \mathbf{v})_h.$$

Thus, the tangent space of  $\mathcal{J}$  at  $(\mathbf{p}, \mathbf{v})$  consists of tangent vectors  $\dot{\mathbf{p}} \oplus \dot{\mathbf{v}} \in T_{\mathbf{x}}M \oplus T_{\mathbf{v}}S(\mathbf{V})$  such that

$$dU_i(\dot{\mathbf{p}}, \dot{\mathbf{v}}) = 0, \quad \forall i = 1, \dots, m.$$

We let  $\omega_U$  denote the  $m$ -form

$$\omega_U := dU_1 \wedge \dots \wedge dU_m \in \Omega^m(\mathcal{U}),$$

and we denote by  $\|\omega_U\|$  its norm with respect to the product metric on  $M \times S(\mathbf{U})$ . Denote by  $|d\widehat{V}|$  the volume density on  $M \times S(\mathbf{U})$  induced by the product metric. The equality (2.7) implies that

$$|d\widehat{V}| = \frac{1}{\|\omega_U\|} |\omega_U \wedge dV_E|.$$

Hence

$$J_{\rho}|d\widehat{V}| = \frac{1}{\|\omega_U\|} |\omega_U \wedge \rho^*dA|.$$

We deduce

$$\begin{aligned} J_{\rho}(\mathbf{p}_0, \mathbf{v}_0) &= J_{\rho}(\mathbf{p}_0, \mathbf{v}_0)|d\widehat{V}|(\mathbf{e}_1, \dots, \mathbf{e}_m, \mathbf{u}_1, \dots, \mathbf{u}_{N-1}) \\ &= \frac{1}{\|\omega_U\|} |\omega_U \wedge \rho^*dS|(\mathbf{e}_1, \dots, \mathbf{e}_m, \mathbf{u}_1, \dots, \mathbf{u}_{N-1}) = \frac{1}{\|\omega_U\|} \underbrace{|\omega_U(\mathbf{e}_1, \dots, \mathbf{e}_m)|}_{=: \Delta_U(\mathbf{p}_0, \mathbf{v}_0)} \Big|_{(\mathbf{p}_0, \mathbf{v}_0)}. \end{aligned}$$

Hence,

$$(2.8) \quad \int_{S(\mathbf{U})} \#\rho^{-1}(\mathbf{w})|dA_h(\mathbf{v})| = \int_{\mathcal{J}} \frac{\Delta_U}{\|\omega_U\|} |dV_{\mathcal{J}}(\mathbf{p}, \mathbf{v})|.$$

SUBLEMMA 2.5. *We have the equality*

$$(2.9) \quad J_{\lambda} = \frac{1}{\|\omega_U\|},$$

where  $J_{\lambda}$  denotes the Jacobian of the projection  $\lambda : \mathcal{J} \rightarrow M$ .

*Proof of Sublemma 2.5.* Along  $\mathcal{U}$  we have

$$|\widehat{dV}| = \frac{1}{\|\omega_U\|} |\omega_U \wedge dV_{\mathcal{J}}|,$$

while the definition of the Jacobian implies that

$$|dV_{\mathcal{J}}| = \frac{1}{J_{\lambda}} |dV_g \wedge dA_{S(\mathcal{U}_{\mathbf{p}}^0)}|.$$

Therefore, it suffices to show that along  $\mathcal{U}$  we have

$$|\widehat{dV}| = |\omega_U \wedge dV_g \wedge dA_{S(\mathcal{U}_{\mathbf{p}}^0)}|,$$

i.e.,

$$\left| \omega_U \wedge dV_g \wedge dA_{S(\mathcal{U}_{\mathbf{p}}^0)}(\mathbf{e}_1, \dots, \mathbf{e}_m, \mathbf{u}_1, \dots, \mathbf{u}_{N-1}) \right| = 1.$$

Since  $dU_i(\mathbf{u}_k) = 0, \forall k \geq m + 1$  we deduce that

$$\left| \omega_U \wedge dV_g \wedge dA_{S(\mathcal{U}_{\mathbf{p}}^0)}(\mathbf{e}_1, \dots, \mathbf{e}_m, \mathbf{u}_1, \dots, \mathbf{u}_{N-1}) \right| = |\omega_U(\mathbf{u}_1, \dots, \mathbf{u}_m)|.$$

Thus, it suffices to show that  $|\omega_U(\mathbf{u}_1, \dots, \mathbf{u}_m)| = 1$ . This follows from the elementary identities

$$dU_i(\mathbf{u}_j) = (\mathbf{u}_i, \mathbf{u}_j)_h = \delta_{ij}, \quad \forall 1 \leq i, j \leq m,$$

where  $\delta_{ij}$  is the Kronecker symbol.  $\square$

Using (2.9) in (2.8) and the coarea formula we deduce

$$(2.10) \quad \int_{S(\mathcal{U})} \#\rho^{-1}(\mathbf{w}) |dA_h(\mathbf{v})| = \int_M \left( \int_{S(\mathcal{U}_{\mathbf{p}}^0)} \Delta_U(\mathbf{p}, \mathbf{v}) |dA_{S(\mathcal{U}_{\mathbf{p}}^0)}(\mathbf{v})| \right) |dV_g(\mathbf{p})|.$$

Observe that at a point  $(\mathbf{p}, \mathbf{v}) \in \lambda^{-1}(\mathcal{O}) \subset \mathcal{J}$  we have

$$dU_i(\mathbf{e}_j) = (D_{\mathbf{e}_j} \mathbf{u}_i(\mathbf{p}), \mathbf{v})_h.$$

We can rewrite this in terms of the shape operator  $\Xi_{\mathbf{p}} : T_{\mathbf{p}}M \otimes \mathbf{K}_{\mathbf{p}} \rightarrow \mathcal{U}_{\mathbf{p}}^0$ . More precisely,

$$dU_i(\mathbf{e}_j) = (\Xi_{\mathbf{p}}(\mathbf{e}_j, \mathbf{u}_i), \mathbf{v})_h.$$

Hence,

$$\Delta_U(\mathbf{x}, \mathbf{v}) = |\det(\Xi_{\mathbf{p}}(\mathbf{e}_j, \mathbf{u}_i), \mathbf{v})_h|,$$

We conclude that

$$\int_{S_h(\mathcal{U})} \#\rho^{-1}(\mathbf{v}) |dA_h(\mathbf{v})| = \int_M \left( \int_{S(\mathcal{U}_{\mathbf{p}}^0)} |\det \Xi_{\mathbf{p}} \cdot \mathbf{v}| |dA_{S(\mathcal{U}_{\mathbf{p}}^0)}(\mathbf{v})| \right) |dV_M(\mathbf{p})|.$$

This proves (2.3)  $\square$

To proceed further observe that the left-hand side of (2.3) is plainly independent of the metric  $g$  on  $M$ . This raises the hope that if we judiciously choose the metric on  $M$ , then we can obtain a more manageable expression for  $\mu(M, \mathbf{U})$ . One choice presents itself.

Let  $\sigma$  be the pullback to  $M$  of the metric on  $\mathbf{U}$  via the immersion  $\mathbf{e}\mathbf{v} : M \rightarrow \mathbf{U}$ . More concretely, for any  $\mathbf{p} \in M$  and any  $X, Y \in T_{\mathbf{p}}M$ , we have

$$\sigma_{\mathbf{p}}(X, Y) = (\mathcal{A}_{\mathbf{p}}X, \mathcal{A}_{\mathbf{p}}Y)_h.$$

Fix  $\mathbf{p} \in M$  and a  $\sigma$ -orthonormal frame  $(\mathbf{e}_i)_{1 \leq i \leq m}$  of  $TM$  defined in a neighborhood  $\mathcal{O}$  of  $\mathbf{p}$ . Then the collection  $\mathbf{u}_j = \mathcal{A}\mathbf{e}_j$ ,  $1 \leq j$ , is a local orthonormal frame of  $\mathbf{K}|_{\mathcal{O}}$ . The shape operator has the simple description

$$\Xi_{\mathbf{p}}(\mathbf{e}_i, \mathbf{u}_j) = (D_{\mathbf{e}_i}\mathcal{A}\mathbf{e}_j)^0.$$

Fix an orthonormal basis  $(\Psi_n)_{1 \leq n \leq N}$  of  $\mathbf{U}$  so that every  $\mathbf{v} \in \mathbf{U}$  has a decomposition

$$\mathbf{v} = \sum_{\alpha} v_n \Psi_n, \quad v_n \in \mathbb{R}.$$

Then

$$\mathcal{A}_{\mathbf{p}}\mathbf{e}_j(\mathbf{p}) = \sum_n (\partial_{\mathbf{e}_j} \Psi_n)_{\mathbf{p}} \Psi_n, \quad D_{\mathbf{e}_i} \mathcal{A}^{\dagger} \mathbf{e}_j(\mathbf{p}) = \sum_n (\partial_{\mathbf{e}_i \mathbf{e}_j}^2 \Psi_n)_{\mathbf{p}} \Psi_n,$$

and

$$((D_{\mathbf{e}_i} \mathcal{A}\mathbf{e}_j)_{\mathbf{p}}, \mathbf{v})_h = \sum_{\alpha} v_n (\partial_{\mathbf{e}_i \mathbf{e}_j}^2 \Psi_{\alpha})_{\mathbf{p}} = \partial_{\mathbf{e}_i \mathbf{e}_j}^2 \mathbf{v}(\mathbf{p}).$$

If  $\mathbf{v} \in \mathbf{U}_{\mathbf{p}}^0$ , then the Hessian of  $\mathbf{v}$  at  $\mathbf{p}$  is a well-defined, symmetric bilinear form  $\text{Hess}_{\mathbf{p}}(\mathbf{v})$  on  $T_{\mathbf{p}}M$  that can be identified via the metric  $\sigma$  with a symmetric linear operator

$$\text{Hess}_{\mathbf{p}}(\mathbf{v}, \sigma) : T_{\mathbf{p}}M \rightarrow T_{\mathbf{p}}M.$$

If we fix a  $\sigma$ -orthonormal frame  $(\mathbf{e}_i)$  of  $T_{\mathbf{p}}M$ , then the operator  $\text{Hess}_{\mathbf{p}}(\mathbf{v}, \sigma)$  is described by the symmetric  $m \times m$  matrix with entries  $\partial_{\mathbf{e}_i \mathbf{e}_j}^2 \mathbf{v}(\mathbf{x})$ . We deduce that

$$|\det \Xi_{\mathbf{p}} \cdot \mathbf{v}| = |\det \text{Hess}_{\mathbf{p}}(\mathbf{v}, \sigma)|, \quad \forall \mathbf{v} \in S(\mathbf{U}_{\mathbf{p}}^0).$$

In particular, we deduce that

$$(2.11) \quad \mathcal{N}(\mathbf{U}, h) = \frac{1}{\sigma_{N-1}} \int_M \left( \int_{S(\mathbf{U}_{\mathbf{p}}^0)} |\det \text{Hess}_{\mathbf{p}}(\mathbf{v}, \sigma)| |dA_{S(\mathbf{U}_{\mathbf{p}}^0)}(\mathbf{v})| |dV_{\sigma}(\mathbf{p})| \right).$$

This is precisely the main theorem of Chern and Lashof, [11].

Finally, we want to express (2.11) entirely in terms of the adjunction map  $\mathcal{A}$ . For any  $\mathbf{p} \in M$  and any  $\mathbf{v} \in \mathbf{U}_{\mathbf{p}}$ , we define the density

$$\mu_{\mathbf{p}, \mathbf{v}} : \Lambda^m T_{\mathbf{p}}M \rightarrow \mathbb{R},$$

$$\begin{aligned} \mu_{\mathbf{p}, \mathbf{v}}(X_1 \wedge \cdots \wedge X_m) &= \left| \det(\partial_{X_i X_j}^2 \mathbf{v}(\mathbf{p}))_{1 \leq i, j \leq m} \right| \cdot \left( \det((\mathcal{A}_{\mathbf{p}}X_i, \mathcal{A}_{\mathbf{p}}X_j)_h)_{1 \leq i, j \leq m} \right)^{-1/2} \\ &= \left| \det(\text{Hess}_{\mathbf{p}}(\mathbf{v})(X_i, X_j))_{1 \leq i, j \leq m} \right| \cdot \left( \det(\sigma(X_i, X_j))_{1 \leq i, j \leq m} \right)^{-1/2}, \end{aligned}$$

for any basis  $X_1, \dots, X_m$  of  $T_{\mathbf{p}}M$ . Observe that for any  $\sigma$ -orthonormal frame  $\mathbf{e}_1, \dots, \mathbf{e}_m$  of  $T_{\mathbf{p}}M$  we have

$$\mu_{\mathbf{p},\mathbf{v}}(\mathbf{e}_1 \wedge \dots \wedge \mathbf{e}_m) = |\det \text{Hess}_{\mathbf{p}}(\mathbf{v}, \sigma)|.$$

If we integrate  $\mu_{\mathbf{p},\mathbf{v}}$  over  $\mathbf{v} \in S(\mathbf{U}_{\mathbf{p}}^0)$ , we obtain a density

$$|d\mu_{\mathbf{U}}(\mathbf{p})| : \Lambda^m T_{\mathbf{p}}M \rightarrow \mathbb{R},$$

$$|d\mu_{\mathbf{U}}(\mathbf{p})|(X_1 \wedge \dots \wedge X_m) = \int_{S(\mathbf{U}_{\mathbf{p}}^0)} \mu_{\mathbf{p},\mathbf{v}}(X_1 \wedge \dots \wedge X_m) |dA_{S(\mathbf{U}_{\mathbf{p}}^0)}(\mathbf{v})|,$$

$\forall X_1, \dots, X_m \in T_{\mathbf{p}}M$ .

Clearly  $|d\mu_{\mathbf{U}}(\mathbf{p})|$  varies smoothly with  $\mathbf{p}$ , and thus it defines a density  $|d\mu_{\mathbf{U}}(-)|$  on  $M$ . We want to emphasize that this density *depends on the metric on  $\mathbf{U}$*  but *it is independent of any metric on  $M$* . We will refer to it as *the density of  $\mathbf{U}$* . By construction

$$\mathcal{N}(\mathbf{U}, h) = \frac{1}{\sigma_{n-1}} \int_M |d\mu_{\mathbf{U}}(\mathbf{p})|.$$

If we now return to our original metric  $g$  on  $M$ , then we can express  $|d\mu_{\mathbf{U}}(-)|$  as a product

$$|d\mu_{\mathbf{U}}(\mathbf{p})| = \delta_g(\mathbf{p}) \cdot |dV_g(\mathbf{p})|,$$

where  $\delta_g = \delta_{g,\mathbf{U}} : M \rightarrow \mathbb{R}$  is a smooth nonnegative function.

To find a more useful description of  $\rho_g$ , we choose local coordinates  $(x^1, \dots, x^m)$  near  $\mathbf{p}$  such that  $(\partial_{x^i})$  is a  $g$ -orthonormal basis of  $T_{\mathbf{p}}M$ . Then

$$\mu_{\mathbf{p},\mathbf{v}}(\partial_{x_1} \wedge \dots \wedge \partial_{x_m}) = \left| \det(\partial_{x_i x_j}^2 \mathbf{v}(\mathbf{p}))_{1 \leq i, j \leq m} \right| \cdot \left( \det((\mathcal{A}_{\mathbf{p}} \partial_{x_i}, \mathcal{A}_{\mathbf{p}} \partial_{x_j})_h)_{1 \leq i, j \leq m} \right)^{-1/2}.$$

Observe that the matrix  $(\partial_{x_i x_j}^2 \mathbf{v}(\mathbf{p}))_{1 \leq i, j \leq m}$  describes the Hessian operator

$$\text{Hess}_{\mathbf{p}}(\mathbf{v}, g) : T_{\mathbf{p}}M \rightarrow T_{\mathbf{p}}M$$

induced by the Hessian of  $\mathbf{v}$  at  $\mathbf{p}$  and the metric  $g$ .

The scalar  $(\det((\mathcal{A}_{\mathbf{p}} \partial_{x_i}, \mathcal{A}_{\mathbf{p}} \partial_{x_j})_h)_{1 \leq i, j \leq m})^{1/2}$  is precisely the Jacobian  $J_g(\mathbf{p})$  of the adjunction map  $\mathcal{A}_{\mathbf{p}} : T_{\mathbf{p}}M \rightarrow \mathbf{U}$  defined in terms of the metric  $g$  on  $T_{\mathbf{p}}M$  and the metric  $h$  on  $\mathbf{U}$ . We set

$$\Delta_{\mathbf{x}}(\mathbf{V}, g) := \int_{S(\mathbf{U}_{\mathbf{p}}^0)} |\det \text{Hess}_{\mathbf{x}}(\mathbf{v}, g)| |dA_{S(\mathbf{U}_{\mathbf{p}}^0)}(\mathbf{v})|.$$

Since  $|dV_g(\mathbf{p})|(\partial_{x_1} \wedge \dots \wedge \partial_{x_m}) = 1$ , we deduce

$$(2.12) \quad \delta_{g,\mathbf{V}}(\mathbf{p}) = \Delta_{\mathbf{p}}(\mathbf{V}, g) \cdot J_g(\mathbf{p})^{-1}.$$

This proves the first equality in (2.2). The second equality follows from the first by invoking (A.6) and the explicit formula  $(\sigma)$  for  $\sigma_{N-1}$ .  $\square$

**2.2. A Gaussian random field perspective.** For our concrete purposes it is convenient to give a probabilistic interpretation to the integral formula (2.2). For the reader's convenience we have gathered in Appendix A the basic probabilistic notions and facts needed in the sequel.

Consider again the metric  $\sigma = \sigma_U$ , the pullback of the metric  $h$  on  $U$  via the evaluation map. We will refer to it as the *stochastic metric* associated to the sample space  $(U, h)$ . It is convenient to have a local description of the stochastic metric.

Fix an orthonormal basis  $\psi_1, \dots, \psi_N$  of  $U$ . The evaluation map  $\mathbf{ev}^U : M \rightarrow U$  is then given by

$$M \ni \mathbf{x} \mapsto \sum_n \psi_n(\mathbf{x}) \cdot \psi_n \in U.$$

If  $\mathbf{p} \in M$  and  $\mathcal{U}$  is an open coordinate neighborhood of  $\mathbf{p}$  with coordinates  $x = (x^1, \dots, x^m)$ , then

$$(2.13) \quad \sigma_{\mathbf{p}}(\partial_{x^i}, \partial_{x^j}) = \sum_n \frac{\partial \psi_n}{\partial x^i}(\mathbf{p}) \frac{\partial \psi_n}{\partial x^j}(\mathbf{p}), \quad \forall 1 \leq i, j \leq m.$$

Note that if the collection  $(\partial_{x^i})_{1 \leq i \leq m}$  forms a  $g$ -orthonormal frame of  $T_{\mathbf{p}}M$ , then

$$(2.14) \quad J_g(\mathbf{p})^2 = \det \left[ \sigma_{\mathbf{p}}(\partial_{x^i}, \partial_{x^j}) \right]_{1 \leq i, j \leq m}.$$

To the sample space  $(U, h)$  we associate in a tautological fashion a Gaussian random field on  $M$  as follows. The measure  $d\gamma_h$  in (2.1) is a probability measure and thus  $(U, d\gamma_h)$  is naturally a probability space. We have a natural map

$$\xi : M \times U \rightarrow \mathbf{R}, \quad M \times U \ni (\mathbf{p}, \mathbf{u}) \mapsto \xi_{\mathbf{p}}(\mathbf{u}) := \mathbf{u}(\mathbf{p}).$$

The collection of random variables  $(\xi_{\mathbf{p}})_{\mathbf{p} \in M}$  is a Gaussian random field on  $M$ .

Using the orthonormal basis  $(\psi_k)$  of  $U$  we obtain a linear isometry

$$\mathbb{R}^N \ni \mathbf{t} = (t_1, \dots, t_n) \mapsto \mathbf{u}_{\mathbf{t}} = \sum_k t_k \psi_k \in U,$$

with inverse  $\mathbf{u} \mapsto t_k(\mathbf{u}) = h(\mathbf{u}, \psi_k)$ . For any  $\mathbf{p} \in M$  and any  $\mathbf{t} \in \mathbb{R}^N$  we have

$$\xi_{\mathbf{p}}(\mathbf{u}_{\mathbf{t}}) = \sum_k t_k \psi_k(\mathbf{p}).$$

The covariance kernel of this field is the function  $\mathcal{E} = \mathcal{E}_U : M \times M \rightarrow \mathbb{R}$  given by

$$(2.15) \quad \begin{aligned} \mathcal{E}(\mathbf{p}, \mathbf{q}) &= \mathbf{E}(\xi_{\mathbf{p}}, \xi_{\mathbf{q}}) = \sum_{j,k=1}^N \left( \int_{\mathbb{R}^N} t_j t_k d\gamma_N(\mathbf{t}) \right) \psi_j(\mathbf{p}) \psi_k(\mathbf{q}) \\ &= \sum_{k=1}^M \psi_k(\mathbf{p}) \psi_k(\mathbf{q}), \end{aligned}$$

where  $d\gamma_N$  is the canonical Gaussian measure on  $\mathbb{R}^N$ .

If  $\mathbf{p} \in M$  and  $\mathcal{U}$  is an open coordinate neighborhood of  $\mathbf{p}$  with coordinates  $x = (x^1, \dots, x^m)$  such that  $x(\mathbf{p}) = 0$ , then we can rewrite (2.13) in terms of the covariance kernel alone

$$(2.16) \quad \sigma_{\mathbf{p}}(\partial_{x^i}, \partial_{x^j}) = \frac{\partial^2 \mathcal{E}(x, y)}{\partial x^i \partial y^j} \Big|_{x=y=0}.$$

Note that any vector field  $X$  determines a new Gaussian random field on  $M$ , the derivative of  $\mathbf{u}$  along  $X$ . We obtain the Gaussian random variables  $\mathbf{u} \mapsto (X\mathbf{u})_{\mathbf{p}}$ ,  $\mathbf{u} \mapsto (Y\mathbf{u})_{\mathbf{p}}$ , and we have

$$(2.17) \quad \sigma_{\mathbf{p}}(X, Y) = \mathbf{E}((X\mathbf{u})_{\mathbf{p}}, (Y\mathbf{u})_{\mathbf{p}}).$$

The last equality justifies the attribute stochastic attached to the metric  $\sigma$ .

We denote by  $\nabla$  the Levi-Civita connection of the metric  $g$ . The Hessian of a smooth function  $f : M \rightarrow \mathbb{R}$  with respect to the metric  $g$  is the symmetric  $(0, 2)$ -tensor  $\nabla^2 f$  on  $M$  defined by the equality

$$(2.18) \quad \nabla^2 f(X, Y) := XYf - (\nabla_X Y)f, \quad \forall X, Y \in \text{Vect}(M).$$

If  $\mathbf{p}$  is a critical point of  $f$  then  $\nabla_{\mathbf{p}}^2 f$  is the usual Hessian of  $f$  at  $\mathbf{p}$ . More generally, if  $(x^1, \dots, x^m)$  are  $g$ -normal coordinates at  $\mathbf{p}$ , then

$$\nabla_{\mathbf{p}}^2 f(\partial_{x^i}, \partial_{x^j}) = \partial_{x^i x^j}^2 f(\mathbf{p}), \quad \forall 1 \leq i, j \leq m.$$

For any  $\mathbf{p} \in M$  and any  $f \in C^\infty(M)$  we use the metric  $g_{\mathbf{p}}$  to identify the bilinear form  $\nabla_{\mathbf{p}}^2 f$  on  $T_{\mathbf{p}}M$  with an element of  $\mathcal{S}(T_{\mathbf{p}}M)$ , the vector space of symmetric endomorphisms of the Euclidean space  $(T_{\mathbf{p}}M, g_{\mathbf{p}})$ . For any  $\mathbf{p} \in M$  we have two random Gaussian vectors

$$\mathbf{U} \ni \mathbf{u} \mapsto \nabla_{\mathbf{p}}^2 \mathbf{u} \in \mathcal{S}(T_{\mathbf{p}}M), \quad \mathbf{U} \ni \mathbf{u} \mapsto d\mathbf{u}(\mathbf{p}) \in T_{\mathbf{x}}^*M.$$

Note that the expectation of both random vectors are trivial while (2.16) shows that the covariance form of  $d\mathbf{u}(\mathbf{p})$  is the metric  $\sigma_{\mathbf{p}}$ .

To proceed further we need to make an additional assumption on the sample space  $\mathbf{U}$ . Namely, in the remainder of this section we will assume that it is 2-ample. In this case the map

$$\mathbf{U} \ni \mathbf{u} \mapsto \nabla_{\mathbf{p}}^2 \mathbf{u} \in \mathcal{S}(T_{\mathbf{p}}M)$$

is surjective so the Gaussian random vector  $\nabla_{\mathbf{p}}^2 \mathbf{u}$  is nondegenerate. A simple application of the co-area formula shows that the integral  $I_{\mathbf{p}}$  in (2.2) can be expressed as a conditional expectation

$$I_{\mathbf{p}} = \mathbf{E}(|\det \nabla_{\mathbf{p}}^2 \mathbf{u}| \mid d\mathbf{u}(\mathbf{p}) = 0).$$

Observing that

$$(2.19) \quad J_g(\mathbf{p}) = (\det \mathbf{S}_{d\mathbf{u}(\mathbf{p})})^{\frac{1}{2}},$$

we deduce that

$$(2.20) \quad \mathcal{N}(\mathbf{U}, h) = \frac{1}{(2\pi)^{\frac{m}{2}}} \int_M (\det \mathbf{S}_{d\mathbf{u}(\mathbf{p})})^{-\frac{1}{2}} \mathbf{E}(|\det \nabla_{\mathbf{p}}^2 \mathbf{u}| \mid d\mathbf{u}(\mathbf{p}) = 0) |dV_g(\mathbf{p})|.$$

The last equality is the main conclusion of the Expectation Metatheorem, [1, Thm. 11.2.1] or the expectation formula in [4, Thm. 6.2]. We can simplify the equality (2.20) even more by taking full advantage of the Gaussian nature of the various random variables involved in this equality.

The covariance form of the pair of random variables  $\nabla_{\mathbf{p}}^2 \mathbf{u}$  and  $d\mathbf{u}(p)$  is the bilinear map

$$\Omega : \mathcal{S}(T_{\mathbf{p}}M)^\vee \times T_{\mathbf{p}}M \rightarrow \mathbb{R},$$

$$\Omega(\xi, \eta) = \mathbf{E}(\langle \xi, \nabla_{\mathbf{p}}^2 \mathbf{u} \cdot \langle d\mathbf{u}, \eta \rangle \rangle), \quad \forall \xi \in \mathcal{S}_m^\vee, \quad \eta \in T_{\mathbf{p}}M.$$

Using the natural inner products on  $\mathcal{S}(T_{\mathbf{p}}M)$  and  $T_{\mathbf{p}}M$  defined by  $g_{\mathbf{p}}$  we can regard the covariance form as a linear operator

$$\Omega_{\mathbf{p}} : T_{\mathbf{p}}M \rightarrow \mathcal{S}(T_{\mathbf{p}}M).$$

Similarly, we can identify the covariance forms of  $\nabla_{\mathbf{p}}^2 u$  and  $du$  with symmetric positive definite operators

$$\mathbf{S}_{\nabla_{\mathbf{p}}^2 u} : \mathcal{S}(T_{\mathbf{p}}M) \rightarrow \mathcal{S}(T_{\mathbf{p}}M)$$

and respectively

$$\mathbf{S}_{d\mathbf{u}(p)} : T_{\mathbf{p}}M \rightarrow T_{\mathbf{p}}M.$$

Using the regression formula (A.3) we deduce that

$$(2.21) \quad \mathbf{E}(|\det \nabla_{\mathbf{p}}^2 \mathbf{u}| \mid d\mathbf{u}(p) = 0) = \mathbf{E}(|\det Y_{\mathbf{p}}|),$$

where  $Y_{\mathbf{p}} : \mathbf{U} \rightarrow \mathcal{S}(T_{\mathbf{p}}M)$  is a Gaussian random vector with mean value zero and covariance operator

$$(2.22) \quad \Xi_{\mathbf{p}} = \Xi_{Y_{\mathbf{p}}} := \mathbf{S}_{\nabla_{\mathbf{p}}^2 u} - \Omega \mathbf{S}_{d\mathbf{u}(p)}^{-1} \Omega^\dagger : \mathcal{S}(T_{\mathbf{p}}M) \rightarrow \mathcal{S}(T_{\mathbf{p}}M).$$

Since  $\mathbf{U}$  is 2-ample the operator  $\Xi_{\mathbf{p}}$  is invertible and we have

$$(2.23) \quad \mathbf{E}(|\det Y_{\mathbf{p}}|) = (2\pi)^{-\frac{\dim \mathcal{S}(T_{\mathbf{p}}M)}{2}} (\det \Xi_{\mathbf{p}})^{-\frac{1}{2}} \int_{\mathcal{S}(T_{\mathbf{p}}M)} |\det Y| e^{-\frac{(\Xi_{\mathbf{p}}^{-1} Y, Y)}{2}} dV_g(Y).$$

We deduce that *when  $\mathbf{U}$  is 2-ample* we have

$$(2.24) \quad \mathcal{N}(\mathbf{U}, h) = \frac{1}{(2\pi)^{\frac{m}{2}}} \int_M (\det \mathbf{S}_{d\mathbf{u}(p)})^{-\frac{1}{2}} \mathbf{E}(|\det Y_{\mathbf{p}}|) |dV_g(p)|,$$

where  $Y_{\mathbf{p}}$  is a Gaussian random symmetric endomorphism of  $T_{\mathbf{p}}M$  with expectation 0 and covariance operator  $\Xi_{\mathbf{p}}$  described by (2.22).

To compute the above integral we choose normal coordinates  $(x^1, \dots, x^n)$  near  $\mathbf{p}$  and thus we can orthogonally identify  $T_{\mathbf{p}}M$  with  $\mathbb{R}^m$ . We can view the Hessian  $\nabla_{\mathbf{p}}^2 \mathbf{u}$  as a random variable

$$H^{\mathbf{p}} : \mathbf{U} \rightarrow \mathcal{S}_m := \mathcal{S}(\mathbb{R}^m), \quad \mathbf{U} \ni \mathbf{u} \mapsto H^{\mathbf{p}}(\mathbf{u}) \in \mathcal{S}_m, \quad H_{ij}^{\mathbf{p}}(\mathbf{u}) = \partial_{x^i x^j}^2 \mathbf{u}(\mathbf{p}),$$

and the differential  $d\mathbf{u}(p)$  as a random variable

$$D^{\mathbf{p}} : \mathbf{U} \rightarrow \mathbb{R}^m, \quad \mathbf{u} \mapsto D^{\mathbf{p}}\mathbf{u} \in \mathbb{R}^m, \quad D_i^{\mathbf{p}}\mathbf{u} = \partial_{x^i} \mathbf{u}(\mathbf{p}).$$



The covariance operator  $S_{d\mathbf{u}(p)}$  of the random variable  $D\mathbf{p}$  is given by the symmetric  $m \times m$  matrix with entries

$$(2.25) \quad \sigma_{\mathbf{p}}(\partial_{x^i}, \partial_{x^j}) = \frac{\partial^2 \mathcal{E}(x, y)}{\partial x^i \partial y^j} \Big|_{x=y=0}.$$

To compute the covariance form  $\Sigma_{H\mathbf{p}}$  of the random matrix  $H\mathbf{p}$  we observe first that we have a canonical basis  $(\xi_{ij})_{1 \leq i \leq j \leq m}$  of  $\mathcal{S}_m^\vee$  so that  $\xi_{ij}$  associates to a symmetric matrix  $A$  the entry  $a_{ij}$  located in the position  $(i, j)$ . Then

$$(2.26) \quad \begin{aligned} \Sigma_{H\mathbf{p}}(\xi_{ij}, \xi_{k\ell}) &= \mathbf{E}(H_{ij}^{\mathbf{p}}(\mathbf{u}), H_{k\ell}^{\mathbf{p}}(\mathbf{u})) = \mathbf{E}(\partial_{x^i x^j}^2 \mathbf{u}(\mathbf{x}) \partial_{x^k x^\ell}^2 \mathbf{u}(\mathbf{x})) \\ &= \sum_{n=1}^N \partial_{x^i x^j}^2 \psi_n(\mathbf{x}) \partial_{x^k x^\ell}^2 \psi_n(\mathbf{x}) = \frac{\partial^4 \mathcal{E}(x, y)}{\partial x^i \partial x^j \partial y^k \partial y^\ell} \Big|_{x=y=0}. \end{aligned}$$

Similarly we have

$$(2.27) \quad \Omega(\xi_{ij}, \partial_{x^k}) = \mathbf{E}(\partial_{x^i x^j}^2 \mathbf{u}(\mathbf{p}), \partial_{x^k} \mathbf{u}(\mathbf{p})) = \frac{\partial^3 \mathcal{E}(x, y)}{\partial x^i \partial x^j \partial y^k} \Big|_{x=y=0}.$$

To identify  $\Omega$  with an operator it suffices to observe that  $(\partial_{x^k})$  is an orthonormal basis of  $T_{\mathbf{p}}M$ , while the collection  $\{\hat{\xi}_{ij}\}_{i \leq j} \subset \mathcal{S}_m^\vee$ ,

$$\hat{\xi}_{ij} = \begin{cases} \xi_{ij}, & i = j \\ \sqrt{2}\xi_{ij}, & i < j \end{cases}$$

is an orthonormal basis of  $\mathcal{S}_m^\vee$ . If we denote by  $\hat{E}_{ij}$  the dual orthonormal basis of  $\mathcal{S}_m$ , then

$$\Omega \partial_{x^k} = \sum_{i \leq j} \Omega(\hat{\xi}_{ij}, \partial_{x^k}) \hat{E}_{ij}.$$

REMARK 2.6. If the metric  $g$  coincides with the stochastic metric  $\sigma$ , then the covariance operator  $\Omega$  is trivial. For a proof of this and of many other nice properties of the metric  $\sigma$  we refer to [1, §12.2].  $\square$

**2.3. Zonal domains of spherical harmonics of large degree.** In the conclusion of this section we want to discuss an immediate application of the above results to critical sets of random spherical harmonics.

Let  $(M, g)$  be the unit round sphere  $S^2$ . The spectrum of the Laplacian on  $S^2$  is

$$\lambda_n = n(n + 1), \quad n = 0, 1, 2, \dots, \quad \dim \ker(\lambda_n - \Delta) = 2n + 1 = d_n.$$

The space  $\mathbf{U}_n = \ker(\lambda_n - \Delta)$  has a well known description: it consists of spherical harmonics, i.e., restrictions to  $S^2$  of harmonic polynomials of degree  $n$  in three variables. We want to describe the behavior of  $\mathcal{N}(\mathbf{U}_n)$  as  $n \rightarrow \infty$ , where  $\mathbf{U}_n$  is equipped with the  $L^2$ -metric. In other words we want to find the expected number of critical points of a spherical harmonic of very large degree.

In this case the covariance kernel  $\mathcal{E}_n(\mathbf{p}, \mathbf{q})$  of  $\mathbf{U}_n$  has a very simple description. More precisely, if  $(\Psi_k)_{1 \leq k \leq 2n+1}$  is an orthonormal basis of  $\mathbf{U}_n$ , then the classical addition theorem, [28, §1.2] shows that

$$\mathcal{E}_n(\mathbf{p}, \mathbf{q}) = \sum_k \Psi_k(\mathbf{p}) \Psi_k(\mathbf{q}) = \frac{2n + 1}{4\pi} P_n(\mathbf{p} \bullet \mathbf{q}), \quad \forall \mathbf{p}, \mathbf{q} \in S^2,$$

where  $\bullet$  denotes the inner product in  $\mathbb{R}^3$ , and  $P_n$  denotes the  $n$ -th Legendre polynomial,

$$P_n(t) = (-1)^n \frac{1}{2^n n!} \frac{d^n}{dt^n} (1 - t^2)^n.$$

In this case the stochastic metric  $\sigma = \sigma_n$  is obviously  $SO(3)$ -invariant and it is a (constant) multiple of the round metric. In view of Remark 2.6 this implies that for any  $\mathbf{p} \in S^2$  the random variables

$$\mathbf{U}_n \ni \mathbf{u} \mapsto \text{Hess}_{\mathbf{p}}(\mathbf{u}, g) \quad \text{and} \quad \mathbf{U}_n \ni \mathbf{u} \mapsto d\mathbf{u}(\mathbf{p})$$

are independent and we deduce that

$$\mathcal{N}(\mathbf{U}_n) = \frac{1}{2\pi} \int_{S^2} \frac{1}{J_g(\mathbf{p})} \left( \int_{\mathbf{U}_n} |\det \text{Hess}_{\mathbf{p}}(\mathbf{u}, g)| \underbrace{\frac{e^{-\frac{1}{2}|\mathbf{u}|^2}}{(2\pi)^{\frac{\dim \mathbf{U}_n}{2}} |d\mathbf{u}|}}_{=: d\gamma_n(\mathbf{u})} \right) dV_g(\mathbf{p}).$$

Clearly, the integrand in the above formula is invariant with respect to the  $SO(3)$ -action on  $S^2$  and we thus have

$$(2.28) \quad \mathcal{N}(\mathbf{U}_n) = \frac{2}{J_g(\mathbf{p}_0)} \int_{\mathbf{U}_n} |\det \text{Hess}_{\mathbf{p}_0}(\mathbf{u}, g)| d\gamma_n(\mathbf{u}),$$

where  $\mathbf{p}_0$  a fixed (but arbitrary) point on  $S^2$ . To compute the term in the right-hand side of the above equality we use the equalities (2.25) and (2.26).

Fix normal coordinates  $(x^1, x^2)$  in a neighborhood  $\mathcal{O}$  of  $\mathbf{p}_0$  so we can view  $\mathcal{E}_n$  as a function  $\mathcal{E}_n(x, y)$ . The location of a point  $\mathbf{p} \in \mathcal{O}$  is described by a smooth function

$$\mathcal{O} \ni (x^1, x^2) \mapsto \mathbf{p}(x^1, x^2) \in \mathbb{R}^3.$$

The tangent vector  $\partial_{x^i}$ , viewed as a vector in  $\mathbb{R}^3$ , corresponds with the derivative  $\mathbf{p}_{x^i} := \partial_{x^i} \mathbf{p}$  of the above function. At  $\mathbf{p}_0$  we have

$$(2.29) \quad \mathbf{p}_{x^i} \bullet \mathbf{p}_{x^j} = \delta_{ij} \quad \text{and} \quad \mathbf{p}_{x^i} \bullet \mathbf{p}_0 = 0, \quad \forall i, j.$$

The arcs  $C_1 = \{x^2 = 0\}$  and  $C_2 = \{x^1 = 0\}$  are portions of great circles intersecting orthogonally at  $\mathbf{p}_0$ . Note that  $x^1$  is the arclength parameter along  $C_i$ ,  $i = 1, 2$ . The vectors  $\mathbf{p}_{x^i}$  are unit tangent vectors along these arcs. This shows that at  $\mathbf{p}_0$  we have

$$\mathbf{p}_{x^i x^i} = -\mathbf{p}_0.$$

Since the arcs  $C_1$  and  $C_2$  are planar their torsion is trivial and the Frenet formulæ imply that at  $\mathbf{p}_0$  we have

$$\mathbf{p}_{x^i x^j} = 0, \quad \forall i \neq j.$$

The last two equalities can be rewritten in compact form as

$$(2.30) \quad \mathbf{p}_{x^i x^j} = -\delta_{ij} \mathbf{p}_0, \quad \forall i, j.$$

We set

$$(2.31) \quad \begin{aligned} s_n &:= \frac{2n+1}{\pi} P'_n(1) = \frac{2n+1}{4\pi} \times \frac{n(n+1)}{2} \sim \frac{1}{4\pi} n^3 \\ t_n &:= \frac{2n+1}{\pi} P''_n(1) = \frac{(2n+1)}{4\pi} \times \frac{(n+2)(n+1)n(n-1)}{16} \sim \frac{1}{32\pi} n^5. \end{aligned}$$

We deduce

$$(2.32) \quad \begin{aligned} \sigma(\partial_{x^j}, \partial_{x^k}) &= \partial_{x_j} \partial_{y^k} \mathcal{E}(\mathbf{p}, \mathbf{q})|_{\mathbf{p}=\mathbf{q}=\mathbf{p}_0} \\ &= \frac{2n+1}{4\pi} \left( P'_n(\mathbf{p} \bullet \mathbf{q}) \mathbf{p}_{x^j} \bullet \mathbf{q}_{y^k} + P_n^{(2)}(\mathbf{p} \bullet \mathbf{q})(\mathbf{p}_{x^j} \bullet \mathbf{q})(\mathbf{p} \bullet \mathbf{q}_{y^k}) \right)_{\mathbf{p}=\mathbf{q}=\mathbf{p}_0} \\ &= s_n \delta_{jk}, \end{aligned}$$

and

$$(2.33) \quad J_g(\mathbf{p}_0) = s_n.$$

To compute  $\partial_{x^i x^j y^k y^\ell}^4 \mathcal{E}_n(\mathbf{p}, \mathbf{q})$  at  $\mathbf{p} = \mathbf{q} = \mathbf{p}_0$  we will use (2.29) and (2.30) to cut down the complexity of the final formula. We deduce that at  $\mathbf{p} = \mathbf{q} = \mathbf{p}_0$  we have

$$\begin{aligned} &\partial_{x^i x^j y^k y^\ell}^4 \mathcal{E}_n(\mathbf{p}, \mathbf{q}) \\ &= \frac{2n+1}{4\pi} \left( P'_n(\mathbf{p} \bullet \mathbf{q}) \mathbf{p}_{x^i x^j} \bullet \mathbf{q}_{y^\ell y^k} + P_n^{(2)}(\mathbf{p} \bullet \mathbf{q})(\mathbf{p}_{x^i x^j} \bullet \mathbf{q})(\mathbf{p} \bullet \mathbf{q}_{y^k y^\ell}) \right)_{\mathbf{p}=\mathbf{q}} \\ &\quad + \frac{2n+1}{4\pi} \left( P_n^{(2)}(\mathbf{p} \bullet \mathbf{q})(\mathbf{p}_{x^i} \bullet \mathbf{q}_{y^\ell})(\mathbf{p}_{x^j} \bullet \mathbf{q}_{y^k}) + P_n^{(2)}(\mathbf{p} \bullet \mathbf{q})(\mathbf{p}_{x^j} \bullet \mathbf{q}_{y^\ell})(\mathbf{p}_{x^i} \bullet \mathbf{q}_{y^k}) \right)_{\mathbf{p}=\mathbf{q}}, \end{aligned}$$

and thus

$$(2.34) \quad \partial_{x^i x^j y^k y^\ell}^4 \mathcal{E}_n(\mathbf{p}, \mathbf{q})_{\mathbf{p}=\mathbf{q}} = (s_n + t_n) \delta_{ij} \delta_{kl} + t_n (\delta_{i\ell} \delta_{jk} + \delta_{ik} \delta_{j\ell}).$$

Denote by  $d\mathbf{\Gamma}_n$  the pushforward of the Gaussian measure  $d\gamma_n$  via the Hessian map

$$\mathbf{U}_n \ni \mathbf{u} \mapsto \text{Hess}_{\mathbf{p}_0}(\mathbf{u}, g) \in \mathcal{S}(T_{\mathbf{p}_0} S^2) = \mathcal{S}_2.$$

We deduce from (2.34) that the covariance form  $\Sigma_n$  of  $d\mathbf{\Gamma}_n$  satisfies the equality

$$\Sigma_n = \Sigma_{a_n, b_n, c_n}, \quad a_n = s_n + 3t_n, \quad b_n = s_n + t_n, \quad c_n = t_n,$$

where  $\Sigma_{a,b,c}$  is defined by the conditions (B.2a) and (B.2b). Observe that  $a_n, b_n, c_n$  satisfy (B.4), i.e.,  $a_n = b_n + 2c_n$ . As explained in Appendix B, this implies that  $d\mathbf{\Gamma}_n$  is  $O(2)$ -invariant. Set

$$a_n^* = \frac{a_n}{t_n}, \quad b_n^* = \frac{b_n}{t_n}, \quad c_n^* = \frac{c_n}{t_n},$$

and denote by  $d\mathbf{\Gamma}_n^*$  the Gaussian measure on  $\mathcal{S}_2$  with covariance matrix  $\Sigma_{a_n^*, b_n^*, c_n^*}$ . Using (A.7) we deduce that

$$\int_{\mathcal{S}_2} |\det X| d\mathbf{\Gamma}_n(X) = t_n \int_{\mathcal{S}_2} |\det X| d\mathbf{\Gamma}_n^*(X).$$

From (2.28) and (2.33) we now deduce

$$(2.35) \quad \mathcal{N}(\mathbf{U}_n) = \frac{2t_n}{s_n} \int_{\mathcal{S}_2} |\det X| d\mathbf{\Gamma}_n^*(X).$$

Observe that as  $n \rightarrow \infty$  we have

$$\frac{2t_n}{s_n} \sim \frac{n^2}{4}, \quad a_n^* \sim 3, \quad b_n^* \sim 1, \quad c_n^* \sim 1,$$

so that

$$(2.36) \quad \mathcal{N}(\mathbf{U}_n) \sim \frac{n^2}{4} \int_{\mathbb{S}_2} |\det X| d\mathbf{\Gamma}_{3,1,1}(X),$$

where  $d\mathbf{\Gamma}_{3,1,1}(X)$  is the Gaussian measure on  $\mathbb{S}_2$  with covariance form  $\mathbf{\Sigma}_{3,1,1}$ . More precisely (see (B.11))

$$d\mathbf{\Gamma}_{3,1,1}(X) = \frac{1}{4(2\pi)^{\frac{3}{2}}} e^{-\frac{1}{4}(\text{tr } X^2 - \frac{1}{4}(\text{tr } X)^2)} \cdot \sqrt{2} \prod_{1 \leq i \leq j \leq 2} dx_{ij}.$$

In Appendix C we show that

$$(2.37) \quad \int_{\mathbb{S}_2} |\det X| d\mathbf{\Gamma}_{3,1,1}(X) = \frac{4}{\sqrt{3}},$$

and we deduce from (2.28) that

$$(2.38) \quad \mathcal{N}(\mathbf{U}_n) \sim \frac{n^2}{\sqrt{3}} \text{ as } n \rightarrow \infty.$$

Let us observe that for  $n$  very large, a typical spherical harmonic  $\mathbf{u} \in \mathbf{U}_n$  is a Morse function on  $S^2$  and 0 is a regular value. The nodal set  $\{\mathbf{u} = 0\}$  is disjoint union of smoothly embedded circles. We denote by  $\mathcal{D}_{\mathbf{u}}$  the set of connected components of the complement of the nodal set are called the *nodal domains* of  $\mathbf{u}$  and we denote  $\delta(\mathbf{u})$  the cardinality of  $\mathcal{D}_{\mathbf{u}}$ . A result of Pleijel and Peetre, [6, 36, 39], shows that

$$(2.39) \quad \delta(\mathbf{u}) \leq \frac{4}{j_0^2} n^2 \approx 0.692n^2,$$

where  $j_0$  denotes the first positive zero of the Bessel function  $J_0$ .

We think of  $\delta(\mathbf{u})$  as a random variable and we denote by  $\delta_n$  its expectation,

$$\delta_n = \frac{1}{(2\pi)^{\dim \mathbf{U}_n} 2} \int_{\mathbf{U}_n} \delta(\mathbf{u}) e^{-\frac{1}{2}|\mathbf{u}|^2} |d\mathbf{u}|.$$

Denote by  $p(\mathbf{u})$  the number of local minima and maxima of  $\mathbf{u}$ , and by  $s(\mathbf{u})$  the number of saddle points. Then

$$\mathcal{N}(\mathbf{y}) = p(\mathbf{u}) + s(\mathbf{u}), \quad p(\mathbf{u}) - s(\mathbf{u}) = \chi(S^2) = 2.$$

This proves that

$$p(\mathbf{y}) = \frac{1}{2}(\mathcal{N}(\mathbf{u}) + 2).$$

For every nodal region  $D$ , we denote by  $p(\mathbf{u}, D)$  the number of local minima and maxima<sup>3</sup> of  $\mathbf{u}$  on  $D$ . Note that  $p(\mathbf{u}, D) > 0$  for any  $D$  and thus the number  $p(\mathbf{u}) = \sum_{D \in \mathcal{D}_{\mathbf{u}}} p(\mathbf{u}, D)$  can be viewed as a weighted count of nodal domains. Moreover

$$\delta(\mathbf{u}) \leq p(\mathbf{u}).$$

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<sup>3</sup>A simple application of the maximum principle shows that on each nodal domain, all the local extrema of  $\mathbf{y}$  are of the same type: either all local minima or all local maxima. Thus  $p(\mathbf{u}, D)$  can be visualized as the number of peaks of  $|\mathbf{u}|$  on  $D$ .

We set

$$p(\mathbf{U}_n) := \frac{1}{(2\pi)^{\frac{\dim \mathbf{U}_n}{2}}} \int_{\mathbf{U}_n} e^{-\frac{1}{2}|\mathbf{u}|^2} p(\mathbf{u}) |d\mathbf{u}|.$$

The equality (2.38) implies that

$$p(\mathbf{U}_n) \sim \frac{1}{2\sqrt{3}}n^2 \text{ as } n \rightarrow \infty, \quad \frac{1}{2\sqrt{3}} \approx 0.288.$$

This shows that while

$$\max_{\mathbf{u} \in \mathbf{U}_n} \delta(\mathbf{u}) \leq 0.692n^2,$$

the expectation  $\delta_n$  is less than half this theoretical maximum,

$$\delta_n \approx 0.288.$$

Recently, Nazarov and Sodin [30], have proved that there exists a positive constant  $a > 0$  such that

$$\delta_n \sim an^2 \text{ as } n \rightarrow \infty.$$

Additionally, for large  $n$ , with high probability,  $\delta(\mathbf{u})$  is close to  $an^2$  (see [30] for a precise statement). This shows that

$$(2.40) \quad a \leq \frac{1}{2\sqrt{3}} \approx 0.288.$$

More information about lower bounds on  $a$  can be found in Maria Năstăsescu’s senior thesis [29].

### 3. The proof of Theorem 1.1.

**3.1. Asymptotic estimates of the spectral function.** We fix an orthonormal basis of  $L^2(M, g)$  consisting of eigenfunctions  $\Psi_n$  of  $\Delta_g$ ,

$$\Delta_g \Psi_n = \lambda_n \Psi_n, \quad n = 0, 1, \dots, \quad \lambda_0 \leq \lambda_1 \leq \dots \leq \lambda_n \leq \dots.$$

The collection  $(\Psi_n)_{\lambda_n \leq L}$  is therefore an orthonormal basis of  $\mathbf{U}_L$  so that the covariance kernel of the Gaussian field determined by  $\mathbf{U}_L$  is

$$\mathcal{E}_L(\mathbf{p}, \mathbf{q}) = \sum_{\lambda_n \leq L} \Psi_n(\mathbf{p}) \Psi_n(\mathbf{q}).$$

This function is also known as the *spectral function* associated to the Laplacian. Equivalently,  $\mathcal{E}_L$  can be identified with the Schwartz kernel of the orthogonal projection onto  $\mathbf{U}_L$ . Observe that

$$\int_M \mathcal{E}_L(\mathbf{p}, \mathbf{p}) |dV_g(\mathbf{p})| = \dim \mathbf{U}_L.$$

In the groundbreaking work [21], L. Hörmander used the kernel of the wave group  $e^{it\sqrt{\Delta}}$  to produce refined asymptotic estimates for the spectral function. More precisely he showed (see [21] or [23, §17.5])

$$(3.1) \quad \mathcal{E}_L(p, p) = \frac{\omega_m}{(2\pi)^m} L^{\frac{m}{2}} + O(L^{\frac{m-1}{2}}) \text{ as } L \rightarrow \infty,$$

uniformly with respect to  $\mathbf{p} \in M$ . Above,  $\omega_m$  denotes the volume of the unit ball in  $\mathbb{R}^m$ . This implies immediately the classical Weyl estimates

$$(3.2) \quad \dim \mathbf{U}_L \sim \frac{\omega_m}{(2\pi)^m} \text{vol}_g(M) L^{\frac{m}{2}}.$$

Hörmander’s approach can be refined to produce asymptotic estimates for the behavior of the derivatives the spectral function in a neighborhood of the diagonal. We describe below these estimates following closely the presentation in [7]. For more general results we refer to [41, Thm. 1.8.5, 1.8.7].

We set  $\lambda := L^{\frac{1}{2}}$ . Fix a point  $\mathbf{p}$  and normal coordinates  $x = (x^1, \dots, x^m)$  at  $\mathbf{p}$ . Note that  $x(\mathbf{p}) = 0$ . For any multi-indices  $\alpha, \beta \in \mathbb{Z}_{\geq 0}^m$  we have (see [7, Thm. 1.1, Prop. 2.3])

$$(3.3) \quad \frac{\partial^{\alpha+\beta} \mathcal{E}_L(x, y)}{\partial x^\alpha \partial y^\beta} \Big|_{x=y=0} = C_m(\alpha, \beta) \lambda^{m+|\alpha|+|\beta|} + O(\lambda^{m+|\alpha|+|\beta|-1}),$$

where

$$(3.4) \quad C_m(\alpha, \beta) = \begin{cases} 0, & \alpha - \beta \notin (2\mathbb{Z})^m, \\ \frac{(-1)^{\frac{|\alpha|-|\beta|}{2}}}{(2\pi)^m} \int_{\mathbf{B}^m} \mathbf{x}^{\alpha+\beta} |d\mathbf{x}|, & \alpha - \beta \in (2\mathbb{Z})^m, \end{cases}$$

and  $\mathbf{B}^m$  denotes the unit ball

$$\mathbf{B}^m = \{ \mathbf{x} \in \mathbb{R}^m; |\mathbf{x}| = 1 \}.$$

The estimates (3.3) are uniform in  $\mathbf{p} \in M$ . Using (A.6) we deduce (compare with (B.13) )

$$\frac{1}{(2\pi)^m} \int_{\mathbf{B}^m} \mathbf{x}^{\alpha+\beta} |d\mathbf{x}| = \frac{1}{(4\pi)^{\frac{m}{2}} \Gamma(1 + \frac{|\alpha|+|\beta|+m}{2})} \int_{\mathbb{R}^m} \mathbf{x}^{\alpha+\beta} \frac{e^{-|\mathbf{x}|^2}}{\pi^{\frac{m}{2}}} |d\mathbf{x}|.$$

We set

$$K_m = C_m(\alpha, \alpha), \quad |\alpha| = 1,$$

so that

$$(3.5) \quad K_m = \frac{1}{(4\pi)^{\frac{m}{2}} \Gamma(2 + \frac{m}{2})} \int_{\mathbb{R}^m} x_1^2 \frac{e^{-|\mathbf{x}|^2}}{\pi^{\frac{m}{2}}} |d\mathbf{x}| = \frac{1}{2(4\pi)^{\frac{m}{2}} \Gamma(2 + \frac{m}{2})}.$$

For any  $i \leq j$  define  $\alpha_{ij} \in \mathbb{Z}^m$  so that

$$\mathbf{x}^{\alpha_{ij}} = x_i x_j.$$

For  $i \leq j$  and  $k \leq \ell$  we set

$$(3.6) \quad C_m(i, j; k, \ell) = C_m(\alpha_{ij}, \alpha_{k\ell}) = \frac{1}{(4\pi)^{\frac{m}{2}} \Gamma(3 + \frac{m}{2})} \int_{\mathbb{R}^m} x_i x_j x_k x_\ell \frac{e^{-|\mathbf{x}|^2}}{\pi^{\frac{m}{2}}} |d\mathbf{x}|.$$

For  $i < j$  we have

$$(3.7) \quad C_m(i, i; j, j) = \frac{1}{(4\pi)^{\frac{m}{2}} \Gamma(3 + \frac{m}{2})} \int_{\mathbb{R}^m} x_i^2 x_j^2 \frac{e^{-|\mathbf{x}|^2}}{\pi^{\frac{m}{2}}} |d\mathbf{x}| = \frac{1}{4(4\pi)^{\frac{m}{2}} \Gamma(3 + \frac{m}{2})} =: c_m.$$

$$C_m(i, j; i, j) = C_m(i, i; j, j),$$

Finally

$$(3.8) \quad C_m(i, i; i, i) = \frac{1}{(4\pi)^{\frac{m}{2}} \Gamma(3 + \frac{m}{2})} \int_{\mathbb{R}^m} x_i^4 \frac{e^{-|x|^2}}{\pi^{\frac{m}{2}}} |d\mathbf{x}| = \frac{3}{4(4\pi)^{\frac{m}{2}} \Gamma(3 + \frac{m}{2})} = 3c_m,$$

and

$$C_m(i, j; k, \ell) = 0, \quad \forall k \leq \ell, \quad (i, j) \neq (k, \ell).$$

**3.2. Probabilistic consequences of the previous estimates.** We denote by  $\sigma^L$  the stochastic metric on  $M$  determined by the sample space  $\mathbf{U}_L$ ,  $L \gg 0$ . As explained in Subsection 2.2 the covariance form of the random vector  $\mathbf{U}_L \ni \mathbf{u} \mapsto d\mathbf{u}(\mathbf{p}) \in T_{\mathbf{p}}^*M$  is  $\sigma_{\mathbf{p}}^L$ , and from (3.3) we deduce

$$(3.9) \quad \begin{aligned} \sigma_{\mathbf{p}}^L(\partial_{x^i}, \partial_{x^j}) &= \frac{\partial^2 \mathcal{E}_L(x, y)}{\partial x^i \partial y^j} \Big|_{x=y=0} = K_m \lambda^{m+2} \delta_{ij} + O(\lambda^{m+1}) \\ &= K_m \lambda^{m+2} g_{\mathbf{p}}(\partial_{x^i}, \partial_{x^j}) + O(\lambda^{m+1}) \quad \text{as } L \rightarrow \infty, \quad \text{uniformly in } \mathbf{p}. \end{aligned}$$

In particular, if  $\mathbf{S}_{d\mathbf{u}(\mathbf{p})}^L$  denotes the covariance operator of the random vector  $d\mathbf{u}(\mathbf{p})$ , then we deduce from the above equality that

$$(3.10) \quad \mathbf{S}_{d\mathbf{u}(\mathbf{p})}^L = K_m \lambda^{m+2} \mathbb{1}_m + O(\lambda^{m+1}), \quad \text{uniformly in } \mathbf{p},$$

and invoking (2.19) we deduce

$$(3.11) \quad J_g^L(\mathbf{p}) = (\det \mathbf{S}_{d\mathbf{u}(\mathbf{p})}^L)^{\frac{1}{2}} = K_m^{\frac{m}{2}} \lambda^{\frac{m(m+2)}{2}} + O(\lambda^{\frac{m(m+2)}{2}-1}), \quad \text{uniformly in } \mathbf{p}.$$

Denote by  $\Sigma_{H\mathbf{p}}^L$  the covariance form of the random matrix

$$\mathbf{U}_L \ni \mathbf{u} \mapsto \nabla_{\mathbf{p}}^2 \mathbf{u} \in \mathcal{S}(T_{\mathbf{p}}M) = \mathcal{S}_m.$$

Using (2.26) and (3.3) we deduce

$$(3.12) \quad \Sigma_{H\mathbf{p}}^L = c_m \lambda^{m+4} \Sigma_{3,1,1} + O(\lambda^{m+3}), \quad \text{uniformly in } \mathbf{p},$$

where the positive definite, symmetric bilinear form  $\Sigma_{3,1,1} : \mathcal{S}_m^\vee \times \mathcal{S}_m^\vee \rightarrow \mathbb{R}$  is described by the equalities (B.2a) and (B.2b). We denote by  $\Gamma_{3,1,1}$  the centered Gaussian measure on  $\mathcal{S}_m$  with covariance form  $\Sigma_{3,1,1}$ .

The equality (2.27) coupled with (3.3) imply that the covariance operator  $\Omega_{\mathbf{p}}^L$  satisfies

$$(3.13) \quad \Omega_{\mathbf{p}}^L = O(\lambda^{m+2}), \quad \text{uniformly in } \mathbf{p}.$$

Using (3.10), (3.12) and (3.13) we deduce that the covariance operator  $\Xi_{\mathbf{p}}^L$  defined as in (2.22) satisfies the estimate

$$(3.14) \quad \Xi_{\mathbf{p}}^L = c_m \lambda^{m+4} \widehat{Q}_{3,1,1} + O(\lambda^{m+2}), \quad \text{as } L \rightarrow \infty, \quad \text{uniformly in } \mathbf{p},$$

where  $\widehat{Q}_{3,1,1}$  is the covariance operator associated to the covariance form  $\Sigma_{3,1,1}$  and it is described explicitly in (B.3). If we denote by  $d\Gamma_L$  the Gaussian measure on  $\mathcal{S}_m$  with covariance operator  $\Xi_{\mathbf{p}}^L$ , we deduce that

$$d\Gamma_L(Y) = \frac{1}{(2\pi)^{\frac{N_m}{2}} (\det \Xi_{\mathbf{p}}^L)^{\frac{1}{2}}} e^{-\frac{(\Xi_{\mathbf{p}}^L Y, Y)}{2}} \cdot \underbrace{2^{\frac{1}{2}} \binom{m}{2} \prod_{i \leq j} dy_{ij}}_{|dY|},$$

where

$$N_m = \dim \mathcal{S}_m = \frac{m(m+1)}{2}.$$

Let us observe that  $|dY|$  is the Euclidean volume element on  $\mathcal{S}_m$  defined by the natural inner product on  $\mathcal{S}_m$ ,  $(X, Y) = \text{tr}(XY)$ . We set

$$c_L := c_m \lambda^{m+4}, \quad Q_{\mathbf{p}}^L = \frac{1}{c_L} \Xi_{\mathbf{p}}^L.$$

Using (A.7) we deduce that

$$\begin{aligned} & \frac{1}{(2\pi)^{\frac{N_m}{2}} (\det \Xi_{\mathbf{p}}^L)^{\frac{1}{2}}} \int_{\mathcal{S}_m} |\det Y| e^{-\frac{(\Xi_{\mathbf{p}}^L Y, Y)}{2}} |dY| \\ &= \frac{(c_L)^{\frac{m}{2}}}{(2\pi)^{\frac{N_m}{2}} (\det Q_{\mathbf{p}}^L)^{\frac{1}{2}}} \int_{\mathcal{S}_m} |\det Y| e^{-\frac{(Q_{\mathbf{p}}^L Y, Y)}{2}} |dY|. \end{aligned}$$

From the estimate (3.14) we deduce that

$$Q_{\mathbf{p}}^L \rightarrow \widehat{Q}_{3,1,1} \text{ as } L \rightarrow \infty, \text{ uniformly in } \mathbf{p}.$$

We conclude that

$$(3.15) \quad \mathbf{E}(|\det Y_{\mathbf{p}}|) = \int_{\mathcal{S}_m} |\det Y| d\Gamma_L(Y) \sim c_m^{\frac{m}{2}} \lambda^{\frac{m(m+4)}{2}} \int_{\mathcal{S}_m} |\det Y| d\Gamma_{3,1,1}(Y).$$

The measure  $d\Gamma_{3,1,1}$  is described explicitly in (B.11), more precisely

$$d\Gamma_{3,1,1}(Y) = \frac{1}{(2\pi)^{\frac{N_m}{2}} \sqrt{\mu_m}} \cdot e^{-\frac{1}{4}(\text{tr } Y^2 - \frac{1}{m+2}(\text{tr } Y)^2)} |dY|,$$

where  $\mu_m$  is given by (B.12). Using (2.24), (3.11) and (3.15) we deduce that

$$\begin{aligned} \mathbf{E}(N_L) &\sim \left(\frac{c_m}{K_m}\right)^{\frac{m}{2}} \lambda^{\frac{m(m+4)}{2} - \frac{m(m+2)}{2}} \text{vol}_g(M) \int_{\mathcal{S}_m} |\det Y| d\Gamma_{3,1,1}(Y) \\ &\stackrel{(3.2)}{\sim} \left(\frac{c_m}{K_m}\right)^{\frac{m}{2}} \frac{(2\pi)^m}{\omega_m} \dim \mathbf{U}_L. \end{aligned}$$

Observe that

$$\frac{c_m}{K_m} = \frac{\Gamma(2 + \frac{m}{2})}{2\Gamma(3 + \frac{m}{2})} = \frac{1}{m+4}, \quad \omega_m = \frac{\pi^{\frac{m}{2}}}{\Gamma(1 + \frac{m}{2})} \frac{(2\pi)^m}{\omega_m} = (4\pi)^{\frac{m}{2}} \Gamma\left(1 + \frac{m}{2}\right).$$

This completes the proof of (1.1) and (1.7).  $\square$



**3.3. On the asymptotic behavior of the stochastic metric.** We denote by  $g(L)$  the metric

$$g(L) := \lambda^{-(m+2)} \sigma^L = L^{-\frac{(m+2)}{2}} \sigma^L,$$

where  $K_m$  is described by (3.5). The estimate (3.9) shows that

$$g(L) \xrightarrow{C^0} g \text{ as } L \rightarrow \infty,$$

where  $K_m$  is described by (3.5). The metrics  $g(L)$  are closely related to the metrics constructed in [5, Thm. 5]. We want to discuss here possible ways to improve the topology of the convergence.

Observe that if  $g(L)$  were to converge in the  $C^2$ -topology to  $g$ , then the sectional curvatures of  $g(L)$  would have to be uniformly bounded. Conversely, the results of S. Peters [37] show that the  $C^0$  convergence coupled with an uniform bound on the sectional curvatures would yield a  $C^{1,\alpha}$  convergence.

The results in [1, §12.2.1] describe a simple way of expressing the sectional curvatures of  $\sigma^L$  in terms of the spectral function  $\mathcal{E}_L$ . Here are the details.

Denote by  $\nabla^L$  the Levi-Civita connection of the metric  $\sigma^L$ . Fix a point  $\mathbf{p} \in M$  and  $g$ -normal coordinates  $(x^1, \dots, x^m)$  at  $\mathbf{p}$ . We set

$$\mathcal{E}_{i_1, \dots, i_a; j_1, \dots, j_b}^L := \frac{\partial^{a+b} \mathcal{E}_L(x, y)}{\partial x^{i_1} \dots \partial x^{i_a} \partial y^{j_1} \dots \partial y^{j_b}} \Big|_{x=y=0},$$

$$\sigma(L)_{ij} := \sigma_{\mathbf{p}}^L(\partial_{x^i}, \partial_{x^j}), \quad 1 \leq i, j \leq m,$$

and we denote by  $(\sigma(L)^{ij})_{1 \leq i, j \leq m}$  the inverse matrix of  $(\sigma(L)_{ij})_{1 \leq i, j \leq m}$ . From [1, Eq. (12.2.6)] we deduce

$$\Gamma(L)_{ijk} := \sigma_{\mathbf{p}}^L(\nabla_{\partial_{x^i}}^L \partial_{x^j}, \partial_{x^k}) = \mathcal{E}_{ij;k}^L.$$

We set

$$\Gamma(L)_{ij}^k := \sum_{\ell} \sigma(L)^{k\ell} \Gamma(L)_{ij\ell} = \sum_{\ell} \sigma(L)^{k\ell} \mathcal{E}_{ij;\ell}^L,$$

so that

$$(\nabla_{\partial_{x^i}}^L \partial_{x^j})_{\mathbf{p}} = \sum_k \Gamma(L)_{ij}^k \partial_{x^k}.$$

For  $\mathbf{u} \in U_L$  we set

$$(3.16) \quad H_{ij}^L(\mathbf{u}) := (\partial_{\partial_{x^i}} \partial_{x^j} \mathbf{u} - (\nabla_{\partial_{x^i}}^L \partial_{x^j})_{\mathbf{p}} \mathbf{u})_{\mathbf{p}} = \partial_{\partial_{x^i}} \partial_{x^j} \mathbf{u}(\mathbf{p}) - \sum_k \Gamma(L)_{ij}^k \partial_{x^k} \mathbf{u}(\mathbf{p}).$$

We think of the matrix  $H_{ij}^L(\mathbf{u})$  as an element  $H^L(\mathbf{u}) \in T_{\mathbf{p}}^* M \otimes T_{\mathbf{p}}^* M$ ,

$$H^L(\mathbf{u}) = \sum_{i,j} H_{ij}^L dx^i \otimes dx^j$$

and we set

$$\begin{aligned} H^L(\mathbf{u}) \wedge H^L(\mathbf{u}) &:= \sum_{i,j,k,\ell} H_{ij}^L(\mathbf{u})H_{k\ell}^L(\mathbf{u})dx^i \wedge dx^k \otimes dx^j \wedge dx^\ell \\ &=: \sum_{i < k, j < \ell} Q_{ikj\ell}^L(\mathbf{u})dx^i \wedge dx^k \otimes dx^j \wedge dx^\ell. \end{aligned}$$

Note that

$$Q_{ikj\ell}^L(\mathbf{u}) = 2(H_{ij}^L(\mathbf{u})H_{k\ell}^L(\mathbf{u}) - H_{kj}^L(\mathbf{u})H_{i\ell}^L(\mathbf{u})).$$

We denote by  $R^L$  the Riemann tensor of  $\sigma^L$  and we set

$$R_{ijk\ell}^L := \sigma^L(R^L(\partial_{x^i}, \partial_{x^j})\partial_{x^k}, \partial_{x^\ell})_{\mathbf{p}}.$$

The map  $U_L \ni \mathbf{u} \mapsto Q_{ikj\ell}^L(\mathbf{u}) \in \mathbb{R}$  is a random variable and according to [1, Lemma 12.2.1] we have<sup>4</sup>

$$(3.17) \quad 2R_{ikj\ell}^L = -\mathbf{E}(Q_{ikj\ell}^L).$$

In particular we deduce that

$$-R_{ijij}^L = \mathbf{E}(H_{ii}^L H_{jj}^L - (H_{ij}^L)^2).$$

From (3.9) we deduce that

$$\sigma(L)_{ij} = \varepsilon_{ij}^L \sim K_m \lambda^{m+2} \delta_{ij} + O(\lambda^{m+1}) \text{ as } L \rightarrow \infty.$$

Hence

$$\sigma(L)^{ij} \sim \frac{1}{K_m \lambda^{m+2}} (\delta^{ij} + O(\lambda^{-1})).$$

From (3.3) we deduce that as  $\lambda \rightarrow \infty$  we have

(3.18a)

$$\Gamma(L)_{ij}^k \sim \sum_{\ell} \frac{1}{K_m \lambda^{m+2}} (\delta^{k\ell} + O(\lambda^{-1})) \varepsilon_{ij;\ell}^L \sim \frac{1}{K_m \lambda^{m+2}} \varepsilon_{ij;k}^L + O(\lambda^{-1}) = O(1),$$

(3.18b)

$$\mathbf{E}(\partial_{x^i x^j}^2 \mathbf{u}(\mathbf{p}), \partial_{x^k} \mathbf{u}(\mathbf{p})) = \varepsilon_{ij;k}^L = O(\lambda^{m+2})$$

Using the estimates (2.26), (3.16), (3.18a) and (3.18b) in (3.17) we deduce

$$\mathbf{E}(H_{ii}^L H_{jj}^L - (H_{ij}^L)^2) = (\varepsilon_{ii;jj}^L - \varepsilon_{ij;ij}^L) + O(\lambda^{m+2}).$$

We deduce that the sectional curvature of  $\sigma^L$  along the plane spanned by  $\partial_{x^i}, \partial_{x^k}$  is

$$K_{ij}^L = -\frac{R_{ijij}}{\sigma(L)_{ii}\sigma(L)_{jj} - \sigma(L)_{ij}^2} = \frac{1}{K_m^2 \lambda^{2m+4}} (\varepsilon_{ii;jj}^L - \varepsilon_{ij;ij}^L) + O\left(\frac{\varepsilon_{ii;jj}^L - \varepsilon_{ij;ij}^L}{\lambda^{2m+5}}\right).$$

---

<sup>4</sup>Alternatively, in our case, the equalities (3.17) are simple consequences of Theorema Egregium, [31, §4.2.4, Eq. (4.2.12)].

On the other hand

$$\mathbf{E}(\partial_{x^i x^j}^2 \mathbf{u}(\mathbf{p}), \partial_{x^k x^\ell}^2 \mathbf{u}(\mathbf{p})) = \mathcal{E}_{ij; k\ell}^L \sim C_m(i, j; k, \ell) \lambda^{m+4} + O(\lambda^{m+3}), \quad i \leq j, \quad k \leq \ell,$$

where  $C_m(i, j; k, \ell)$  is defined by (3.6), and we deduce

$$(3.19) \quad \mathcal{E}_{ii; jj}^L - \mathcal{E}_{ij; ij}^L = (C_m(i, i; j, j) - C_m(i, j; i, j)) \lambda^{m+4} + O(\lambda^{m+3}) = O(\lambda^{m+3}).$$

Hence

$$K_{ij}^L = \frac{1}{K_m^2 \lambda^{2m+4}} (\mathcal{E}_{ii; jj}^L - \mathcal{E}_{ij; ij}^L) + O(\lambda^{-m-2}).$$

The sectional curvature of  $g(L) = \lambda^{-m-2} \sigma^L$  along the plane spanned by  $\partial_{x^i}, \partial_{x^j}$  is

$$\overline{K}_{ij}^L = \lambda^{m+2} K_{ij}^L = \frac{1}{K_m^2 \lambda^{m+2}} (\mathcal{E}_{ii; jj}^L - \mathcal{E}_{ij; ij}^L) + O(1).$$

We deduce that the sectional curvatures of  $g(L)$  are uniformly bounded if and only if

$$(3.20) \quad \mathcal{E}_{ii; jj}^L - \mathcal{E}_{ij; ij}^L = O(\lambda^{m+2}) \quad \text{uniformly over } M.$$

Note that the estimates (3.20) are stronger than the estimates (3.19) which are direct consequences of the Bin-Hörmander estimates (3.3).

Let us point out that (3.20) hold when  $(M, g)$  is a homogeneous space equipped with an invariant metric. Indeed, in this case the metric  $g(L)$  has the same symmetries as  $g$  and thus there exists a constant  $c_L > 0$  such that  $g(L) = c_L g$ . Then  $c_L \rightarrow K_m$  as  $L \rightarrow \infty$  so that  $g(L) \rightarrow K_m g$  in the  $C^\infty$ -topology and therefore  $\overline{K}_{ij}^L = O(1)$ .

#### 4. The proof of Theorem 1.2.

**4.1. Reduction to the classical Gaussian orthogonal ensemble.** We begin by describing the large  $m$  behavior of the integral

$$I_m := \frac{1}{(2\pi)^{\frac{m(m+1)}{4}} \sqrt{\mu_m}} \int_{S_m} |\det X| e^{-\frac{1}{4}(\text{tr } X^2 - \frac{1}{m+2}(\text{tr } X)^2)} |dX|,$$

where we recall that

$$\mu_m = 2^{\binom{m}{2} + m - 1} (m + 2).$$

We will use a trick of Fyodorov [17]; see also [16, §1.5]. Recall first the classical equality

$$\int_{\mathbb{R}} e^{-(at^2 + bt + c)} |dt| = \left(\frac{\pi}{a}\right)^{\frac{1}{2}} e^{\frac{\Delta}{4a}}, \quad \Delta = b^2 - 4ac, \quad a > 0.$$

For any real numbers  $u, v, w$ , we have

$$\begin{aligned} ut^2 + v \text{tr}(X + wt \mathbb{1}_m)^2 &= (u + mw^2)t^2 + 2vw(\text{tr } X)t + v \text{tr } X^2 \\ &=: a(u, v, w)t^2 + b(u, v, w)t + c(u, v, w). \end{aligned}$$

We seek  $u, v, w$  such that

$$\frac{v^2 w^2}{u + mw^2} (\text{tr } X)^2 - \frac{v}{u + mw^2} \text{tr } X^2 = \frac{b^2 - 4ac}{4a} = -\frac{1}{4} \left( \text{tr } X^2 - \frac{1}{m+2} (\text{tr } X)^2 \right).$$

We have

$$\frac{v}{u + mw^2} = \frac{1}{4}, \quad \frac{v^2 w^2}{u + mw^2} = \frac{1}{4(m+2)},$$

and we deduce

$$vw^2 = \frac{1}{(m+2)}, \quad v = \frac{1}{4}(u + mw^2) \iff u = 4v - mw^2.$$

Hence

$$w^2 = \frac{1}{v(m+2)}, \quad u = 4v - \frac{m}{v(m+2)}.$$

We choose  $v = \frac{1}{2}$  so that

$$w^2 = \frac{2}{(m+2)}, \quad u = 2 - \frac{2m}{(m+2)} = \frac{4}{m+2}, \quad a(u, v, w) = 4v = 2,$$

$$e^{-\frac{1}{4}(\operatorname{tr} X^2 - \frac{1}{m+2}(\operatorname{tr} X)^2)} = \left(\frac{2}{\pi}\right)^{\frac{1}{2}} \int_{\mathbb{R}} e^{-\frac{4t^2}{m+2}} e^{-\frac{1}{2} \operatorname{tr}(X+t\sqrt{\frac{2}{m+2}}\mathbb{1}_m)^2} dt$$

$$(s = \sqrt{\frac{2}{m+2}}t)$$

$$= \left(\frac{2(m+2)}{\pi}\right)^{\frac{1}{2}} \int_{\mathbb{R}} e^{-\frac{1}{2} \operatorname{tr}(X+\frac{s}{\sqrt{2}}\mathbb{1}_m)^2} e^{-s^2} ds = \left(\frac{m+2}{2}\right)^{\frac{1}{2}} \int_{\mathbb{R}} e^{-\frac{1}{2} \operatorname{tr}(X-s\mathbb{1}_m)^2} \cdot \underbrace{\frac{e^{-2s^2}}{\sqrt{\frac{\pi}{2}}}}_{d\gamma(s)} ds.$$

Hence

$$\begin{aligned} I_m &= \frac{(m+2)^{\frac{1}{2}}}{\underbrace{2^{\frac{1}{2}}(2\pi)^{\frac{m(m+1)}{4}}\sqrt{\mu_m}}_{=:A_m}} \int_{\mathbb{R}} \left( \int_{\mathcal{S}_m} |\det X| e^{-\frac{1}{2} \operatorname{tr}(X-s\mathbb{1}_m)^2} |dX| \right) d\gamma(s) \\ (4.1) \quad &= A_m \int_{\mathbb{R}} \underbrace{\left( \int_{\mathcal{S}_m} |\det(x\mathbb{1}_m - Y)| e^{-\frac{1}{2} \operatorname{tr} Y^2} |dY| \right)}_{=:f_m(x)} d\gamma(x). \end{aligned}$$

For any  $O(n)$ -invariant function  $f : \mathcal{S}_n \rightarrow \mathbb{R}$  we have a Weyl integration formula (see [3, 16, 26]),

$$\frac{1}{(2\pi)^{\frac{\dim \mathcal{S}_m}{2}}} \int_{\mathcal{S}_n} f(X) |dX| = \frac{1}{Z_n} \int_{\mathbb{R}^n} f(\lambda) |\Delta_m(\lambda)| |d\lambda|,$$

where

$$\Delta_n(\lambda) := \prod_{1 \leq i < j \leq n} (\lambda_j - \lambda_i),$$

and the constant  $Z_n$  is defined by the equality [3, Eq. (2.5.11)] ,

$$(4.2) \quad Z_m := \int_{\mathbb{R}^n} e^{-\frac{1}{2}|\lambda|^2} |\Delta_m(\lambda)| |d\lambda| = 2^{\frac{n}{2}} n! \prod_{j=1}^n \Gamma\left(\frac{j}{2}\right).$$

Now observe that for any  $\lambda_0 \in \mathbb{R}$  we have (with  $f_m$  defined in (4.1))

$$\begin{aligned} f_m(\lambda_0) &= \frac{(2\pi)^{\frac{\dim S_m}{2}}}{\mathbf{Z}_m} \int_{\mathbb{R}^m} e^{-\frac{|\lambda|^2}{2}} \left( \prod_{j=1}^n |\lambda_j - \lambda_0| \right) |\Delta_m(\lambda)| |d\lambda| \\ &= \frac{e^{\frac{1}{2}\lambda_0^2} (2\pi)^{\frac{\dim S_m}{2}}}{\mathbf{Z}_m} \int_{\mathbb{R}^m} e^{-\frac{1}{2}\sum_{i=1}^m \lambda_i^2} |\Delta_{m+1}(\lambda_0, \lambda_1, \dots, \lambda_m)| |d\lambda_1 \cdots d\lambda_m| \\ &= \frac{e^{\frac{1}{2}\lambda_0^2} (2\pi)^{\frac{\dim S_m}{2}} \mathbf{Z}_{m+1}}{\mathbf{Z}_m} \underbrace{\frac{1}{\mathbf{Z}_{m+1}} \int_{\mathbb{R}^m} e^{-\frac{1}{2}\sum_{i=1}^m \lambda_i^2} |\Delta_{m+1}(\lambda_0, \lambda_1, \dots, \lambda_m)| |d\lambda_1 \cdots d\lambda_m|}_{=:\rho_{m+1}(\lambda_0)}. \end{aligned}$$

The function  $R_n(x) = n\rho_n(x)$  is known in random matrix theory as the *1-point correlation function* of the Gaussian orthogonal ensemble of symmetric  $n \times n$  matrices, [13, §4.4.1], [18, §3], [26, §4.2]. We conclude that

$$I_m = \frac{(2\pi)^{\frac{\dim S_m}{2}} A_m \mathbf{Z}_{m+1}}{\mathbf{Z}_m} \int_{\mathbb{R}} \rho_{m+1}(x) e^{\frac{x^2}{2}} d\gamma(x) = \frac{A_m \mathbf{Z}_{m+1}}{\mathbf{Z}_m} \int_{\mathbb{R}} \rho_{m+1}(x) \sqrt{\frac{2}{\pi}} e^{-\frac{3x^2}{2}} dx.$$

We have

$$\frac{\mathbf{Z}_{m+1}}{\mathbf{Z}_m} = 2^{\frac{1}{2}}(m+1)\Gamma\left(\frac{m+1}{2}\right),$$

$$\begin{aligned} \sqrt{\frac{2}{\pi}} \frac{(2\pi)^{\frac{\dim S_m}{2}} A_m \mathbf{Z}_{m+1}}{\mathbf{Z}_m} &= (2\pi)^{\frac{\dim S_m}{2}} \sqrt{\frac{2}{\pi}} 2^{\frac{1}{2}}(m+1)\Gamma\left(\frac{m+1}{2}\right) \frac{(m+2)^{\frac{1}{2}}}{2^{\frac{1}{2}}(2\pi)^{\frac{m(m+1)}{4}} \sqrt{\mu_m}} \\ &= \sqrt{\frac{2}{\pi}}(m+1)\Gamma\left(\frac{m+1}{2}\right) \frac{(m+2)^{\frac{1}{2}}}{\sqrt{\mu_m}} \\ &= \sqrt{\frac{2}{\pi}}(m+1)\Gamma\left(\frac{m+1}{2}\right) \frac{(m+2)^{\frac{1}{2}}}{2^{\frac{1}{2}}\binom{m-1}{2} + \frac{m-1}{2}}(m+2)^{\frac{1}{2}} \\ &= \sqrt{\frac{2}{\pi}}(m+1) \frac{\Gamma\left(\frac{m+1}{2}\right)}{2^{\frac{1}{2}}\binom{m-1}{2} + \frac{m-1}{2}} = \sqrt{\frac{2}{\pi}} \frac{2\Gamma\left(\frac{m+3}{2}\right)}{2^{\frac{1}{2}}\binom{m-1}{2} + \frac{m-1}{2}}. \end{aligned}$$

We deduce

$$(4.3) \quad I_m = \sqrt{\frac{2}{\pi}} \frac{2\Gamma\left(\frac{m+3}{2}\right)}{2^{\frac{1}{2}}\binom{m-1}{2} + \frac{m-1}{2}} \int_{\mathbb{R}} \rho_{m+1}(x) e^{-\frac{3x^2}{2}} dx.$$

We set

$$\bar{\rho}_n(s) := \sqrt{n}\rho_n(\sqrt{n}s),$$

and we deduce

$$\begin{aligned} \int_{\mathbb{R}} \rho_n(x) e^{-\frac{3x^2}{2}} dx &= \int_{\mathbb{R}} \rho_n(\sqrt{n}s) e^{-\frac{3ns^2}{2}} ds = n \int_{\mathbb{R}} e^{-\frac{3ns^2}{2}} \bar{\rho}_n(s) ds \\ (4.4) \quad &= \left(\frac{2\pi}{3n}\right)^{\frac{1}{2}} \int_{\mathbb{R}} \underbrace{\frac{(3n)^{\frac{1}{2}} e^{-\frac{3ns^2}{2}}}{(2\pi)^{\frac{1}{2}}}}_{=:w_n(s)} \cdot \bar{\rho}_n(s) ds. \end{aligned}$$

To proceed further we use as guide Wigner’s theorem, [3, 13, 16, 26] stating that the sequences of probability measures

$$\bar{\rho}_n(x)dx = \sqrt{n}\rho_n(\sqrt{n}x)dx = \frac{1}{\sqrt{n}}R_n(\sqrt{n}x)dx$$

converges weakly to the semi-circle probability measure<sup>5</sup>  $\rho(x)dx$ ,

$$(4.5) \quad \rho(x) = \frac{1}{\pi} \begin{cases} \sqrt{2-x^2}, & |x| \leq \sqrt{2} \\ 0, & |x| > \sqrt{2}. \end{cases}$$

We observe that the Gaussian measures  $w_n(s)ds$  converge to the Dirac delta measure concentrated at the origin. This suggests that

$$(4.6) \quad \lim_{n \rightarrow \infty} \int_{\mathbb{R}} \bar{\rho}_n(s)w_n(s)ds = \rho(0) = \frac{\sqrt{2}}{\pi}.$$

We will show that this is indeed the case by slightly refining the arguments in one particular proof of Wigner’s theorem; see [16, §7.1.6],[18, §6.1] or [26, A.9]. For the moment we will take (4.6) for granted and show that it immediately implies (1.4).

Using (4.6) in (4.3) and (4.4) we deduce that

$$I_m \sim \sqrt{\frac{2}{\pi}} \frac{2\Gamma\left(\frac{m+3}{2}\right)}{2^{\frac{1}{2}\left(\frac{m-1}{2}\right)+\frac{m-1}{2}}} \times \left(\frac{2\pi}{3(m+1)}\right)^{\frac{1}{2}} \times \frac{\sqrt{2}}{\pi} \text{ as } m \rightarrow \infty.$$

We now invoke Stirling’s formula to conclude that

$$(4.7) \quad \log I_m \sim \log \Gamma\left(\frac{m+3}{2}\right) \sim \frac{m}{2} \log m, \text{ as } m \rightarrow \infty.$$

Form (1.3) we deduce that

$$\log C(m) = \log I_m + \frac{m}{2} \log 4\pi + \log \Gamma\left(1 + \frac{m}{2}\right) - \frac{m}{2} \log(m+4).$$

Stirling’s formula and (4.7) imply that

$$\log C(m) \sim \log I_m \sim \frac{m}{2} \log m \text{ as } m \rightarrow \infty.$$

This proves (1.4).  $\square$

**4.2. Wigner’s semicircle law revisited.** We can now present the postponed proof of (4.6). The 1-point correlation function  $R_n(x)$  can be expressed explicitly in terms of Hermite polynomials, [26, Eq. (7.2.32) and §A.9],

$$(4.8) \quad R_n(x) = \underbrace{\sum_{k=0}^{n-1} \psi_k(x)^2}_{=: \mathbf{k}_n(x)} + \underbrace{\left(\frac{n}{2}\right)^{\frac{1}{2}} \psi_{n-1}(x) \int_{\mathbb{R}} \varepsilon(x-t)\psi_n(t)dt + \alpha_n(x)}_{=: \mathbf{l}_n(x)},$$

---

<sup>5</sup>There are different rescalings of the semicircle measures in the literature. Our conventions agree with those in [26].

where

$$\psi_n(x) = \frac{1}{(2^n n! \sqrt{\pi})^{\frac{1}{2}}} e^{-\frac{x^2}{2}} H_n(x), \quad H_n(x) = (-1)^n e^{x^2} \frac{d^n}{dx^n} (e^{-x^2}),$$

$$\alpha_n(x) = \begin{cases} 0, & n \in 2\mathbb{Z}, \\ \frac{\psi_{n-1}(x)}{\int_{\mathbb{R}} \psi_{n-1}(x) dx}, & n \in 2\mathbb{Z} + 1, \end{cases}$$

and

$$\varepsilon(x) = \begin{cases} \frac{1}{2}, & x > 0 \\ 0, & x = 0, \\ -\frac{1}{2}, & x < 0. \end{cases}$$

From the Christoffel-Darboux formula [43, Eq. (5.5.9)] we deduce

$$\pi^{\frac{1}{2}} e^{x^2} \sum_{k=0}^{n-1} \psi_k(x)^2 = \sum_{k=1}^{n-1} \frac{1}{2^k k!} H_k(x)^2 = \frac{1}{2^n (n-1)!} (H'_n(x) H_{n-1}(x) - H_n(x) H'_n(x))$$

Using the recurrence formula  $H'_n = 2xH_n - H_{n+1}$  we deduce

$$H'_n(x) H_{n-1}(x) - H_n(x) H'_n(x) = H_n^2(x) - H_{n-1}(x) H_{n+1}(x)$$

and

$$\mathbf{k}_n(x) = \frac{e^{-x^2}}{2^n (n-1)! \pi^{\frac{1}{2}}} (H_n^2(x) - H_{n-1}(x) H_{n+1}(x)).$$

We set

$$\bar{\mathbf{k}}_n(x) := \frac{\mathbf{k}_n(\sqrt{n}x)}{\sqrt{n}}, \quad \bar{\boldsymbol{\ell}}_n(x) := \frac{\boldsymbol{\ell}_n(\sqrt{n}x)}{\sqrt{n}}, \quad \bar{R}_n(x) = \frac{1}{\sqrt{n}} R_n(\sqrt{n}x) = \bar{\rho}_n(x)$$

so that

$$\bar{R}_n(x) = \bar{\mathbf{k}}_n(x) + \bar{\boldsymbol{\ell}}_n(x).$$

LEMMA 4.1.

$$(4.9) \quad \lim_{n \rightarrow \infty} \sup_{x \in \mathbb{R}} |\bar{\boldsymbol{\ell}}_n(x)| = 0.$$

*Proof.* Using the generating series [43, Eq. (5.5.7)]

$$\sum_{n=0}^{\infty} H_n(x) \frac{T^n}{n!} = e^{2Tx - T^2}$$

we deduce that

$$\sum_{n=0}^{\infty} \left( \int_{\mathbb{R}} e^{-\frac{x^2}{2}} H_n(x) dx \right) \frac{T^n}{n!} = e^{T^2} \int_{\mathbb{R}} e^{-\frac{(x-2T)^2}{2}} dx = \sqrt{2\pi} e^{T^2},$$

so that

$$\frac{1}{(2n)!} \int_{\mathbb{R}} e^{-\frac{x^2}{2}} H_{2n}(x) dx = \frac{\sqrt{2\pi}}{n!} \text{ and } \int_{\mathbb{R}} \psi_{2n}(x) dx = \frac{\sqrt{2(2n)!}}{2^n n! \pi^{\frac{1}{4}}} \sim \text{const} \cdot n^{\frac{1}{4}} \text{ as } n \rightarrow \infty.$$

Using [13, Thm. 6.55] or [43, Thm. 8.91.3] we deduce that

$$\sup_{x \in \mathbb{R}} |\psi_n(x)| = O(n^{-\frac{1}{12}})$$

and thus

$$\sup_{x \in \mathbb{R}} |\alpha_n(x)| = O(n^{-\frac{1}{12} - \frac{1}{4}}) = O(n^{-\frac{1}{3}}) \text{ as } n \rightarrow \infty.$$

We set

$$F_n(x) = \int_{\mathbb{R}} \varepsilon(x-t)\psi_n(t)dt.$$

Using [13, Thm. 6.55 + Eq. (6.26)] we deduce  $\sup_{x \in \mathbb{R}} |F_n(x)| = O(n^{-\frac{1}{12}})$ . This proves (4.9).  $\square$

From the above lemma we deduce that

$$\int_{\mathbb{R}} (\bar{\rho}_n(s) - \rho(s)) w_n(s) ds = \int_{\mathbb{R}} (\bar{\mathbf{k}}_n(s) - \rho(s)) w_n(s) ds + O(n^{-\frac{1}{12}}) \text{ as } n \rightarrow \infty.$$

LEMMA 4.2.

$$\lim_{n \rightarrow \infty} \int_{\mathbb{R}} (\bar{\mathbf{k}}_n(s) - \rho(s)) w_n(s) ds = 0.$$

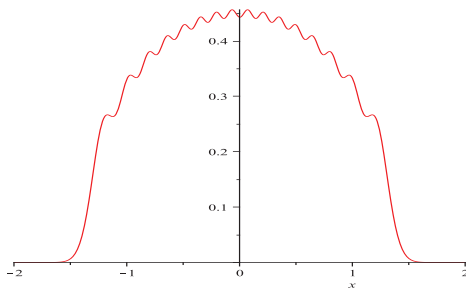


FIG. 1. The graph of  $\bar{\mathbf{k}}_{16}(x)$ ,  $|x| \leq 2$ .

*Proof.* Fix  $c \in (0, \sqrt{2})$  so that the interval  $(-c, c)$  lies inside the oscillatory regime of  $H_n(\sqrt{nt})$ . We have

$$\begin{aligned} & \int_{\mathbb{R}} (\bar{\mathbf{k}}_n(s) - \rho(s)) w_n(s) ds \\ &= \int_{|s| \leq c} (\bar{\mathbf{k}}_n(s) - \rho(s)) w_n(s) ds + \int_{|s| > c} (\bar{\mathbf{k}}_n(s) - \rho(s)) w_n(s) ds \\ &\leq \sup_{|s| \leq c} |\bar{\mathbf{k}}_n(s) - \rho(s)| + \sup_{|s| > c} |(\bar{\mathbf{k}}_n(s) - \rho(s))| \int_{|s| > c} w_n(s) ds. \end{aligned}$$



Using the Plancherel-Rotach formulæ ([13, Eq. (6.126)], [38], [43, Thm. 8.22.9]) and arguing as in [16, §7.1.6] or [18, §6.1] we deduce that

$$\lim_{n \rightarrow \infty} \sup_{|s| \leq c} |\bar{\mathbf{k}}_n(s) - \rho(s)| = 0.$$

On the other hand

$$\lim_{n \rightarrow \infty} \int_{|s| > c} w_n(s) ds = 0,$$

and [43, Thm.8.91.3] implies that

$$\sup_{|s| > c} |(\bar{\mathbf{k}}_n(s) - \rho(s))| = O(1) \text{ as } n \rightarrow \infty.$$

□

Since  $w_n(s)ds$  converges to the  $\delta$ -measure concentrated at the origin we deduce

$$\lim_{n \rightarrow \infty} \int_{\mathbb{R}} \rho(s)w_n(s)ds = \rho(0) = \frac{\sqrt{2}}{\pi}.$$

This proves (4.6).

**Appendix A. Gaussian measures and Gaussian random fields.** For the reader's convenience we survey here a few basic facts about Gaussian measures. For more details we refer to [9]. A *Gaussian measure* on  $\mathbb{R}$  is a Borel measure  $\gamma_{m,\sigma}$  of the form

$$\gamma_{m,\sigma}(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-m)^2}{2\sigma^2}} dx.$$

The scalar  $m$  is called the *mean* while  $\sigma$  is called the *standard deviation*. We allow  $\sigma$  to be zero in which case

$$\gamma_{m,0} = \delta_m = \text{the Dirac measure on } \mathbb{R} \text{ concentrated at } m.$$

Suppose that  $\mathbf{V}$  is a finite dimensional vector space. A *Gaussian measure* on  $\mathbf{V}$  is a Borel measure  $\gamma$  on  $\mathbf{V}$  such that, for any  $\xi \in \mathbf{V}^\vee$ , the pushforward  $\xi_*(\gamma)$  is a Gaussian measure on  $\mathbb{R}$ ,  $\xi_*(\gamma) = \gamma_{m(\xi),\sigma(\xi)}$ .

The map  $\mathbf{V}^\vee \ni \xi \mapsto m(\xi) \in \mathbb{R}$  is linear, and thus can be identified with a vector  $\mathbf{m}_\gamma \in \mathbf{V}$  called the *barycenter* or *expectation* of  $\gamma$  that can be alternatively defined by the equality  $\mathbf{m}_\gamma = \int_{\mathbf{V}} \mathbf{v} d\gamma(\mathbf{v})$ . Moreover, there exists a nonnegative definite, symmetric bilinear map

$$\Sigma : \mathbf{V}^\vee \times \mathbf{V}^\vee \rightarrow \mathbb{R} \text{ such that } \sigma(\xi)^2 = \Sigma(\xi, \xi), \quad \forall \xi \in \mathbf{V}^\vee.$$

The form  $\Sigma$  is called the *covariance form* and can be identified with a linear operator  $\mathbf{S} : \mathbf{V}^\vee \rightarrow \mathbf{V}$  such that

$$\Sigma(\xi, \eta) = \langle \xi, \mathbf{S}\eta \rangle, \quad \forall \xi, \eta \in \mathbf{V}^\vee,$$

where  $\langle -, - \rangle : \mathbf{V}^\vee \times \mathbf{V} \rightarrow \mathbb{R}$  denotes the natural bilinear pairing between a vector space and its dual. The operator  $\mathbf{S}$  is called the *covariance operator* and it is explicitly described by the integral formula

$$\langle \xi, \mathbf{S}\eta \rangle = \Lambda(\xi, \eta) = \int_{\mathbf{V}} \langle \xi, \mathbf{v} - \mathbf{m}_\gamma \rangle \langle \eta, \mathbf{v} - \mathbf{m}_\gamma \rangle d\gamma(\mathbf{v}).$$

The Gaussian measure is said to be *nondegenerate* if  $\Sigma$  is nondegenerate, and it is called *centered* if  $\mathbf{m} = 0$ . A nondegenerate Gaussian measure on  $\mathbf{V}$  is uniquely determined by its covariance form and its barycenter.

EXAMPLE A.1. *Suppose that  $\mathbf{U}$  is an  $n$ -dimensional Euclidean space with inner product  $(-, -)$ . We use the inner product to identify  $\mathbf{U}$  with its dual  $\mathbf{U}^\vee$ . If  $A : \mathbf{U} \rightarrow \mathbf{U}$  is a symmetric, positive definite operator, then*

$$(A.1) \quad d\gamma_A(\mathbf{x}) = \frac{1}{(2\pi)^{\frac{n}{2}} \sqrt{\det A}} e^{-\frac{1}{2}(A^{-1}\mathbf{u}, \mathbf{u})} |d\mathbf{u}|$$

is a centered Gaussian measure on  $\mathbf{U}$  with covariance form described by the operator  $A$ .  $\square$

If  $\mathbf{V}$  is a finite dimensional vector space equipped with a Gaussian measure  $\gamma$  and  $\mathbf{L} : \mathbf{V} \rightarrow \mathbf{U}$  is a linear map then the pushforward  $\mathbf{L}_*\gamma$  is a Gaussian measure on  $\mathbf{U}$  with barycenter

$$\mathbf{m}_{\mathbf{L}_*\gamma} = \mathbf{L}(\mathbf{m}_\gamma)$$

and covariance form

$$\Sigma_{\mathbf{L}_*\gamma} : \mathbf{U}^\vee \times \mathbf{U}^\vee \rightarrow \mathbb{R}, \quad \Sigma_{\mathbf{L}_*\gamma}(\eta, \eta) = \Sigma_\gamma(\mathbf{L}^\vee\eta, \mathbf{L}^\vee\eta), \quad \forall \eta \in \mathbf{U}^\vee,$$

where  $\mathbf{L}^\vee : \mathbf{U}^\vee \rightarrow \mathbf{V}^\vee$  is the dual (transpose) of the linear map  $\mathbf{L}$ . Observe that if  $\gamma$  is nondegenerate and  $\mathbf{L}$  is surjective, then  $\mathbf{L}_*\gamma$  is also nondegenerate.

Suppose  $(\mathcal{S}, \mu)$  is a probability space. A *Gaussian* random vector on  $(\mathcal{S}, \mu)$  is a (Borel) measurable map

$$X : \mathcal{S} \rightarrow \mathbf{V}, \quad \mathbf{V} \text{ finite dimensional vector space}$$

such that  $X_*\mu$  is a Gaussian measure on  $\mathbf{V}$ . We will refer to this measure as the *associated Gaussian measure*, we denote it by  $\gamma_X$  and we denote by  $\Sigma_X$  (respectively  $\mathbf{S}_X$ ) its covariance form (respectively operator),

$$\Sigma_X(\xi_1, \xi_2) = \mathbf{E}(\langle \xi_1, X - \mathbf{E}(X) \rangle \langle \xi_2, X - \mathbf{E}(X) \rangle).$$

Note that the expectation of  $\gamma_X$  is precisely the expectation of  $X$ . The random vector is called *nondegenerate*, respectively *centered*, if the Gaussian measure  $\gamma_X$  is such.

Suppose that  $X_j : \mathcal{S} \rightarrow \mathbf{V}_1$ ,  $j = 1, 2$ , are two *centered* Gaussian random vectors such that the direct sum  $X_1 \oplus X_2 : \mathcal{S} \rightarrow \mathbf{V}_1 \oplus \mathbf{V}_2$  is also a centered Gaussian random vector with associated Gaussian measure

$$\gamma_{X_1 \oplus X_2} = p_{X_1 \oplus X_2}(\mathbf{x}_1, \mathbf{x}_2) |d\mathbf{x}_1 d\mathbf{x}_2|.$$

We obtain a bilinear form

$$\mathbf{cov}(X_1, X_2) : \mathbf{V}_1^\vee \times \mathbf{V}_2^\vee \rightarrow \mathbb{R}, \quad \mathbf{cov}(X_1, X_2)(\xi_1, \xi_2) = \Sigma(\xi_1, \xi_2),$$

called the *covariance form*. The random vectors  $X_1$  and  $X_2$  are independent if and only if they are uncorrelated, i.e.,

$$\mathbf{cov}(X_1, X_2) = 0.$$

We can form the random vector  $\mathbf{E}(X_1|X_2)$ , the conditional expectation of  $X_1$  given  $X_2$ . If  $X_1$  and  $X_2$  are independent then  $\mathbf{E}(X_1|X_2) = \mathbf{E}(X_1)$ , while at the other extreme we have  $\mathbf{E}(X_1|X_1) = X_1$ .

To find a formula for  $\mathbf{E}(X_1|X_2)$  in general we fix Euclidean metrics  $(-, -)_{\mathbf{V}_j}$  on  $\mathbf{V}_j$ . We can then identify  $\mathbf{cov}(X_1, X_2)$  with a linear operator  $\mathbf{Cov}(X_1, X_2) : \mathbf{V}_2 \rightarrow \mathbf{V}_1$ , via the equality

$$\begin{aligned} \mathbf{E}(\langle \xi_1, X_1 \rangle \langle \xi_2, X_2 \rangle) &= \mathbf{cov}(X_1, X_2)(\xi_1, \xi_2) \\ &= \langle \xi_1, \mathbf{Cov}(X_1, X_2)\xi_2^\dagger \rangle, \quad \forall \xi_1 \in \mathbf{V}_1^\vee, \quad \xi_2 \in \mathbf{V}_2^\vee, \end{aligned}$$

where  $\xi_2^\dagger \in \mathbf{V}_2$  denotes the vector metric dual to  $\xi_2$ . The operator  $\mathbf{Cov}(X_1, X_2)$  is called the *covariance operator* of  $X_1, X_2$ . For a proof of the next classical result we refer to [4, Prop. 1.2].

LEMMA A.2 (Regression formula). *If  $X_1$  and  $X_2$  are as above and, additionally,  $X_2$  is nondegenerate, then*

$$(A.2) \quad \mathbf{E}(X_1|X_2) = \mathbf{Cov}(X_1, X_2)\mathbf{S}_{X_2}^{-1}(X_2 - \mathbf{E}(X_2)) + \mathbf{E}(X_1).$$

□

The conditional probability density of  $X_1$  given that  $X_2 = \mathbf{x}_2$  is the function

$$p_{(X_1|X_2=\mathbf{x}_2)}(\mathbf{x}_1) = \frac{p_{X_1 \oplus X_2}(\mathbf{x}_1, \mathbf{x}_2)}{\int_{\mathbf{V}_1} p_{X_1 \oplus X_2}(\mathbf{x}_1, \mathbf{x}_2) |d\mathbf{x}_1|}.$$

For a measurable function  $f : \mathbf{V}_1 \rightarrow \mathbb{R}$  the conditional expectation  $\mathbf{E}(f(X_1)|X_2 = \mathbf{x}_2)$  is the (deterministic) scalar

$$\mathbf{E}(f(X_1)|X_2 = \mathbf{x}_2) = \int_{\mathbf{V}_1} f(\mathbf{x}_1) p_{(X_1|X_2=\mathbf{x}_2)}(\mathbf{x}_1) |d\mathbf{x}_1|.$$

Again, if  $X_2$  is nondegenerate, then we have the *regression formula*

$$(A.3) \quad \mathbf{E}(f(X_1)|X_2 = \mathbf{x}_2) = \mathbf{E}(f(Y + C\mathbf{x}_2))$$

where  $Y : \mathcal{S} \rightarrow \mathbf{V}_1$  is a Gaussian vector with

$$(A.4) \quad \mathbf{E}(Y) = \mathbf{E}(X_1) - C\mathbf{E}(X_2), \quad \mathbf{S}_Y = \mathbf{S}_{X_1} - \mathbf{Cov}(X_1, X_2)\mathbf{S}_{X_2}^{-1}\mathbf{Cov}(X_2, X_1),$$

and  $C$  is given by

$$(A.5) \quad C = \mathbf{Cov}(X_1, X_2)\mathbf{S}_{X_2}^{-1}.$$

Let us point out that if  $X : \mathcal{S} \rightarrow \mathbf{U}$  is a Gaussian random vector and  $\mathbf{L} : \mathbf{U} \rightarrow \mathbf{V}$  is a linear map, then the random vector  $\mathbf{L}X : \mathcal{S} \rightarrow \mathbf{V}$  is also Gaussian. Moreover

$$\mathbf{E}(\mathbf{L}X) = \mathbf{L}\mathbf{E}(X), \quad \Sigma_{\mathbf{L}X}(\xi, \xi) = \Sigma_X(\mathbf{L}^\vee\xi, \mathbf{L}^\vee\xi), \quad \forall \xi \in \mathbf{V}^\vee,$$

where  $\mathbf{L}^\vee : \mathbf{V}^\vee \rightarrow \mathbf{U}^\vee$  is the linear map dual to  $\mathbf{L}$ . Equivalently,  $\mathbf{S}_{\mathbf{L}X} = \mathbf{L}\mathbf{S}_X\mathbf{L}^\vee$ .

A *random field* (or *function*) on a set  $\mathbf{T}$  is a map  $\xi : \mathbf{T} \times (\mathcal{S}, \mu) \rightarrow \mathbb{R}$ ,  $(t, s) \mapsto \xi_t(s)$  such that

- $(\mathcal{S}, \mu)$  is a probability space, and

- for any  $t \in \mathbf{T}$  the function  $\xi_t : \mathcal{S} \rightarrow \mathbb{R}$  is measurable, i.e., it is a random variable.

Thus, a random field on  $\mathbf{T}$  is a family of random variables  $\xi_t$  parameterized by the set  $\mathbf{T}$ . For simplicity we will assume that all these random variables have finite second moments. For any  $t \in \mathbf{T}$  we denote by  $\mu_{t_1}$  the expectation of  $\xi_t$ . The *covariance function* or *kernel* of the field is the function  $C_\xi : \mathbf{T} \times \mathbf{T} \rightarrow \mathbb{R}$  defined by

$$C_\xi(t_1, t_2) = \mathbf{E}((\xi_{t_1} - \mu_{t_1})(\xi_{t_2} - \mu_{t_2})) = \int_{\mathcal{S}} (\xi_{t_1}(s) - \mu_{t_1})(\xi_{t_2}(s) - \mu_{t_2}) d\mu(s).$$

The field is called *Gaussian* if for any finite subset  $F \subset \mathbf{T}$  the random vector

$$\mathcal{S} \in s \mapsto (\xi_t(s))_{t \in F} \in \mathbb{R}^F$$

is a Gaussian random vector. Almost all the important information concerning a Gaussian random field can be extracted from its covariance kernel. For more information about random fields we refer to [1, 4, 12, 20].

In the conclusion of this section we want to describe a few simple integral formulas.

PROPOSITION A.3. *Suppose  $\mathbf{V}$  is an Euclidean space of dimension  $N$ ,  $f : \mathbf{U} \rightarrow \mathbb{R}$  is a locally integrable, positively homogeneous function of degree  $k \geq 0$ , and  $A : \mathbf{U} \rightarrow \mathbf{U}$  is a positive definite symmetric operator. Denote by  $B(\mathbf{U})$  the unit ball of  $\mathbf{V}$  centered at the origin, and by  $S(\mathbf{U})$  its boundary. Then the following hold*

$$\begin{aligned} \text{(A.6)} \quad \frac{1}{\pi^{\frac{N}{2}}(k+N)} \int_{S(\mathbf{U})} f(\mathbf{u})|dA(\mathbf{u})| &= \frac{1}{\pi^{\frac{N}{2}}} \int_{B(\mathbf{U})} f(\mathbf{u})|d\mathbf{u}| \\ &= \frac{1}{\Gamma(1 + \frac{k+N}{2})} \int_{\mathbf{U}} f(\mathbf{u}) \frac{e^{-|\mathbf{u}|^2}}{\pi^{\frac{N}{2}}} |d\mathbf{u}|. \end{aligned}$$

$$\text{(A.7)} \quad \int_{\mathbf{U}} f(\mathbf{u}) d\gamma_{tA}(\mathbf{u}) = t^{\frac{k}{2}} \int_{\mathbf{U}} f(\mathbf{u}) d\gamma_A(\mathbf{u}) \quad \forall t > 0,$$

where  $d\gamma_A$  is the Gaussian measure defined by (A.1).

*Proof.* We have

$$\int_{B(\mathbf{U})} f(\mathbf{u})|d\mathbf{u}| = \int_0^1 t^{k+N-1} \left( \int_{S(\mathbf{U})} f(\mathbf{u})|dA(\mathbf{u})| \right) = \frac{1}{k+N} \int_{S(\mathbf{U})} f(\mathbf{u})|dA(\mathbf{u})|.$$

On the other hand

$$\begin{aligned} &\frac{1}{\pi^{\frac{N}{2}}} \int_{\mathbf{U}} f(\mathbf{u})e^{-|\mathbf{u}|^2} |d\mathbf{u}| \\ &= \frac{1}{\pi^{\frac{N}{2}}} \left( \int_0^\infty t^{k+N-1} e^{-t^2} dt \right) \int_{S(\mathbf{U})} f(\mathbf{u})|dA(\mathbf{u})| \\ &= \frac{1}{2\pi^{\frac{N}{2}}} \Gamma\left(\frac{k+N}{2}\right) \int_{S(\mathbf{U})} f(\mathbf{u})|dA(\mathbf{u})| = \frac{k+N}{2\pi^{\frac{N}{2}}} \Gamma\left(\frac{k+N}{2}\right) \int_{B(\mathbf{U})} f(\mathbf{u})|d\mathbf{u}|. \\ &= \frac{1}{\pi^{\frac{N}{2}}} \Gamma\left(1 + \frac{k+N}{2}\right) \int_{B(\mathbf{U})} f(\mathbf{u})|d\mathbf{u}|. \end{aligned}$$

This proves (A.6). The equality (A.7) follows by using the change in variables  $\mathbf{u} = t^{\frac{1}{2}}\mathbf{v}$ .  $\square$

**Appendix B. Gaussian random symmetric matrices.** We want to describe in some detail a 3-parameter family of centered Gaussian measures on  $\mathcal{S}_m$ , the vector space of real symmetric  $m \times m$  matrices,  $m > 1$ .

For any  $1 \leq i \leq j$  define  $\xi_{ij} \in \mathcal{S}_m^\vee$  so that for any  $A \in \mathcal{S}_m$

$$\xi_{ij}(A) = a_{ij} = \text{the } (i, j)\text{-th entry of the matrix } A.$$

The collection  $(\xi_{ij})_{1 \leq i \leq j \leq m}$  is a basis of the dual space  $\mathcal{S}_m^\vee$ . We denote by  $(E_{ij})_{1 \leq i \leq j}$  the dual basis of  $\mathcal{S}_m$ . More precisely,  $E_{ij}$  is the symmetric matrix whose  $(i, j)$  and  $(j, i)$  entries are 1 while all the other entries are equal to zero. For any  $A \in \mathcal{S}_m$  we have

$$A = \sum_{i \leq j} \xi_{ij}(A) E_{ij}.$$

The space  $\mathcal{S}_m$  is equipped with an inner product

$$(-, -) : \mathcal{S}_m \times \mathcal{S}_m \rightarrow \mathbb{R}, \quad (A, B) = \text{tr}(AB), \quad \forall A, B \in \mathcal{S}_m.$$

This inner product is invariant with respect to the action of  $\text{SO}(m)$  on  $\mathcal{S}_m$ . We set

$$\widehat{E}_{ij} := \begin{cases} E_{ij}, & i = j \\ \frac{1}{\sqrt{2}} E_{ij}, & i < j. \end{cases}$$

The collection  $(\widehat{E}_{ij})_{i \leq j}$  is a basis of  $\mathcal{S}_m$  orthonormal with respect to the above inner product. We set

$$\widehat{\xi}_{ij} := \begin{cases} \xi_{ij}, & i = j \\ \sqrt{2} \xi_{ij}, & i < j. \end{cases}$$

The collection  $(\widehat{\xi}_{ij})_{i \leq j}$  the orthonormal basis of  $\mathcal{S}_m^\vee$  dual to  $(\widehat{E}_{ij})$ . The volume density induced by this metric is

$$|dX| := \prod_{i \leq j} d\widehat{\xi}_{ij} = 2^{\frac{1}{2} \binom{m}{2}} \prod_{i \leq j} dx_{ij}.$$

To any numbers  $a, b, c$  satisfying the inequalities

$$(B.1) \quad a - b, \quad c, \quad a + (m - 1)b > 0.$$

we will associate a centered Gaussian measure  $\Gamma_{a,b,c}$  on  $\mathcal{S}_m$  uniquely determined by its covariance form

$$\Sigma = \Sigma_{a,b,c} : \mathcal{S}_m^\vee \times \mathcal{S}_m^\vee \rightarrow \mathbb{R}$$

defined as follows:

$$(B.2a) \quad \Sigma(\xi_{ii}, \xi_{ii}) = a, \quad \Sigma(\xi_{ii}, \xi_{jj}) = b, \quad \forall i \neq j,$$

$$(B.2b) \quad \Sigma(\xi_{ij}, \xi_{ij}) = c, \quad \Sigma(\xi_{ij}, \xi_{k\ell}) = 0, \quad \forall i < j, \quad k \leq \ell, \quad (i, j) \neq (k, \ell).$$

To see that  $\Sigma_{a,b,c}$  is positive definite if  $a, b, c$  satisfy (B.1) we decompose  $\mathcal{S}_m^\vee$  as a direct sum of subspaces

$$\mathcal{S}_m^\vee = \mathcal{D}_m \oplus \mathcal{O}_m,$$

$$\mathcal{D}_m = \text{span} \{ \xi_{ii}; 1 \leq i \leq m \}, \quad \mathcal{O}_m = \text{span} \{ \xi_{ij}; 1 \leq i < j \leq m \}, \quad \dim \mathcal{O}_m = \binom{m}{2}$$

With respect to this decomposition, and the corresponding bases of these subspaces the matrix  $Q_{a,b,c}$  describing  $\Sigma_{a,b,c}$  with respect to the basis  $(\xi_{ij})$  has a direct sum decomposition

$$Q_{a,b,c} = G_m(a, b) \oplus c \mathbb{1}_{\binom{m}{2}},$$

where  $G_m(a, b)$  is the  $m \times m$  symmetric matrix whose diagonal entries are equal to  $a$  while all the off diagonal entries are all equal to  $b$ .

The the spectrum of  $G_m(a, b)$  consists of two eigenvalues:  $(a - b)$  with multiplicity  $(m - 1)$  and the simple eigenvalue  $a - b + mb$ . Indeed, if  $C_m$  denotes the  $m \times m$  matrix with all entries equal to 1, then  $G_m(a, b) = (a - b)\mathbb{1}_m + bC_m$ . The matrix  $C_m$  has rank 1 and a single nonzero eigenvalue equal to  $m$  with multiplicity 1. This proves that  $Q_{a,b,c}$  is positive definite since its spectrum is positive. We denote by  $d\Gamma_{a,b,c}$  the centered Gaussian measure on  $\mathcal{S}_m$  with covariance form  $\Sigma_{a,b,c}$ .

Since  $\mathcal{S}_m$  is equipped with an inner product we can identify  $\Sigma_{a,b,c}$  with a symmetric, positive definite bilinear form on  $\mathcal{S}_m$ . We would like to compute the matrix  $\widehat{Q} = \widehat{Q}_{a,b,c}$  that describes  $\Sigma_{a,b,c}$  with respect to the orthonormal basis  $(\widehat{E}_{ij})_{1 \leq i \leq j}$ . We have

$$\widehat{Q}(\widehat{E}_{ii}, \widehat{E}_{ii}) = Q(\widehat{\xi}_{ii}, \widehat{\xi}_{ii}) = a, \quad \widehat{Q}(\widehat{E}_{ii}, \widehat{E}_{jj}) = b, \quad \forall i \neq j,$$

$$\widehat{Q}(\widehat{E}_{ij}, \widehat{E}_{ij}) = Q(\widehat{\xi}_{ij}, \widehat{\xi}_{ij}) = 2Q(\xi_{ij}, \xi_{ij}) = 2c, \quad \forall i < j,$$

Thus

$$(B.3) \quad \widehat{Q}_{a,b,c} = G_m(a, b) \oplus 2c \mathbb{1}_{\binom{m}{2}}.$$

If  $|\cdot|_{a,b,c}$  denotes the Euclidean norm on  $\mathcal{S}_m$  determined by  $\Sigma_{a,b,c}$  then for

$$A = \sum_{i \leq j} a_{ij} E_{ij} = \sum_i a_{ii} \widehat{E}_{ii} + \sqrt{2} \sum_{i < j} a_{ij} \widehat{E}_{ij}.$$

we have

$$\begin{aligned} |A|_{a,b,c}^2 &= a \sum_i a_{ii}^2 + 2b \sum_{i < j} a_{ii} a_{jj} + 4c \sum_{i < j} a_{ij}^2 \\ &= (a - b - 2c) \sum_i a_{ii}^2 + b \left( \sum_i a_{ii} \right)^2 + 2c \left( \sum_i a_{ii}^2 + 2 \sum_{i < j} a_{ij}^2 \right) \\ &= (a - b - 2c) \sum_i a_{ii}^2 + b(\text{tr } A)^2 + 2c \text{tr } A^2. \end{aligned}$$

Observe that when

$$(B.4) \quad a - b = 2c$$

we have

$$(B.5) \quad |A|_{a,b,c}^2 = b(\operatorname{tr} A)^2 + 2c \operatorname{tr} A^2$$

so that the norm  $|\cdot|_{a,b,c}$  and the Gaussian measure  $d\Gamma_{a,b,c}$  are  $O(m)$ -invariant. Let us point out that the space  $\mathcal{S}_m$  equipped with the Gaussian measure  $d\Gamma_{2,0,1}$  is the well known *GOE*, the *Gaussian orthogonal ensemble*.

To obtain a more concrete description of  $\Gamma_{a,b,c}$  we first identify  $\Sigma_{a,b,c}$  with a symmetric operator  $\widehat{Q}_{a,b,c} : \mathcal{S}_m \rightarrow \mathcal{S}_m$ . Using (B.3) we deduce that

$$\widehat{Q}_{a,b,c} = G(a, b) \oplus 2c \mathbb{1}_{\binom{m}{2}}.$$

Observe that

$$(B.6) \quad \det \widehat{Q}_{a,b,c} = (a - b)(a + (m - 1)b)^{m-1} (2c)^{\binom{m}{2}},$$

and

$$(B.7) \quad \widehat{Q}_{a,b,c}^{-1} = \widehat{Q}_{a',b',c'} = G_m(a', b') \oplus 2c' \mathbb{1}_{\binom{m}{2}},$$

where  $2c' = \frac{1}{2c}$  and the real numbers  $a', b'$  are determined from the linear system

$$(B.8) \quad \begin{cases} a' - b' &= \frac{1}{a-b} \\ a' + (m - 1)b' &= \frac{1}{a+(m-1)b}. \end{cases}$$

We then have

$$(B.9) \quad d\Gamma_{a,b,c}(X) = \frac{1}{(2\pi)^{\frac{m(m+1)}{4}} (\det \widehat{Q}_{a,b,c})^{\frac{1}{2}}} e^{-\frac{1}{2}(\widehat{Q}_{a,b,c}^{-1} X, X)} 2^{\frac{1}{2}\binom{m}{2}} \prod_{i \leq j} dx_{ij},$$

where

$$(B.10) \quad (\widehat{Q}_{a,b,c}^{-1} X, X) = \left( a' - b' - \frac{1}{2c} \right) \sum_i x_{ii}^2 + b'(\operatorname{tr} X)^2 + \frac{1}{2c} \operatorname{tr} X^2.$$

The special case  $b = c > 0$ ,  $a = 3c$  is particularly important for our considerations. We denote by  $(-, -)_c$  and respectively  $d\Gamma_c$  the inner product and respectively the Gaussian measure on  $\mathcal{S}_m$  corresponding to the covariance form  $\Sigma_{3c,c,c}$ .

If we set  $\widehat{Q}_c := \widehat{Q}_{3c,c,c}$  then we deduce from (B.7) that

$$\widehat{Q}_c^{-1} = \widehat{Q}_{a',b',c'} = G_m(a', b') \oplus \frac{1}{2c} \mathbb{1}_{\binom{m}{2}},$$

where

$$\begin{cases} a' - b' &= \frac{1}{2c} = 2c' \\ a' + (m - 1)b' &= \frac{1}{(m+2)c}. \end{cases}$$

We deduce

$$mb' = \frac{1}{(m+2)c} - \frac{1}{2c} = -\frac{m}{2c(m+2)} \Rightarrow b' = -\frac{1}{2c(m+2)}.$$

Note that the invariance condition (B.4)  $a' - b' = 2c'$  is automatically satisfied so that

$$(\widehat{Q}_c^{-1}X, X) = \frac{1}{2c} \operatorname{tr} X^2 - \frac{1}{2c(m+2)} (\operatorname{tr} X)^2.$$

Using (B.6) and (B.9) we deduce

$$(B.11) \quad d\Gamma_c(X) = \frac{1}{(2\pi c)^{\frac{m(m+1)}{4}} \sqrt{\mu_m}} \cdot e^{-\frac{1}{4c}(\operatorname{tr} X^2 - \frac{1}{m+2}(\operatorname{tr} X)^2)} \underbrace{2^{\frac{1}{2}\binom{m}{2}} \prod_{i \leq j} |dx_{ij}|}_{|dX|},$$

where

$$(B.12) \quad \mu_m := 2^{\binom{m}{2} + (m-1)} (m+2).$$

The inner product  $(-, -)_c$  has the alternate description

$$(B.13) \quad \begin{aligned} (A, B)_c = I_c(A, B) &:= 4c \int_{\mathbb{R}^m} (A\mathbf{x}, \mathbf{x})(B\mathbf{x}, \mathbf{x}) \frac{e^{-|\mathbf{x}|^2}}{\pi^{\frac{m}{2}}} |d\mathbf{x}| \\ &= c \int_{\mathbb{R}^m} (A\mathbf{x}, \mathbf{x})(B\mathbf{x}, \mathbf{x}) \frac{e^{-\frac{|\mathbf{x}|^2}{2}}}{(2\pi)^{\frac{m}{2}}} |d\mathbf{x}|, \quad \forall A, B \in \mathcal{S}_m. \end{aligned}$$

**Appendix C. A Gaussian integral.**

*The Proof of (2.37).* We want to find the value of the integral

$$I = \frac{1}{4(2\pi)^{\frac{3}{2}}} \int_{\mathcal{S}_2} |\det X| e^{-\frac{1}{4}(\operatorname{tr} X^2 - \frac{1}{4}(\operatorname{tr} X)^2)} \cdot \sqrt{2} \prod_{1 \leq i \leq j \leq 2} dx_{ij}.$$

We first make the change in coordinates

$$x_{11} = x + y, \quad x_{22} = x - y, \quad x_{12} = z.$$

Then

$$\det X = x^2 - y^2 - z^2, \quad \operatorname{tr} X = 2x, \quad \operatorname{tr} X^2 = 2(x^2 + y^2 + z^2).$$

Hence

$$I = \frac{1}{2(2\pi)^{\frac{3}{2}}} \int_{\mathbb{R}^3} |x^2 - y^2 - z^2| e^{-\frac{1}{4}(x^2 + 2y^2 + 2z^2)} \sqrt{2} |dxdydz|$$

$$(x = \sqrt{2}u)$$

$$\begin{aligned} &= \frac{1}{(2\pi)^{\frac{3}{2}}} \int_{\mathbb{R}^3} |2u^2 - y^2 - z^2| e^{-\frac{1}{2}(u^2 + y^2 + z^2)} |dudydz| \\ &= \frac{2}{\pi^{\frac{3}{2}}} \int_{\mathbb{R}^3} |2u^2 - y^2 - z^2| e^{-(u^2 + y^2 + z^2)} |dudydz|. \end{aligned}$$



We now make the change in variables  $y = r \cos \theta$ ,  $y = r \sin \theta$ ,  $r > 0$ ,  $\theta \in [0, 2\pi)$  and we deduce

$$\begin{aligned} I &= \frac{2}{\pi^{\frac{3}{2}}} \int_{\mathbb{R}} \int_0^\infty \left( \int_0^{2\pi} |2u^2 - r^2| e^{-u^2+r^2} d\theta \right) r dr du \\ &= \frac{8\pi}{\pi^{\frac{3}{2}}} \int_0^\infty \int_0^\infty |2u^2 - r^2| e^{-(u^2+r^2)} r dr du. \end{aligned}$$

We now make the change in variables

$$u = t \sin \varphi, \quad r = t \cos \varphi, \quad t > 0, \quad 0 \leq \varphi \leq \frac{\pi}{2},$$

and we conclude

$$I = \frac{8}{\pi^{\frac{1}{2}}} \left( \int_0^\infty e^{-t^2} t^4 dt \right) \left( \int_0^{\frac{\pi}{2}} |3 \sin^2 \varphi - 1| \cos \varphi d\varphi \right)$$

( $t = \sqrt{s}$ ,  $x = \sin \varphi$ )

$$= \frac{4}{\pi^{\frac{1}{2}}} \left( \int_0^\infty e^{-s} s^{\frac{3}{2}} ds \right) \left( \int_0^1 |3x^2 - 1| dx \right) = \frac{4}{\pi^{\frac{1}{2}}} \times \Gamma\left(\frac{5}{2}\right) \times \frac{4}{3\sqrt{3}} = \frac{4}{\sqrt{3}}. \quad \square$$

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